Recent Advances in Fractional Stochastic Volatility Models

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Stochastic Volatility

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$$\{S_t; t \in [0,T]\}$$
: Stock Price Process

$$dS_t = \mu S_t dt + \sigma(X_t) S_t dW_t,$$

 $\sigma(\cdot)$ is a deterministic function and X_t is a stochastic process.

- Assumption: X_t is modeled by a diffusion driven by noise other than W.
 - Log-Normal: $dX_t = c_1 X_t dt + c_2 X_t dZ_t$,
 - Mean Reverting OU: $dX_t = \alpha (m X_t) dt + \beta dZ_t$,
 - Feller/Cox-Ingersoll-Ross (CIR): $dX_t = k (\nu X_t) dt + v \sqrt{X_t} dZ_t$.



Stylized Fact: Volatility Persistence



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Stochastic Process with Long Memory

Long-memory

If $\{X_t\}_{t\in\mathbb{N}}$ is a stationary process and there exists $H \in (\frac{1}{2}, 1)$ such that

$$\lim_{t \to \infty} \frac{Corr(X_t, X_1)}{c \ t^{2-2H}} = 1$$

then $\{X_t\}$ has long memory (long-range dependence).





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then $\{X_t\}$ has long memory (long-range dependence).

Equivalently,

Long Memory: $\sum_{t=1}^{\infty} Corr(X_t, X_1) = \infty$, when H > 1/2Antipersistence: $\sum_{t=1}^{\infty} Corr(X_t, X_1) < \infty$, when H < 1/2



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Fractional Stochastic Volatility Model

Log-Returns: $\{Y_t, t \in [0, T]\}$

$$dY_t = \left(r - \frac{{\sigma_t}^2}{2}\right) dt + \sigma_t dW_t.$$

where $\sigma_t = \sigma(X_t)$ and X_t is described by:

$$dX_t = \alpha \ (m - X_t) \ dt + \beta \ dB_t^H$$



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- B_t^H is a fractional Brownian motion with Hurst parameter $H \in (0, 1)$.
- X_t is a fractional Ornstein-Uhlenbeck process (fOU).





Fractional Brownian Motion

Definition

A centered Gaussian process $B^H = \{B_t^H, t \ge 0\}$ is called fractional Brownian motion (fBm) with selfsimilarity parameter $H \in (0, 1)$, if it has the following covariance function

$$\mathbf{E}\left(B_{t}^{H}B_{s}^{H}\right) = \frac{1}{2}\left\{t^{2H} + s^{2H} - |t-s|^{2H}\right\}.$$

and a.s. continuous paths.



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and a.s. continuous paths.

– When $H = \frac{1}{2}$, B^H is a **standard** Brownian motion.



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circFBM Path - N=1000, H=0.5

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circFBM Path - N=1000, H=0.7



circFBM Path - N=1000, H=0.95

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The increments of fBm, $\{B_n - B_{n-1}\}_{n \in \mathbb{N}}$, are

- stationary, i.e. $\mathbb{E}[(B_n B_{n-1}) (B_{n+h} B_{n+h-1})] = \gamma(h).$
- *H*-selfsimilar, i.e. $c^{-H} (B_n B_{n-1}) \sim^{\mathcal{D}} (B_{c n} B_{c(n-1)})$
- dependent
 - When $H > \frac{1}{2}$, the increments exhibit long-memory, i.e.

 $\sum \mathbb{E}\left[B_1\left(B_n-B_{n-1}\right)\right] = +\infty.$

— When $H < rac{1}{2}$, the increments exhibit antipersistence, i.e.

 $\sum \mathbb{E}\left[B_1\left(B_n-B_{n-1}\right)\right] < +\infty.$





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- It has a.s. *Hölder-continuous* sample paths of any order $\gamma < H$.
- Its $\frac{1}{H}$ -variation on [0, t] is finite. In particular, fBm has an infinite quadratic variation for H < 1/2.
- When $H \neq \frac{1}{2}$, B_t^H is not a semimartingale.



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Stochastic Integral wrt fBm

- Pathwise Riemann-Stieltjes integral, when H > 1/2. (Lin; Dai and Heyde).
- Stochastic calculus of variations with respect to a Gaussian process. (Decreusefond and Üstünel; Carmona and Coutin; Alòs, Mazet and Nualart; Duncan, Hu and Pasik-Duncan and Hu and Oksendal).
- Pathwise stochastic integral interpreted in the Young sense, when H > 1/2 (Young; Gubinelli).
- Rough path-theoretic approach by T. Lyons.





Fractional Stochastic Volatility Model

$$\begin{cases} dY_t = \left(r - \frac{\sigma^2(X_t)}{2}\right) dt + \sigma(X_t) dW_t, \\ dX_t = \alpha (m - X_t) dt + \beta dB_t^H \end{cases}$$

Some Properties

- *Hölder continuity*: Hölder continuous of order γ , for all $\gamma < H$.
- Self-similarity: Self-similar in the sense that

$$\{B^{H}_{ct}; t \in \mathbb{R}\} \sim^{\mathcal{D}} \{c^{H}B^{H}_{t}; t \in \mathbb{R}\}, \quad \forall c \in \mathbb{R}$$

This property is approximately inherited by the fOU, for scales smaller than $1/\alpha.$



Fractional Stochastic Volatility Model

Hurst Index	Model
H > 1/2	Long-Memory SV: persistence $\label{eq:ACF} \mbox{ACF Decay} \sim dt^{2H-2}$
H < 1/2	Rough Volatility: anti-persistence ACF Decay $\sim dt^H$

H = 1/2 Classical SV



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Long Memory Stochastic Volatility Models

- Comte and Renault (1998)
- Comte, Coutin and Renault (2010)
- C. and Viens (2010, 2012)
- Gulisashvili, Viens and Zhang (2015)
- Guennoun, Jacquier, and Roome (2015)
- Garnier & Solna (2015, 2016)
- Bezborodov, Di Persio & Mishura (20176)
- Fouque & Hu (2017)



Rough Stochastic Volatility Models

- Gatheral, Jaisson, and Rosenbaum (2014)
- Bayer, Friz, and Gatheral (2015)
- Forde, Zhang (2015)
- El Euch, Rosenbaum (2016, 2017)

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Related Literature

- Willinger, Taggu, Teverovsky (1999): LRD in the stock market
- Bayraktar, Poor, Sircar (2003): Estimation of fractal dimension of S&P 500.
- Björk, Hult (2005): Fractional Black-Scholes market.
- Cheriditio (2003): Arbitrage in fractional BS market.
- Guasoni (2006): No arbitral under transaction cost with fBm.

Research Questions

Option Pricing

Statistical Inference





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- In practice, we have access to discrete-time observations of historical stock prices, while the volatility is unobserved.



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 In practice, we have access to discrete-time observations of historical stock prices, while the volatility is *unobserved*.

Two-steps

• Estimate the *empirical stochastic volatility distribution*.

- Through adjusted particle filtering algorithms.
- Construct a *multinomial recombining tree* to compute the option prices.



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Multinomial Recombining Tree



(i) At each step, a value for the volatility is drawn from the volatility filter.

- (ii) A standard pricing technique using *backward induction* can be used to compute the option price.
- iii) We iterate this procedure by using N repeated volatility samples, constructing a different tree with each sample and averaging over all prices.





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S&P 500 Data





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Statistical Inference



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$$\begin{cases} dY_t = \left(\mu - \frac{\sigma^2(X_t)}{2}\right) dt + \sigma(X_t) dW_t, \\ dX_t = \alpha (m - X_t) dt + \beta dB_t^H \end{cases}$$

- We can also assume that $Corr(B_t^H, W_t) = \rho$ (leverage effects).
- Parameters to estimate: $\theta = (\alpha, m, \beta, \mu, \rho)$ and H.

Remark

- The estimation of H is decoupled from the estimation of the drift components, but not from the estimation of the "diffusion" terms.

Framework

- Observations: Historical stock prices \rightarrow Discrete, even when in high-frequency.
- Unobserved State:

Stochastic Volatility with non-Markovian structure is hidden.



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Extension of classical statistical methods

- For *H* known:

 μ , m, α and β can be estimated with standard techniques (Fouque, Papanicolaou and Sircar, 2000) using high-frequency data.



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 - Traditional volatility proxies: Squared returns, logarithm of squared returns.



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Extension of classical statistical methods

- For *H* known:

 μ , m, α and β can be estimated with standard techniques (Fouque, Papanicolaou and Sircar, 2000) using high-frequency data.

- For H unknown:
 - Traditional volatility proxies: Squared returns, logarithm of squared returns.
 - Use a non-parametric method to estimate H through the squared returns and then go back to classical techniques and estimate the remaining parameters.



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Simulation Based Methods

- Employ a Sequential Monte Carlo (SMC) method to estimate the unobserved state along with the unknown parameters.



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Simulation Based Methods

- Employ a Sequential Monte Carlo (SMC) method to estimate the unobserved state along with the unknown parameters.
- Denote θ the vector of all parameters, except for H.



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Simulation Based Methods

- Employ a Sequential Monte Carlo (SMC) method to estimate the unobserved state along with the unknown parameters.
- Denote θ the vector of all parameters, except for H.
 - Observation equation: $f(Y_t|X_t; \theta)$



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Simulation Based Methods

- Employ a Sequential Monte Carlo (SMC) method to estimate the unobserved state along with the unknown parameters.
- Denote θ the vector of all parameters, except for H.
 - Observation equation: $f(Y_t|X_t; \theta)$
 - State equation: $f(X_t|X_{t-1},\ldots,X_1;\theta)$



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Simulation Based Methods

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 - Observation equation: $f(Y_t|X_t; \theta)$
 - State equation: $f(X_t|X_{t-1},\ldots,X_1;\theta)$
- Key Idea: Sequentially compute

$$f(X_{1:t};\theta|Y_{1:t}) \propto f(X_1) \cdot f(X_2|X_1;\theta) \cdot \ldots \cdot f(X_n|X_{n-1},\ldots,X_1;\theta)$$
$$\cdot \prod_{i=1}^t f(Y_i|X_i;\theta) \cdot f(\theta|Y_t),$$



Learning θ sequentially

Filtering for states and parameter(s): Learning X_t and θ sequentially.



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Learning θ sequentially

Filtering for states and parameter(s): Learning X_t and θ sequentially.

Posterior at
$$t : f(X_t|\theta, Y_t) f(\theta|Y_t)$$

 \Downarrow
Prior at $t + 1 : f(X_{t+1}|\theta, Y_t) f(\theta|Y_t)$
 \Downarrow
Posterior at $t + 1 : f(X_{t+1}|\theta, Y_{t+1}) f(\theta|Y_{t+1})$



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Learning θ sequentially

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```
Posterior at t : f(X_t | \theta, Y_t) f(\theta | Y_t)
                      1
     Prior at t+1: f(X_{t+1}|\theta, Y_t) f(\theta|Y_t)
                      11
Posterior at t+1: f(X_{t+1}|\theta, Y_{t+1}) f(\theta|Y_{t+1})
```

Advantages

- Sequential updates of $f(\theta|Y_t)$, $f(X_t|Y_t)$ and $f(\theta, X_t|Y_t)$.
- 2 Sequential *h*-step ahead forecasts $f(Y_{t+h}|Y_t)$
- Sequential approximations for $f(Y_t|Y_{t-1})$.





• Draw θ from a mixture of Normals:

$$f(\theta|Y_t) \approx \sum_{j=1}^{N} \omega_t^{(j)} \mathcal{N}(\theta|m_t^{(j)}, h^2 V_t)$$



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• Draw θ from a mixture of Normals:

- $\forall t \text{ update } f(\theta|Y_t)$

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• Draw θ from a mixture of Normals:

- $\forall t \text{ update } f(\theta|Y_t)$
- Compute a Monte Carlo approximation of $f(\theta|Y_t)$, by using samples $\theta_t^{(j)}$ and weights $w_t^{(j)}$.

$$f(\theta|Y_t) \approx \sum_{j=1}^{N} \omega_t^{(j)} \mathcal{N}(\theta|m_t^{(j)}, h^2 V_t)$$



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• Draw θ from a mixture of Normals:

- $\forall t \text{ update } f(\theta|Y_t)$
- Compute a Monte Carlo approximation of $f(\theta|Y_t)$, by using samples $\theta_t^{(j)}$ and weights $w_t^{(j)}$.
- Smooth kernel density approximation

$$f(\boldsymbol{\theta}|Y_t) \approx \sum_{j=1}^{N} \omega_t^{(j)} \mathcal{N}(\boldsymbol{\theta}|m_t^{(j)}, h^2 V_t)$$



Filter convergence

Let $\phi : \mathcal{X} \mapsto \mathbb{R}$ be an appropriate test function and assume that we want estimate

$$\bar{\phi}_t = \int \phi_t(x_{1:t}) p(x_{1:t}, \theta_{(t)} | Y_{1:t}) dx_{1:t} d\theta_{(t)}.$$

The SISR algorithm provides us with the estimator

$$\hat{\phi}_t^N = \int \phi_t(x_{1:t}) \pi^N(dx_{1:t}) = \sum_{i=1}^N W_t^i \phi_t\left(X_{1:t-1,t-1}^{(i)}, \tilde{X}_{t,t}^{(i)}\right)$$

CLT for the filter (C. and Spiliopoulos) $\sqrt{N}\left(\hat{\phi}_t^N - \bar{\phi}_t\right) \Rightarrow \mathcal{N}\left(0, \sigma^2(\phi_t)\right)$

as $N \to \infty$.



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Convergence of the Parameter

$$\bar{\theta}^N_{(t)} = \sum_{i=1}^N W^{(i)}_t \theta^{(N,i)}_{(t)}, \quad \text{where} \quad \theta^{(N,i)}_{(t)} \sim \mathcal{N}(m^{(N,i)}_{t-1} | h^2 V^N_{t-1})$$

with $m_{t-1}^N = \alpha \theta_{(t-1)}^{(N,i)} + (1-\alpha) \bar{\theta}_{(t-1)}^N$, $V_{t-1}^N = \frac{1}{N-1} \sum_{i=1}^N \left(W_{t-1}^{(i)} \theta_{(t-1)}^{(N,i)} - \bar{\theta}_{(t-1)}^N \right)^2$.

CLT for the parameter

Assuming that $\mathbb{E}_{\pi_{\star}^{\theta}} \|W_t\|^{2+\delta} < \infty$:

$$\sqrt{N}\left(\bar{\theta}_{(t)}^{(N)} - \bar{\theta}_{(t)}\right) \Rightarrow \mathcal{N}\left(0, \sigma^{2}(\theta_{(t)})\right), \quad \text{as} \quad N \to \infty$$
(1)

Moreover, if the model \mathbb{P}_{θ} is identifiable, then the posterior mean $\bar{\theta}_{(t)}$ consistently estimates the true parameter value θ , as $t \to \infty$.



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S& P 500: Volatility Particle Filter



S& P 500: Parameter Estimators





Model Validation: 1-Step Ahead Prediction





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Model Validation: Residuals





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- Conditionally on the past and the entire volatility path $\ln S_T / S_t \sim \mathcal{N} \left(r - (1/2) \int_t^T \sigma_s^2 ds, \int_t^T \sigma_s^2 ds \right).$ So.

$$\begin{split} C_t &= S_t \left\{ \mathbb{E}^Q \left[\Phi \left(\frac{x_t}{V_{t,T}} + \frac{V_{t,T}}{2} \right) \right] - e^{-x_t} \mathbb{E}^Q \left[\Phi \left(\frac{x_t}{V_{t,T}} - \frac{V_{t,T}}{2} \right) \right] \right\}, \\ \text{where } x_t &= \ln \left(e^{-rT} S_t / K \right) \text{ and } V_{t,T} = \left(\int_t^T \sigma_s^2 ds \right)^{1/2}. \end{split}$$

$$\Delta_t(x_t, \sigma_t) = \mathbb{E}^Q \left[\Phi \left(\frac{x_t}{V_{t,T}} + \frac{V_{t,T}}{2} \right) \right]$$



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- Conditionally on the past and the entire volatility path $\ln S_T / S_t \sim \mathcal{N} \left(r - (1/2) \int_t^T \sigma_s^2 ds, \int_t^T \sigma_s^2 ds \right).$ So.

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- Conditionally on the past and the *entire volatility path* $\ln S_T / S_t \sim \mathcal{N} \left(r - (1/2) \int_t^T \sigma_s^2 ds, \int_t^T \sigma_s^2 ds \right).$ So.

$$\begin{split} C_t &= S_t \left\{ \mathbb{E}^Q \left[\Phi \left(\frac{x_t}{V_{t,T}} + \frac{V_{t,T}}{2} \right) \right] - e^{-x_t} \mathbb{E}^Q \left[\Phi \left(\frac{x_t}{V_{t,T}} - \frac{V_{t,T}}{2} \right) \right] \right\}, \end{split}$$
 where $x_t &= \ln \left(e^{-rT} S_t / K \right)$ and $V_{t,T} = \left(\int_t^T \sigma_s^2 ds \right)^{1/2}.$

- Imperfect *Delta-Sigma* hedging strategy

$$\Delta_t(x_t, \sigma_t) = \mathbb{E}^Q \left[\Phi \left(\frac{x_t}{V_{t,T}} + \frac{V_{t,T}}{2} \right) \right]$$



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Two notions of Implied Volatility

Black-Scholes Implied Volatility:

The unique solution to

$$C_t(x,\sigma) = C_t^{BS}(x,\sigma^i(x,\sigma)),$$

i.e. the volatility parameter that equates the BS price to the HW.

$$\Delta_t(x,\sigma) = \Delta_t^{BS}(x,\sigma^h(x,\sigma))$$



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Two notions of Implied Volatility

Black-Scholes Implied Volatility:

The unique solution to

$$C_t(x,\sigma) = C_t^{BS}(x,\sigma^i(x,\sigma)),$$

i.e. the volatility parameter that equates the BS price to the HW.

Pedging Volatility:

The unique solution to

$$\Delta_t(x,\sigma) = \Delta_t^{BS}(x,\sigma^h(x,\sigma))$$

i.e. the volatility parameter that equates the BS hedge ratio against the underlying asset variations to the HW one.



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Hedging Bias (Definition)

The difference between the BS implied volatility-based hedging ratio and the HW one:

Bias
$$= \Delta_t^{BS}(x, \sigma^i(x, \sigma)) - \Delta_t(x, \sigma)$$



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Theorem (a) Sign of Hedging Bias:

$$\sigma^h(-x,\sigma) \leq \sigma^i(-x,\sigma) = \sigma^i(x,\sigma) \leq \sigma^h(x,\sigma)$$



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Theorem

(a) Sign of Hedging Bias:

$$\sigma^h(-x,\sigma) \leq \sigma^i(-x,\sigma) = \sigma^i(x,\sigma) \leq \sigma^h(x,\sigma)$$

(b) Accuracy of approximation of partial hedging ratio by the BS implicit volatility-based hedging ratio:

> $\Delta^{BS}(x,\sigma) \le \Delta(x,\sigma)$ $\Delta^{BS}(-x,\sigma) \ge \Delta(-x,\sigma)$ $\Delta^{BS}(0,\sigma) = \Delta(0,\sigma)$



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Thank you!



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