

# Averaging and Sensitivity Analysis for Multiscale Stochastic Reaction Networks

Araz Hashemi\*

M. Nunez<sup>†</sup>, P. Plecháč<sup>‡</sup>, D.G. Vlachos<sup>†</sup>

\* Control Science Center, Air Force Research Labs, RQQA

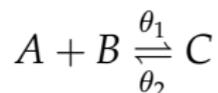
‡ Dept. of Mathematical Sciences, University of Delaware

† Dept. of Chemical Engineering, University of Delaware

araz@udel.edu

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- ▶ Example: Chemical species  $A, B, C$



- ▶  $X(t) = [A(t), B(t), C(t)] \in \mathcal{M} \subset \mathbb{Z}^3$  is viewed as a Markov Chain.

State space  $\mathcal{M}$  typically quite large.

- ▶ When reaction  $r$  fires,  $X \rightarrow X + \zeta_r$ ,  
 $\zeta_r$  is the **stoichiometric vector** corresponding to rxn  $r$
- ▶ **Propensity Functions:**  $a_r(x; \theta) =$  instant. prob. of reaction  $r$  firing  
 Mass Action :  $a_r(x; \theta) = \theta_r \cdot b_r(x)$

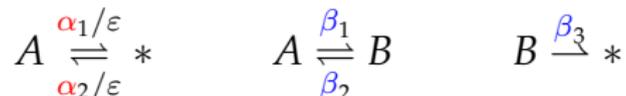
- ▶  $X(t)$  has generator  $Q = Q(\theta)$  determined by propensity functions  $a_r(x; \theta)$  and stoichiometric vectors  $\zeta_r$ .  
Large, but sparse.

$$Q(\theta) = \begin{matrix} & \dots & x + \zeta_2 & \dots & x & \dots & x + \zeta_1 & \dots \\ \vdots & & & & & & & \\ x & \left( \begin{matrix} \dots & a_2(x; \theta) & \dots & -a_0(x; \theta) & \dots & a_1(x; \theta) & \dots \end{matrix} \right) \\ \vdots & & & & & & & \end{matrix}$$

where  $a_0(x; \theta) = \sum_{r=1}^M a_r(x; \theta)$ .

- ▶ We are interested in quantities such as  $\mathbb{E}_\theta \{f(X(T))\}$ .  
Most efficient to use Monte Carlo methods to simulate trajectories and average results.

# ABSORPTION/DESORPTION WITH STIFFNESS $\varepsilon$



$\varepsilon \ll 1$  is the stiffness parameter

Write  $X = [A, *, B]$

**Fast Rxns  $\alpha$ :**

$$\zeta_{\alpha_1} = [-1, 1, 0]$$

$$\zeta_{\alpha_2} = [1, -1, 0]$$

$$a_{\alpha_1}^\varepsilon(X) = \alpha_1/\varepsilon \cdot X_1$$

$$a_{\alpha_2}^\varepsilon(X) = \alpha_2/\varepsilon \cdot X_2$$

**Slow Rxns  $\beta$ :**

$$\zeta_{\beta_1} = [-1, 0, 1]$$

$$\zeta_{\beta_2} = [1, 0, -1]$$

$$\zeta_{\beta_3} = [0, 1, -1]$$

$$a_{\beta_1}(X) = \beta_1 \cdot X_1$$

$$a_{\beta_2}(X) = \beta_2 \cdot X_3$$

$$a_{\beta_3}(X) = \beta_3 \cdot X_3$$

# COMPUTATIONAL PROBLEM

In exact simulation (SSA, Next Reaction, ...), **every** reaction is computed.

- ▶ Prob. of next rxn :  $\mathbb{P}\{r^* = r_k\} \propto a_k(x, \theta)$

As  $\varepsilon \rightarrow 0$ ,  $\mathbb{P}\{r^* \in \alpha\} \rightarrow 1$  and  $\mathbb{P}\{r^* \in \beta\} \rightarrow 0$ .

Fast rxns dominate computations!

- ▶ Each fast rxn advances time clock on micro-scale :

$$\Delta t = O(\varepsilon).$$

Prohibitive computational load to increment to large time horizons (and observe slow dynamics)

# TAU-LEAPING

[Gillespie 2001, Rathinam et al. 2003, Cao et al 2006]

$$X(t) = X(0) + \sum_{j=1}^M R_j(t) \zeta_j$$

$R_j(t)$  = # of reactions of type  $j$  which fire by time  $t$

Random Time Change representation:

$$R_j(t) = Y_j \left( \int_0^t a_j(X(s); \theta) ds \right)$$

$Y_j$  are unit-rate Poisson processes

If the value of  $a_j(X(s); \theta)$  remains essentially constant in  $(t, t + \tau]$ , then

$$R_j(t + \tau) - R_j(t) \approx \text{Poisson}(a_j(X(t); \theta) \cdot \tau)$$

Choosing the appropriate leaping value of  $\tau$  is a challenge

# STOCHASTIC AVERAGING

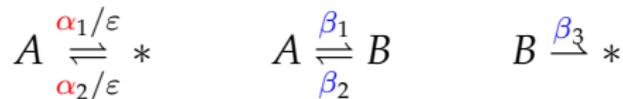
[ Cao et al. 2005, Samant & Vlachos 2005, E et al. 2007, Kang & Kurtz 2013 ]

Similar to deterministic quasi-steady state approximations:

$$\dot{X}_t^\varepsilon = f(X_t^\varepsilon, Y_t^\varepsilon, \varepsilon) \quad X_t^\varepsilon = \text{slow variables}$$

$$\dot{Y}_t^\varepsilon = \frac{1}{\varepsilon} g(X_t^\varepsilon, Y_t^\varepsilon, \varepsilon) \quad Y_t^\varepsilon = \text{fast variables}$$

As  $\varepsilon \rightarrow 0$ , one obtains a singularly perturbed system  $(X, Y)$  which can be solved by  $\dot{X}_t = f(X_t, Y_t)$  under constraint  $g(X, Y) = 0$ .



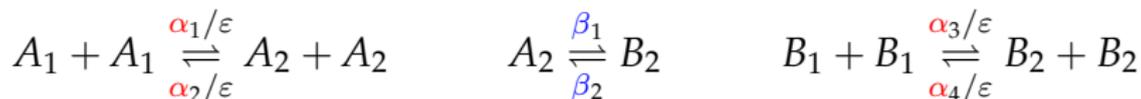
$$Y^\varepsilon(t) = (A, *), X^\varepsilon(t) = B$$

Can then separate the dynamics of fast-species and slow species in the CME under steady-state constraints.

Issues :

1. Rate of convergence in ( $\varepsilon$ )
2. Should be able to separate dynamics even if one cannot separate species into fast and slow

**Example:** Fast dimerization, slow isomerization



All species are 'fast', but there are still separable fast and slow dynamics.

- ▶ For discrete state space, state-space partitioning [Nunez and Vlachos, 2015] allows for a more useful separation of dynamics.
- ▶ We have gone further to identify the partitioned classes as states of a 'meta' Markov chain which corresponds to a 'macro' rxn network.
- ▶ This formulation allows for analysis of rate of convergence, as well as implementations of **sensitivity**

$X = [X_1, \dots, X_{N_S}]$  species

$r_{\alpha_1}, \dots, r_{\alpha_{M_f}}, r_{\beta_1}, \dots, r_{\beta_{M_s}}$  Fast and Slow rxns

$\theta = [\alpha, \beta] = [\alpha_1, \dots, \alpha_{M_f}, \beta_1, \dots, \beta_{M_s}]$  rxn parameters

Propensities:

$$a_{\alpha_j}^\varepsilon(x; \theta) = \frac{\alpha_j}{\varepsilon} b_{\alpha_j}(x) \quad a_{\beta_j}(x; \theta) = \beta_j b_{\beta_j}(x)$$

Take  $\mathcal{M}$  = State space of  $X(t)$ , assume  $|\mathcal{M}| < \infty$  for simplicity.

As  $\varepsilon \rightarrow 0$  only fast rxns  $\alpha$  fire, so define an equivalence relation on states  $s \in \mathcal{M}$  by  $s_i \leftrightarrow s_j$  if mutually accessible through only fast rxns. This gives a partition of  $\mathcal{M}$  into **fast-classes**  $\mathcal{M}_k$ :

$$\mathcal{M} = \bigcup_{k=0}^{N_C} \mathcal{M}_k = \left\{ s_1^{(1)}, \dots, s_{m_1}^{(1)}, s_1^{(2)}, \dots, s_{m_2}^{(2)}, \dots, \dots, s_{m_{N_C}}^{(N_C)} \right\}$$

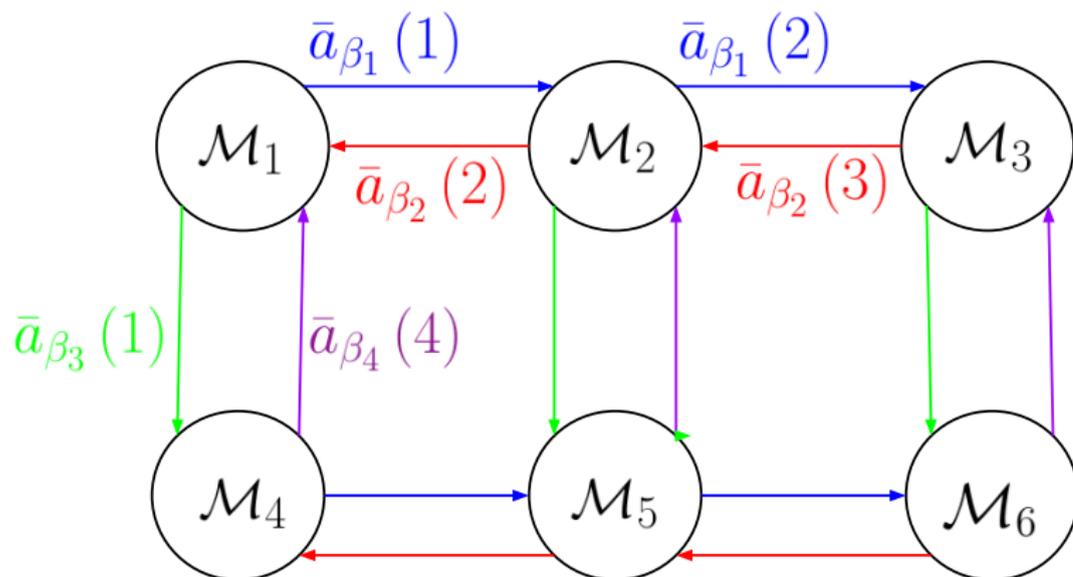
The generator of the exact process can be separated as

$$Q^\varepsilon(\boldsymbol{\theta}) = \frac{1}{\varepsilon} \tilde{Q}(\boldsymbol{\alpha}) + \hat{Q}(\boldsymbol{\beta})$$

- ▶  $\hat{Q}(\boldsymbol{\beta})$  determined only by slow rxns/propensities  $a_{\beta_j}(x, \boldsymbol{\beta})$
- ▶  $\tilde{Q}(\boldsymbol{\alpha}) = \text{diag} [Q^{(1)}(\boldsymbol{\alpha}), Q^{(2)}(\boldsymbol{\alpha}), \dots, Q^{(N_c)}(\boldsymbol{\alpha})]$   
each  $Q^{(k)}$  corresponds to fast-only dynamics of each fast-class  $\mathcal{M}_k$  under unscaled propensities  $a_{\alpha_j}(x, \boldsymbol{\alpha})$
- ▶ Each  $\mathcal{M}_k$  is an irreducible class of  $\mathcal{M}$  under  $\tilde{Q}(\boldsymbol{\alpha})$ , so there is a stationary (steady state) distribution  $\pi^{(k)}(\boldsymbol{\alpha})$
- ▶ If  $X(0) \in \mathcal{M}_k$ , then  $X(t)$  relaxes to steady-state distrn.  $\pi^{(k)}$  before slow dynamics take effect
- ▶ Slow propensities relax to  $\pi^{(k)}$  averages on macro time scale  $t$

$$a_{\beta_j}(X(t); \boldsymbol{\beta}) \approx \bar{a}_{\beta_j}(\mathcal{M}_k, \boldsymbol{\theta}) = \mathbb{E}_{\pi^{(k)}} \left\{ a_{\beta_j}(x; \boldsymbol{\beta}) \right\}$$

## TTS CRN EXAMPLE



Averaged system  $\bar{X}(t)$  forms a meta "macro" Markov chain among fast-classes  $\mathcal{M}_k$  with propensities  $\bar{a}_{\beta}(\bar{X}(t); \beta, \tilde{\pi}(\alpha))$

# TTS SSA

- (1) Initialize  $x$  at macro-time  $T$ ;  $x \in \mathcal{M}_k$  some (unknown)  $k$
- (2) Simulate fast-only rxns  $Q^{(k)}(\alpha)$  until time-averages of observable  $f$  and slow propensities  $a_\beta$  relax to steady-state

$$\frac{1}{t} \int_0^t f\left(X^{(k)}(s)\right) ds \rightarrow \mathbb{E}_{\pi^{(k)}} \{f(X)\} = \bar{f}(k)$$

$$\frac{1}{t} \int_0^t a_\beta\left(X^{(k)}(s); \beta\right) ds \rightarrow \mathbb{E}_{\pi^{(k)}} \{a_\beta(X; \beta)\} = \bar{a}_\beta(k; \beta)$$

- (3) Observe terminal state  $\tilde{x}^{(k)} \sim \pi^{(k)}$ . Compute  $\bar{a}_{\beta_0} = \sum_{j=1}^{M_S} \bar{a}_{\beta_j}$
- (4) Sample time to next slow rxn:  $\Delta T \sim \text{Exp}(\bar{a}_{\beta_0})$
- (5) Sample next slow rxn to fire  $\beta^* \sim 1/\bar{a}_{\beta_0} [\bar{a}_{\beta_1}, \dots, \bar{a}_{\beta_{M_S}}]$
- (6) Update macro-time  $T \leftarrow T + \Delta T$  and move to next fast-class by taking  $x = \tilde{x}^{(k)} + \zeta_{\beta^*}$
- (7) Return to (1) until macro time horizon is reached

- ▶ At the end of a TTS simulation, one obtains terminal state  $X(T) = x \sim p_T^0 = p_T^0(X; \alpha, \beta)$ .  
 $p_T^0$  depends on fast-class stationary distributions  $\tilde{\pi}(\alpha)$  as well as distribution of macro-chain  $\bar{p}_T(\bar{X}; \alpha, \beta)$ .
- ▶ What is the error from exact probability distribution  $p_T^\varepsilon = p_T^\varepsilon(X^\varepsilon; \alpha/\varepsilon, \beta)$ ?

### Theorem 1 (TTS Distribution Error)

Let  $\tilde{\kappa} = \frac{-1}{2} \max\{\text{Re}(\lambda) : \lambda \text{ is a non-zero eigenvalue of a } Q^k\}$ . Then

$$\|p_T^\varepsilon - p_T^0\| \leq O(\varepsilon + \exp\{-\tilde{\kappa}T/\varepsilon\})$$

- ▶ Thus, as long as the macro time horizon  $T$  is greater than the relaxation time of the fast dynamics, the induced error of the TTS simulation is  $O(\varepsilon)$ .

Often, one is interested in steady state behavior as  $T \rightarrow \infty$ .  
As the stiffness parameter  $\varepsilon \rightarrow 0$ , this is given by

$$\lim_{\substack{T \rightarrow \infty \\ \varepsilon \rightarrow 0}} \mathbb{E}_{p_T^\varepsilon} \{f(X^\varepsilon(T))\} = \mathbb{E}_{\bar{\pi}} \{\bar{f}(\bar{X})\}$$

where  $\bar{\pi} = \bar{\pi}(\alpha, \beta)$  is the stationary distribution of the macro-chain associated with generator  $\bar{Q} = \tilde{\pi}(\alpha) \hat{Q}(\beta) \mathbf{1}$   
Write  $\pi = \bar{\pi} \tilde{\pi}$ , where  $\tilde{\pi} = \text{diag} [\pi^{(1)}, \dots, \pi^{(N_C)}]$ .

## Theorem 2 (TTS Expectation Error)

For sufficiently large  $T$ ,  $\|p_T^\varepsilon - \pi\| \leq O(\varepsilon)$  and  $\|\pi^\varepsilon - \pi\| \leq O(\varepsilon)$ ,  
where  $\pi^\varepsilon$  is the stationary distribution of  $\mathcal{M}$  corresponding to  $Q^\varepsilon$ .  
Thus for all bounded functions  $f$  on  $\mathcal{M}$ ,

$$|\mathbb{E}_{\pi^\varepsilon} \{f(X^\varepsilon)\} - \mathbb{E}_{\bar{\pi}} \{\bar{f}(\bar{X})\}| \leq \|f\|_\infty \|\pi^\varepsilon - \pi\| \leq O(\varepsilon)$$

Goal: estimate  $\frac{\partial}{\partial \theta_i} \mathbb{E}_{p^\varepsilon(t)} \{f(X^\varepsilon)\}$  for  $\theta_i \in \alpha$  or  $\theta_i \in \beta$

If sufficient conditions for existence of invariant distribution  $\bar{\pi}$  (e.g. the number of fast-classes  $\{\mathcal{M}_k\}_{k=1}^{N_C}$  is finite), then

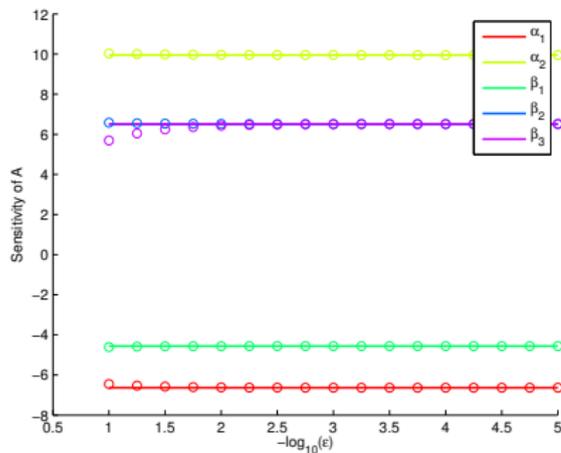
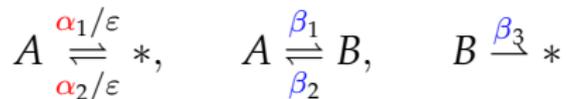
$$\mathbb{E}_{p^\varepsilon(t)} \{f(X^\varepsilon)\} = \mathbb{E}_{\bar{p}(t)} \{\bar{f}(\bar{X})\} + O(\varepsilon)$$

Does it follow that

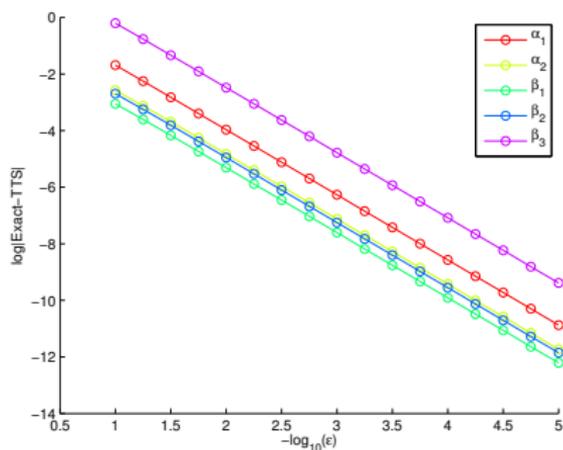
$$\frac{\partial}{\partial \theta_i} \mathbb{E}_{p^\varepsilon(t)} \{f(X^\varepsilon)\} = \frac{\partial}{\partial \theta_i} \mathbb{E}_{\bar{p}(t)} \{\bar{f}(\bar{X})\} + O(\varepsilon)?$$

[Gupta & Khammash 2014] :

$$\lim_{\varepsilon \rightarrow 0} \frac{\partial}{\partial \theta_i} \mathbb{E}_{p^\varepsilon(t)} \{f(X^\varepsilon)\} = \frac{\partial}{\partial \theta_i} \mathbb{E}_{\bar{p}(t)} \{\bar{f}(\bar{X})\}$$

ANALYTIC SENSITIVITIES OF  $|A|$ 

TTS sensitivity (solid lines) and exact system sensitivity (circles) by  $\varepsilon$ .



Log-Error of exact sensitivity from TTS sensitivity

# STIFF VS. RESCALED SENSITIVITIES

$$\boldsymbol{\theta}^\varepsilon = [\boldsymbol{\alpha}/\varepsilon, \boldsymbol{\beta}] = [\boldsymbol{\alpha}^\varepsilon, \boldsymbol{\beta}].$$

What is the difference between  $\partial_{\theta_1^\varepsilon} \mathbb{E}_{\pi^\varepsilon} \{f(X)\} = \partial_{\alpha_1/\varepsilon} \mathbb{E}_{\pi^\varepsilon} \{f(X)\}$   
vs.  $\partial_{\alpha_1} \mathbb{E}_{\pi^\varepsilon} \{f(X)\}$ ?

## Proposition 3

*Stiff vs. Rescaled Sensitivities*

$$\frac{\partial}{\partial \alpha_i^\varepsilon} \mathbb{E}_{\pi^\varepsilon} \{f(X; \boldsymbol{\theta}^\varepsilon)\} = \varepsilon \left[ \frac{\partial}{\partial \alpha_i} \mathbb{E}_{\pi^\varepsilon} \{f(X; \boldsymbol{\theta}^\varepsilon)\} \right]. \quad (1)$$

Pseudo-inverse representation of steady-state sensitivities :

### Lemma 4

*Let  $Q = Q(\boldsymbol{\theta})$  be the generator of an ergodic continuous-time Markov chain determined by a parameter vector  $\boldsymbol{\theta}$ . Let  $\pi = \pi(\boldsymbol{\theta})$  be the corresponding stationary distribution. Then we can represent the derivatives  $\partial_{\theta}\pi(\boldsymbol{\theta})$  as*

$$\frac{\partial \pi}{\partial \theta} = \pi \left( \frac{\partial Q}{\partial \theta} \right) Q^+ [\mathbf{1}\pi - I]. \quad (2)$$

*where  $Q^+$  is the pseudo-inverse of the generator  $Q$ .*

Possible combination with Finite State Projection algorithm?  
[Munsky & Khammash 2006]

Since the macro system  $\bar{X}$  is just another reaction network with propensities  $\bar{a}_\beta(\bar{x}; \beta, \tilde{\pi}(\alpha))$ , we can appeal to existing sensitivity analysis estimation methods.

3 categories: **Finite Difference**, **Likelihood Ratio / Girsanov Transform**, and **Pathwise Derivatives**

**Finite Differences** :

$$\frac{\partial}{\partial \theta_i} \mathbb{E}_{\theta^0} \{f(X(T))\} \approx 1/h \left[ \mathbb{E}_{\theta^0 + h \cdot e_i} \{f(X(T))\} - \mathbb{E}_{\theta^0} \{f(X(T))\} \right]$$

- ▶ Low Variance, but each parameter sensitivity must be estimated separately

## Likelihood Ratio / Girsanov Transform :

$$\begin{aligned} \frac{\partial}{\partial \theta_i} \mathbb{E}_{\theta^0} \{f(X(t))\} &= \int_{\Omega} f(X(t, \omega)) \frac{\frac{\partial}{\partial \theta_i} |_{\theta^0} P(d\omega, t; \theta)}{P(d\omega, t; \theta)} P(d\omega, t; \theta^0) \\ &= \mathbb{E}_{\theta^0} \{f(X(t)) W_{\theta_i}(t)\} \end{aligned}$$

Using the random-time-change representation:

$$X(t) = X(0) + R(t)\zeta = X(0) + \sum_{r=1}^M Y_r \left( \int_0^t a_r(X(s), \theta) ds \right) \zeta_r$$

$W_{\theta}(t)$  can be represented as

$$W_{\theta}(t) = \sum_{r=1}^M \int_0^t \frac{\frac{\partial a_r}{\partial \theta}(X(s^-), \theta^0)}{a_r(X(s^-), \theta^0)} dR_r(s) - \sum_{r=1}^M \int_0^t \frac{\partial a_r}{\partial \theta}(X(s^-), \theta^0) ds$$

$N_S$  = number of samples

**Likelihood Ratio Estimator:**

$$\widehat{\mathbf{LR}}(N_S, \theta_i) \triangleq \frac{1}{N_S} \sum_{n=1}^{N_S} [f(\widehat{x(T)})]_n [\widehat{W_{\theta_i}(T)}]_n.$$

$W_{\theta}(T)$  is a martingale with zero mean but variance which grows  $O(T)$

Makes LR inefficient for estimation at large time horizons

Can reduce variance by *centering*

**Centered Likelihood Ratio Estimator:**

$$\widehat{\mathbf{CLR}}(N_S, \theta_i) \triangleq \widehat{\mathbf{LR}}(N_S, \theta_i) - \frac{1}{N_S^2} \left\{ \sum_{n=1}^{N_S} [f(\widehat{x(T)})]_n \right\} \left\{ \sum_{n=1}^{N_S} [\widehat{W_{\theta_i}(T)}]_n \right\}$$

# STEADY-STATE SENSITIVITIES

How to estimate steady-state sensitivities  $\partial_{\theta_i} \mathbb{E}_{\pi(\theta)} \{f(X)\}$ ?

Terminal Distribution :  $\mathbb{E}_{p_T(\theta)} \{f(X(T))\} = \mathbb{E}_{\pi(\theta)} \{f(X)\} + O(e^{-\kappa T})$

[Gupta & Khammash 2014] :  $\lim_{T \rightarrow \infty} \partial_{\theta_i} \mathbb{E}_{p(T)(\theta)} \{f(X)\} = \partial_{\theta_i} \mathbb{E}_{\pi(\theta)} \{f(X)\}$ , so can estimate via LR/CLR with large time horizon  $f(X(T))W_{\theta_i}(T)$

Variance grows proportional to  $\mathbb{V}\text{ar}\{f(X(T))\}\mathbb{V}\text{ar}\{W_{\theta_i}(T)\} = O(T)$

Bias/Variance tradeoff in choice of  $T$

Instead of using the terminal distribution  $f(X(T))$  in LR estimator, could use ergodic average :  $1/T \int_0^T f(X(s))ds$ .

Larger bias  $O(1/T)$ , but diminishing variance  $O(1/T)$

Then  $\mathbb{V}\text{ar} \left\{ 1/T \int_0^T f(X(s))ds W_{\theta}(T) \right\} \approx O(1)$ , and should have  $O(1/T)$  bias

## CELR FOR STEADY-STATE SENSITIVITIES

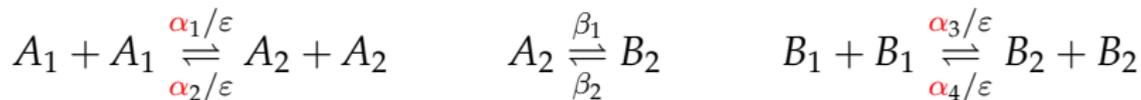
Thus we propose the **Ergodic Likelihood Ratio**

$$\widehat{\mathbf{ELR}}(N_S, \theta_i) \triangleq \frac{1}{N_S} \sum_{n=1}^{N_S} \frac{1}{T} \left[ \int_0^T \widehat{f(x(s))} ds \right]_n [\widehat{W_{\theta_i}(T)}]_n.$$

and again center it for improved efficiency to obtain the **Centered Ergodic Likelihood Ratio**

$$\begin{aligned} \widehat{\mathbf{CELR}}(N_S, \theta_i) &\triangleq \widehat{\mathbf{ELR}}(N_S, \theta_i) \\ &- \frac{1}{N_S^2} \left\{ \sum_{n=1}^{N_S} \frac{1}{T} \left[ \int_0^T \widehat{f(x(s))} ds \right]_n \right\} \left\{ \sum_{n=1}^{N_S} [\widehat{W_{\theta_i}(T)}]_n \right\} \end{aligned}$$

## CELR Estimator

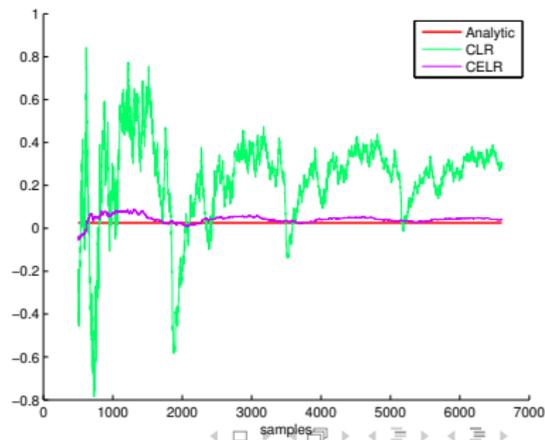
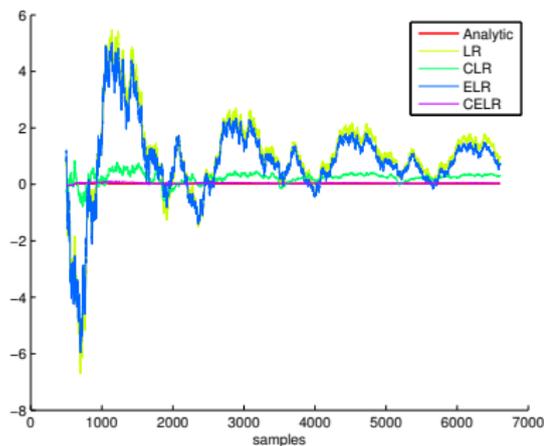


$x(0) = [10, 15, 15, 10]$ ,  $|\mathcal{M}| = 5850$ ,  $T = 100$

$\alpha = [1, 2, 3, 4]$ ,  $\beta = [1.5, 2.5]$ ,  $\varepsilon = 10^{-3}$ ,  $\theta = [\alpha/\varepsilon, \beta]$

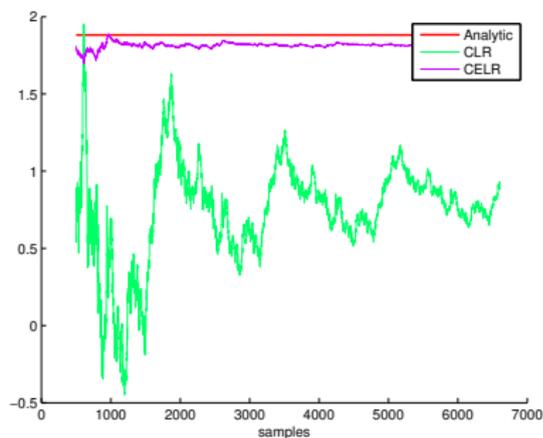
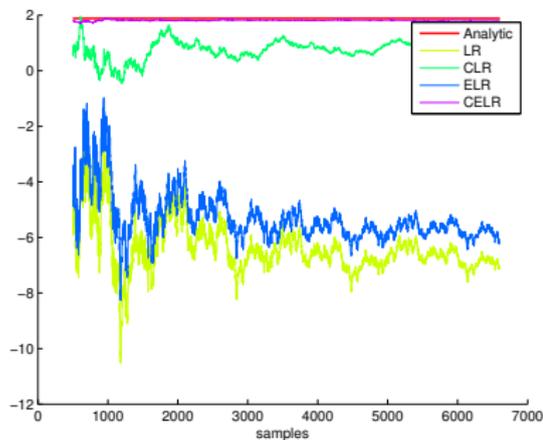
$f(X) = |A_2|$

$\partial_{\theta_1} \mathbb{E}_{\pi^\varepsilon} \{f(X)\} = \partial_{\alpha_1/\varepsilon} \mathbb{E}_{\pi^\varepsilon} \{f(X)\} :$



## CELR Estimator

$$\partial_{\theta_6} \mathbb{E}_{\pi^\varepsilon} \{f(X)\} = \partial_{\beta_2} \mathbb{E}_{\pi^\varepsilon} \{f(X)\} :$$



- ▶  $\lim_{\varepsilon \rightarrow 0} \frac{\partial}{\partial \theta_i} \mathbb{E}_{p^\varepsilon(t)} \{f(X^\varepsilon)\} = \frac{\partial}{\partial \theta_i} \mathbb{E}_{\bar{p}(t)} \{\bar{f}(\bar{X}; \boldsymbol{\theta})\}$
- ▶  $\bar{X}(t)$  is just another reaction network with propensities  $\bar{a}_{\beta_j}(\bar{x}; \boldsymbol{\theta})$  and observable  $\bar{f}(\bar{x}; \boldsymbol{\theta})$
- ▶ Thus the macro-sensitivities can be represented by

$$\frac{\partial}{\partial \theta_i} \mathbb{E}_{\bar{p}_T(\boldsymbol{\theta})} \{\bar{f}(\bar{X}(T); \boldsymbol{\theta})\} = \mathbb{E}_{\bar{p}_T(\boldsymbol{\theta})} \left\{ \frac{\partial}{\partial \theta_i} \bar{f}(\bar{X}(T); \boldsymbol{\theta}) + \bar{f}(\bar{X}; \boldsymbol{\theta}) \bar{W}_{\theta_i}(T) \right\}$$

$$\bar{W}_{\theta_i} = \sum_{r=1}^{M_s} \int_0^T \frac{\frac{\partial}{\partial \theta_i} \bar{a}_{\beta_r}(\bar{X}(s); \boldsymbol{\theta})}{\bar{a}_{\beta_r}(\bar{X}(s); \boldsymbol{\theta})} dR_{\beta_r}(s) - \sum_{r=1}^{M_s} \int_0^T \frac{\partial}{\partial \theta_i} \bar{a}_{\beta_r}(\bar{X}(s); \boldsymbol{\theta}) ds.$$

- ▶ Sensitivity w.r.t. slow parameters  $\theta_i = \beta_i$  straightforward, since

$$\frac{\partial}{\partial \beta_i} \bar{f}(\bar{X}; \boldsymbol{\alpha}) = 0 \quad \frac{\partial}{\partial \beta_i} \bar{a}_{\beta_r}(\bar{X}; \boldsymbol{\alpha}) = \bar{b}_{\beta_r}(\bar{X}; \boldsymbol{\alpha}) \cdot \delta_{i,r}$$

For  $\theta_i = \alpha_i$ , the situation is more complicated.

$$\frac{\partial}{\partial \alpha_i} \mathbb{E}_{\bar{p}_T(\theta)} \{ \bar{f}(\bar{X}(T); \alpha) \} = \mathbb{E}_{\bar{p}_T(\theta)} \left\{ \frac{\partial \bar{f}}{\partial \alpha_i}(\bar{X}; \alpha) + \bar{f}(\bar{X}; \alpha) \bar{W}_{\alpha_i}(T) \right\}$$

- ▶  $\bar{f}(\bar{X}; \alpha)$  and  $\bar{a}_{\beta_r}(\bar{X}; \alpha)$  depend on  $\alpha_i$  only indirectly through underlying fast-class steady-state measure  $\pi^{(\bar{X})}(\alpha)$ .
- ▶ However, we may estimate  $\partial \alpha_i \bar{f}(\bar{X}; \alpha)$  and  $\partial \alpha_i \bar{a}_{\beta_r}(\bar{X}; \alpha)$  using micro-sensitivity estimates inside each fast-class  $\bar{X}$ .

$$\frac{\partial}{\partial \alpha_i} \bar{f}(\bar{X}; \theta) \approx \mathbb{E}_{\tilde{p}_t^{(\bar{X})}} \left\{ \frac{1}{t} \left[ \int_0^t f(\tilde{X}(s)) ds \right] \tilde{W}_{\alpha_i}(t) \right\}$$

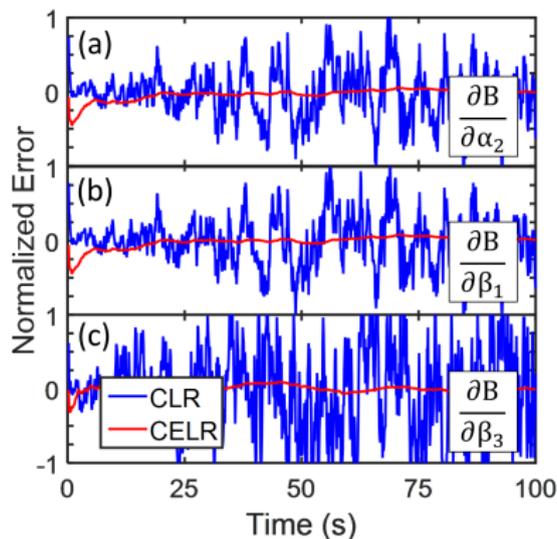
$$\frac{\partial}{\partial \alpha_i} \bar{a}_{\beta_r}(\bar{X}; \theta) \approx \mathbb{E}_{\tilde{p}_t^{(\bar{X})}} \left\{ \frac{1}{t} \left[ \int_0^t a_{\beta_r}(\tilde{X}(s)) ds \right] \tilde{W}_{\alpha_i}(t) \right\},$$

$$A \begin{matrix} \xrightarrow{\alpha_1/\varepsilon} \\ \xleftarrow{\alpha_2/\varepsilon} \end{matrix} *, \quad A \begin{matrix} \xrightarrow{\beta_1} \\ \xleftarrow{\beta_2} \end{matrix} B, \quad B \xrightarrow{\beta_3} *$$

$$(A_0, B_0, *_0) = (30, 60, 10)$$

$$(\alpha_1^\varepsilon, \alpha_2^\varepsilon, \beta_1, \beta_2, \beta_3) = (1/\varepsilon, 1.5/\varepsilon, 2, 1, 0.4)$$

1000 TTS samples computed to time horizon  $T = 100$



Percent Error				
	CLR	CELR	CLR	CELR
	$t = 1.3s$		$t = 100s$	
$\alpha_1$	-1.4	-40.0	-83.5	0.0
$\alpha_2$	-3.0	-40.3	-63.0	-1.1
$\beta_1$	0.2	-39.1	-64.1	1.2
$\beta_2$	-1.1	-16.7	-22.6	-2.8
$\beta_3$	-1.4	-23.3	20.6	-0.1
Half-length of 95% Confidence Interval				
$\alpha_1$	12	8	99	11
$\alpha_2$	12	7	94	11
$\beta_1$	12	7	91	11
$\beta_2$	13	10	108	12
$\beta_3$	21	14	171	18

## Pathwise Derivatives :

Application of I.P.A. of discrete-event systems to rxn networks in the **path space** [Sheppard et al. 2012]

$$X(t, \boldsymbol{\theta}, \omega) = X(0, \boldsymbol{\theta}, \omega) + \sum_{r=1}^M \zeta_r Y_r(S_r(t, \boldsymbol{\theta}, \omega), \omega)$$

**Internal times** :  $S_r(t, ; \omega, \boldsymbol{\theta}) = \int_0^t a_r(X(s; \omega, \boldsymbol{\theta}); \boldsymbol{\theta}) ;$

- ▶ For a fixed  $\omega$ , the event-times of the Poisson processes  $Y_r(\cdot, \omega)$  are fixed
- ▶ When the (increasing, piecewise-linear) internal times  $S_r(t, \boldsymbol{\theta}, \omega)$  hit the event-times of  $Y_r(\cdot, \omega)$ , reaction channel  $r$  will fire
- ▶  $I_+^r(t) = \inf \{u \geq S_r(t, \boldsymbol{\theta}) : Y_r(u) = Y_r(S_r(t, \boldsymbol{\theta}) + 1)\} =$  the internal time for the next event of  $Y_r(\cdot)$  after the current internal time  $S_r(t, \boldsymbol{\theta})$

Viewing parameter-dependence on trajectory  $X_{\theta}(t, \omega)$ , can compute derivatives of jump (rxn) times  $T_n = T_n(\theta, \omega)$  in terms of derivatives of internal times for each rxn

$$T_{n+1} = T_n + \min_{r=1, \dots, M} \left\{ \frac{I_+^r(T_n) - S_r(T_n)}{a_r(X(T_n))} \right\}$$

$\partial\theta_i T_n(\theta; \omega)$  and  $\partial\theta_i S_r(T_n(\theta; \omega), \theta; \omega)$  can be solved recursively from  $T_0 = 0$ .

$\partial\theta_i T_n(\theta, \omega)$  can be computed in terms of :

$$a_r(X_n; \theta) \text{ and } \partial\theta_i a_r(X_n; \theta)$$

For path-space differentiation,

$\partial\theta_i \mathbb{E} \{f(X(T))\} \neq \mathbb{E} \{\partial\theta_i f(X(T))\}$ , since the RHS  $\equiv 0$

Instead, must **smooth** the RHS by instead using an integral of the form  $L(\theta) = \int_A^B F(\theta, X_{\theta}(s)) ds$

Several choices can be considered

- ▶ **RPD** : Regularized Pathwise Derivative [Sheppard et al. 2012]

$$L_{\text{RPD}}(\boldsymbol{\theta}) = \frac{1}{2w} \int_{T-w}^{T+w} f(X_{\boldsymbol{\theta}}(s)) ds \approx f(X_{\boldsymbol{\theta}}(T))$$

Bias-Variance trade-off in  $w$

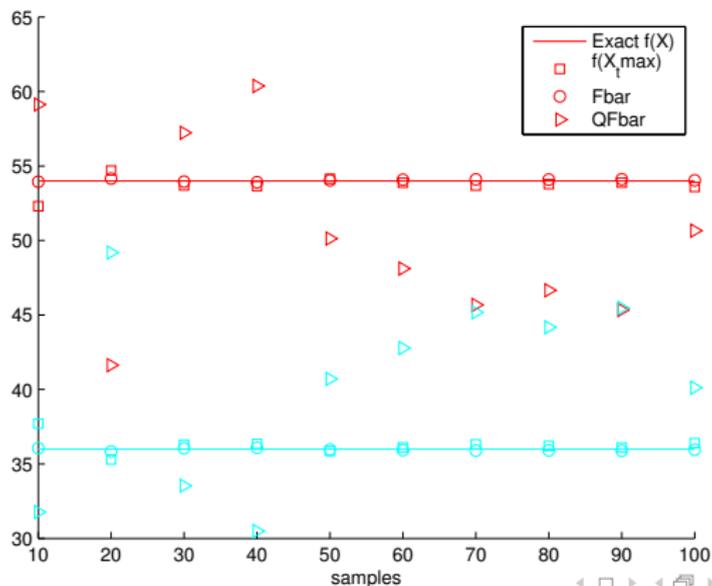
- ▶ **GS** : Generator Smoothing (Dynkin Formula) [Wolf and Anderson 2015]

$$L_{\text{GS}}(\boldsymbol{\theta}) = f(X_{\boldsymbol{\theta}}(0)) + \int_0^T (Q(\boldsymbol{\theta})f)(X_{\boldsymbol{\theta}}(s)) ds$$

Higher variance, but unbiased  $\mathbb{E}\{L_{\text{GS}}(\boldsymbol{\theta})\} = \mathbb{E}\{f(X_{\boldsymbol{\theta}}(T))\}$

For steady-state estimation, the ergodic average will have lower variance than the distribution at terminal time  $T$ . Thus we use the **Ergodic Pathwise Derivative**:

$$L_{\text{EPD}}(\theta) = \frac{1}{T} \int_0^T f(X_{\theta}(s)) ds = \mathbb{E}_{\pi(\theta)}\{f(X)\} + O(1/T)$$



For any of the above functionals  $L(\boldsymbol{\theta})$ , one can compute the pathwise derivative via

$$\begin{aligned} \frac{\partial}{\partial \theta_i} \int_0^t f(X^\theta(s); \boldsymbol{\theta}) ds &= \sum_{l=0}^{N(t)-1} \left[ f(\hat{X}_l^\theta; \boldsymbol{\theta}) \left( \frac{\partial}{\partial \theta_i} \Delta_l^\theta \right) + \left( \frac{\partial}{\partial \theta_i} f(\hat{X}_l^\theta; \boldsymbol{\theta}) \right) \Delta_l^\theta \right] \\ &+ \left[ \frac{\partial}{\partial \theta_i} f(\hat{X}_{N(t)}^\theta; \boldsymbol{\theta}) \right] (t - T_{N(t)}^\theta) - f(\hat{X}_{N(t)}^\theta, \boldsymbol{\theta}) \left[ \frac{\partial}{\partial \theta_i} T_{N(t)}^\theta \right], \end{aligned}$$

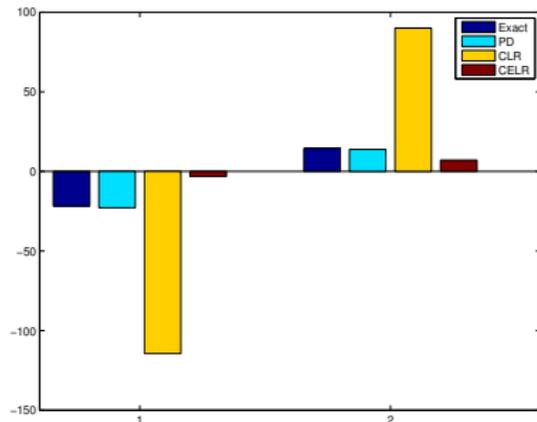
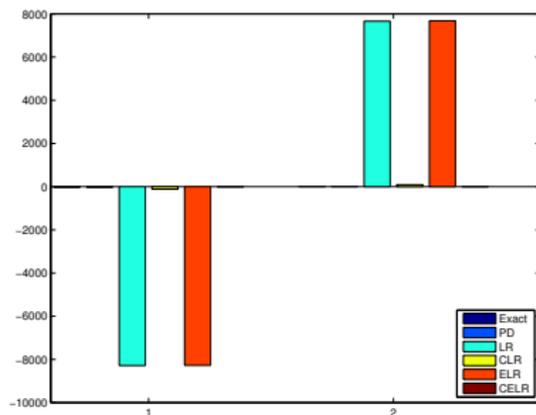
where  $\Delta_l = T_{l+1} - T_l$ .

$$\frac{\partial}{\partial \theta_i} \mathbb{E}_\theta \{f(X^\theta(T))\} = \mathbb{E}_\theta \left\{ \frac{\partial}{\partial \theta_i} L(\boldsymbol{\theta}) \right\} + \text{bias}_{\text{smoothing}} + \text{bias}_{PD}$$

$\text{bias}_{\text{smoothing}} = \mathbb{E}_\theta \{L(\boldsymbol{\theta})\} - \mathbb{E}_\theta \{f(X(T))\}$ , can be controlled

The pathwise derivative bias,  $\text{bias}_{PD} = \partial \theta_i \mathbb{E}_\theta \{L(\boldsymbol{\theta})\} - \mathbb{E}_\theta \{\partial \theta_i L(\boldsymbol{\theta})\}$ , can be more problematic.

**Example: Single-Scale with only fast reactions**  $A \xrightarrow{\alpha_1} * \xrightarrow{\alpha_2} A$ ,  
 $f(X(t)) = |A(t)|$ , time horizon  $t = 37.5$  seconds, small sample size: 100 samples



Comparison of (ergodic) Pathwise Derivative against Girsanov Transform estimates.

When applicable, achieves very accurate estimates with a small number of samples; can be used for estimation of micro-sensitivities

$$\partial_{\alpha_i} \bar{a}_\beta(\bar{X})$$

$$\lim_{\varepsilon \rightarrow 0} \frac{\partial}{\partial \theta_i} \mathbb{E}_{p_T^\varepsilon(\theta)} \{f(X^\varepsilon(T))\} \frac{\partial}{\partial \theta_i} = \mathbb{E}_{\bar{p}_T(\theta)} \left\{ \bar{f}(\bar{X}(T)) \right\}$$

The process  $\bar{X}(T)$  is simply another reaction network with reactions

- ▶ propensities  $\bar{a}_{\beta_r}(\bar{X}; \theta)$
- ▶ jump-times  $\bar{T}_l^\theta$
- ▶ internal times  $\bar{S}_{\beta_r}^\theta(\bar{t}) = \int_0^{\bar{t}} \bar{a}_{\beta_r}(\bar{X}(s)) ds$

Again, if one can compute  $\bar{a}(\bar{X}, \theta)$  and  $\partial_{\theta_i} \bar{a}(\bar{X}, \theta)$ , then one can compute a multiscale pathwise derivative

Can combine Multiscale PD with Multiscale LR for a Multiscale Hybrid PD similar to [Wolf & Anderson 2015].

## SUMMARY

- ▶ TTS framework for multiscale stochastic reaction networks
- ▶ Rate of convergence of probability measures
- ▶ Pseudo-inverse representation of stationary sensitivities
- ▶ CELR estimator for stationary sensitivities
- ▶ EPD estimator for stationary sensitivities
- ▶ Multiscale LR and PD

A. Hashemi, M. Nunez, P. Plechac, and D. G. Vlachos, “Stochastic Averaging and Sensitivity Analysis for Two Scale Reaction Networks,” arXiv:1509.03802, Sep. 2015.