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Generative Models and Stochastic Algorithms for Population Average Estimation and Image Analysis

Stéphanie Allassonnière

CIS, JHU

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Context :	Computational Anato	omy				

Context and motivations :

- * Describing shapes
- * Shape matching : many elaborated registrations theories
- * Defining and infering population average (atlas)
 - Statistical estimation in presence of unobserved variables

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- Proper statistical model = generative
- Consistency issues addressed
- * Discrimination/Classification

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Context : Computational Anatomy						

Outline

Three generative statistical models and stochastic algorithms

- 1 Bayesian Mixed Effect (BME) gray level Template
 - 1.1 Mathematical framework for deformable models
 - 1.2 Past approaches to compute a population average
 - 1.3 Generative statistical models
 - 1.4 Statistical estimation of the model parameters
 - 1.5 Experiments on USPS database and 2D medical images

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- 2 Bayesian Mixed Effect DTI Template
- 3 Noisy ICA

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Outline

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- 2 Bayesian Mixed Effect DTI Template
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Warping a	and past approaches					

Linearized deformations :

- Non rigid deformations
- Let ψ be the deformation from I_0 to I_1 and v a vector field :

$$\psi = \mathbf{Id} + \mathbf{v}$$

• Resulting deformed image given by :

$$I_1(x) \simeq I_0(x - v(x))$$

Advantages :

- Easy to use in computations
- Relevant for certain classes of shapes

Drawbacks :

Invertibility not guaranteed : Overlap → Shape topology may change !

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Warping a	nd past approaches					

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Some solutions previously given :

Only for the template!

• Using one of the data images y_1^n .

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Warping a	nd past approaches					

Only for the template!

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Warping a	and past approaches					

Only for the template!

• Using one of the data images y_1^n .

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Warping a	and past approaches					

Only for the template!

• Procrustes' mean :

$$(\hat{l}_0, \hat{v}_1, \dots, \hat{v}_n) = \arg\min_{l_0, v_1, \dots, v_n} \sum_{i=1}^n \left(\frac{1}{2} \|v_i\|_V^2 + \frac{1}{2\sigma^2} |y_i \circ \phi^{v_i} - l_0|^2\right)$$

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Warping a	and past approaches					

Only for the template!

• Procrustes' mean :

$$(\hat{l}_0, \hat{v}_1, \dots, \hat{v}_n) = \arg\min_{l_0, v_1, \dots, v_n} \sum_{i=1}^n \left(\frac{1}{2} ||v_i||_V^2 + \frac{1}{2\sigma^2} |y_i \circ \phi^{v_i} - l_0|^2\right)$$

• A statistical interpretation : (Glasbey & Mardia)

$$y_i(x + v_i(x)) = I_0(x) + \epsilon_x, \ \epsilon_x \sim \mathcal{N}(0, \sigma^2), \ x \in \Lambda$$

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Problems :

- Needs interpolation
- Unobserved warping variables
- Not a generative statistical model

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Warping a	nd past approaches					

Issues :

Describing databases as some i.i.d. sample of some parametric generative statistical model.

Model parameters = template + deformation law (define the space V) + noise

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Learning the parameters to avoid the previous problems :

- An intrinsic template I₀,
- A weighting term σ^2 ,
- A global geometric behavior in the class.

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Statistical	Generative Model					

Generative Statistical Model :

 \rightarrow Deformable template framework :

$$y_{j,k}^{i} = I_{0}(r_{j,k} - v_{i}(r_{j,k})) + \sigma \varepsilon_{j,k}$$

Conditions chosen for the template and the deformation fields : <u>Parametric Model of Splines</u> : let $(p)_1^{k_p}$ and $(g)_1^{k_g}$ be two sets of control points (fixed, uniformly distributed) :

$$I_0(x) = K_p \alpha(x) = \sum_{k=1}^{k_p} K_p(x, p_k) \alpha(k) ,$$
$$v_\beta(x) = (K_g \beta)(x) = \sum_{k=1}^{k_g} K_g(x, g_k) \beta(k).$$

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Statistical	Generative Model					

What to learn?

Photometry : $\begin{cases} \alpha : \text{ to code the template} \\ \sigma^2 : \text{ the noise variance} \end{cases}$ Geometry : β_1^n : to code each deformation

BUT : does not give the geometrical behavior in the training set. \implies Introduce a prior on β :

$\beta_i \sim \nu(d\beta)$

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Parameters of $\nu = \text{parameters to learn.}$ Deformations $(\beta_i)_{1 \le i \le n} = \text{random hidden variables}$

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Statistical	Generative Model					

Generative Model :

One component per class :

$$\begin{cases} (\Gamma_{g}, \theta_{p}) \sim \nu_{g} \otimes \nu_{p} \text{ with } \theta_{p} = (\alpha, \sigma^{2}) \\ \beta_{1}^{n} \sim \otimes_{i=1}^{n} \mathcal{N}(0, \Gamma_{g}) \mid \Gamma_{g} \\ \gamma_{1}^{n} \sim \otimes_{i=1}^{n} \mathcal{N}(\mathsf{v}_{\beta_{i}} I_{\alpha}, \sigma^{2} \mathsf{Id}_{\Lambda}) \mid \beta_{1}^{n}, \theta_{p} \end{cases}$$

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where $\nu_g(d\Gamma_g), \nu_p(d\sigma^2, d\alpha)$ are prior laws on the parameters. Remark : big structures to learn even in the case of small size training set

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Statistical	Generative Model					

Generative Model (2) :

General case : mixtures of deformable templates (τ_m components per class) Hidden random variables : $(\beta_i)_{1 \le i \le n}$ and the image labels $(\tau_i)_{1 \le i \le n}$.

$$\begin{cases} \rho \sim \nu_{\rho} \\\\ \theta = (\theta_{g}^{\tau}, \theta_{\rho}^{\tau})_{1 \leq \tau \leq \tau_{m}} \sim \otimes_{\tau=1}^{\tau_{m}} (\nu_{g} \otimes \nu_{\rho}) \\\\ \tau_{1}^{n} \sim \otimes_{i=1}^{n} \sum_{\tau=1}^{\tau_{m}} \rho_{\tau} \delta_{\tau} \mid \rho \\\\ \beta_{1}^{n} \sim \otimes_{i=1}^{n} \mathcal{N}(0, \Gamma_{g}^{\tau_{i}}) \mid \theta, \ \tau_{1}^{n} \\\\ y_{1}^{n} \sim \otimes_{i=1}^{n} \mathcal{N}(v_{\beta_{i}} I_{\alpha_{\tau_{i}}}, \sigma_{\tau_{i}}^{2} Id_{\Lambda}) \mid \beta_{1}^{n}, \ \theta, \ \tau_{1}^{n} \end{cases}$$

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Statistical	Generative Model					

Generative Statistical Model :



FIG.: Mixed effect structure for our BME-template

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Estimation	n					

How to learn the parameters? the MAP Estimator :

Parameters θ are estimated by maximum posterior likelihood :

 $\hat{\theta} = \arg \max P(\theta|y)$

where

$$\begin{split} \theta \in \Theta &= \{ \ (\alpha, \sigma^2, \Gamma_g) | \alpha \in \mathbb{R}^{k_p}, \ \sigma^2 > 0, \ \Gamma_g \in \mathcal{Sym}^+_{2k_g, *}(\mathbb{R}) \ \}. \\ \mathcal{Sym}^+_{2k_g, *}(\mathbb{R}) \text{ is the set of positive definite symmetric matrices.} \end{split}$$

Let $\Theta_* = \{ \theta_* \in \Theta \mid E_P(\log q(y|\theta_*)) = \sup_{\theta \in \Theta} E_P(\log q(y|\theta)) \}$ where *P* denotes the distribution governing the observations.

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Estimation	n					

How to do in practice?

Since β_1^n are unobserved variables, a natural approach to reach the MAP estimator is the **EM algorithm**.

Iteration / of the algorithm :

E Step : Compute the posterior law on β_i , i = 1, ..., n.

M Step : Parameter update :

$$heta_{l+1} = \arg \max_{\theta} E\left[\log q(\theta, \beta_1^n, y_1^n) | y_1^n, \theta_l
ight].$$

BUT : the E step is not tractable !

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Estimatio	n					

Details of the maximization step :

Geometry :

$$\theta_{g,l+1} = \Gamma_{g,l+1} = \frac{1}{n+a_g} (n[\beta\beta^t]_l + a_g\Sigma_g).$$

where

$$[\beta\beta^t]_l = \frac{1}{n} \sum_{i=1}^n \int \beta\beta^t \nu_{l,i}(\beta) d\beta,$$

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is the empirical covariance matrix with respect to the posterior density function.

 \rightarrow Importance of the prior !

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Solution proposed : Stochastic version of the EM algorithm :

Idea : Couple SAEM with MCMC procedure (Delyon, Lavielle, Moulines and Kuhn, Lavielle) :

One component case : Iteration $I \rightarrow I + 1$ of the algorithm :

- Simulation step : $\beta^{l+1} \sim \Pi_{\theta_l}(\beta^l, \cdot)$ where $\Pi_{\theta_l}(\beta^l, \cdot)$ is a transition probability of a convergent Markov Chain having the posterior distribution as stationary distribution,
- Stochastic approximation :

 $Q_{l+1}(\theta) = Q_l(\theta) + \Delta_l[\log q(y, \beta^{l+1}, \theta) - Q_l(\theta)]$ where (Δ_l) is a decreasing sequence of positive step-sizes.

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• Maximization step : $\theta_{l+1} = \arg \max Q_{l+1}(\theta)$

[*] $\Pi_{\theta_l}(\beta^l, \cdot)$ given by an hybrid Gibbs sampler

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Estimation	1					

Stochastic version of the EM algorithm (2) :

Since our model is an Exponential Model,

$$q(y, \beta^{l+1}, \theta) = \exp \left\{-\psi(\theta) + \langle S(y, \beta), \phi(\theta) \rangle \right\}$$

the stochastic approximation can be done on the sufficient statistics S so that the algorithm is done via :

$$s_{l+1} = s_l + \Delta_l \left(S(y, \beta^{l+1}) - s_l \right)$$

Let $L(s,\theta) = -\psi(\theta) + \langle s, \phi(\theta) \rangle$, $I(\theta) = \log q(y,\theta)$ and $\hat{\theta}(s) = \arg \max_{s} L(s,\theta(s))$ then

$$\theta_{k+1} = \hat{\theta}(s_{l+1})$$
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Estimation	ı					

Stochastic approximation with truncation on random boundaries :

Set
$$\kappa_0 = 0$$
, $s_0 \in \mathcal{K}_0$ and $\beta_0 \in K$.
 $\forall k \ge 1 \text{ compute } \bar{s} = s_{k-1} + \Delta_{k-1}(S(\bar{\beta}) - s_{k-1})$
where $\bar{\beta}$ is sampled from a transition kernel
 $\Pi_{\theta_{k-1}}(\beta_{k-1}, \cdot)$.
If $\bar{s} \in \mathcal{K}_{\kappa_{k-1}}$ and $|\bar{s} - s_{k-1}| \le \varepsilon_{k-1}$
set $(s_k, \beta_k) = (\bar{s}, \bar{\beta})$ and $\kappa_k = \kappa_{k-1}$,
else set $(s_k, \beta_k) = (\bar{s}, \bar{\beta}) \in \mathcal{K}_0 \times K$ and $\kappa_k = \kappa_{k-1} + 1$, where
 $(\tilde{s}, \tilde{\beta})$ can be chosen through different ways.
 $\theta_k = \arg \max_{\theta} L(s_k, \theta)$

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Multicomponent Model	

Stochastic version of the EM algorithm for the multicomponent model :

Intuitive generalization :

- problem of "trapping states".
- Image analysis interpretation : each iteration tries to deform the data so it is closer to its current component and will not tend to move toward another one

 \rightarrow high dimensional hidden variable β

Solution proposed : Consider another simulation method based on the Gibbs sampler for the deformation and on another law for the class of a given image.

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Multicomp	oonent Model					

The new algorithm :

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Transition step I
ightarrow I+1 using a hybrid Gibbs sampler on (eta, au) :

• for each τ : Run N_l times the hybrid Gibbs Sampler on β given τ .

$$\hat{\beta}_{\tau}^{(l+1)} = \Pi^{N_l}(\beta|\tau)$$

• draw $\tau^{(l+1)}$ through the discrete law with weights :

$$p_{N_l}(\tau) \propto \left(\frac{1}{N} \sum_{i_{mc}=1}^{N_l} \left[\frac{f(\hat{\beta}_{\tau,(i_{mc})})}{q(y,\hat{\beta}_{\tau,(i_{mc})},\tau|\theta,\rho)}\right]\right)^{-1}$$
$$\beta^{(l+1)} = \hat{\beta}_{\tau^{(l+1)},(N_l)}$$

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Multicomponent Model							

Theoretical Results :

With these models and algorithms we have proved some important asymptotic results :

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- Consistency of the MAP estimator
- Convergence of both stochastic algorithms

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Experimen	ts					

MCMC-SAEM :

Template estimation :



FIG.: Left : one component prototype. Right : 2 component prototypes

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Experimer	nts					

MCMC-SAEM :

The Geometric Distribution :

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FIG.: Left : 20 examples of the training set. Right : 20 examples drawn from the prior geometric distribution.

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MCMC-SAEM : Geometry :

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In presence of noise :

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3333333	
5555555	555555555555555555555555555555555555555
66666666	6666666666666
1717777	177117777777
999999999	1901091010900 9901091010909

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Medical Images : Splenium of the Corpus Callossum 47 images of the corpus callosum (and part of the cerebellum)





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FIG.: Left : Gray level mean of the 47 images. Right : template estimated with the stochastic algorithm.

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Experime	ents					

Medical Images : Splenium of the Corpus Callossum 47 images of the corpus callosum (and part of the cerebellum) clustered into 2 components by the multicomponent model.





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FIG.: Results from the estimation with the stochastic algorithm. Left : component 1. Right : component 2.

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Robustness of the algorithm :

Same hyper-parameters as the previous gray level images





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- 2 Bayesian Mixed Effect DTI Template
- 3 Noisy ICA

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BME-DTI	Template					

Bayesian Mixed Effect DTI Template :

Goals of our approach :

* Estimate a Template of Diffusion Tensor Image on a given region of the anatomy

* Just use the Diffusion Weight Images [DWIs] (real vector corresponding to the response to some different gradients)

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BME-D	TI Template					

Previous approach

Subjects	Least square approx.			mean	
		Min of energy	/		
	DWI G1				
Subject 1		\rightarrow	DTI 1	\rightarrow	
	DWI Gm				
	DWI G1				DTI template
Subject 2		\rightarrow	DTI 2	\rightarrow	Diritemplate
	DWI Gm				
	DWI G1				
Subject <i>n</i>		\rightarrow	DTI n	\rightarrow	
	DWI Gm				

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BME-D	TI Template					

Our approach

Subjects	observed	(hidden)	Max likelihood
	DWI G1		\rightarrow
Subject 1		(DTI 1)	\rightarrow
	DWI Gm		\rightarrow
	DWI G1		\rightarrow
Subject 2		(DTI 2)	\rightarrow
	DWI Gm		\rightarrow
	DWI G1		\rightarrow
Subject <i>n</i>		(DTI <i>n</i>)	\rightarrow
	DWI Gm		\rightarrow

DTI template

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BME-DTI	Template					

Results on synthetic data :

• 2 different template tensors

 $FA_1 = 0.6791, FA_2 = 0.6918$

- $ADC_1 = 0.5038$, $ADC_2 = 0.4329$
- 50 random samples with 15 subjects each

	LS	FAM-EM.	SAEM	LS	FAM-EM	SAEM
bias	0.1960	0.7386	0.1409	0.2095	0.5655	0.1269
var	0.7033	0.3489	0.7351	0.4853	0.2648	0.4969
mse	0.8993	1.0875	0.8760	0.6949	0.8303	0.6238
FA	0.6534	0.6193	0.6683	0.6683	0.6332	0.6814
ADC	0.5752	0.5627	0.5460	0.5044	0.4932	0.4747
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BME-DTI Template





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Outline

Three generative statistical models and stochastic algorithms

- 1 Bayesian Mixed Effect (BME) gray level Template
 - 1.1 Mathematical framework for deformable models
 - 1.2 Past approaches to compute a population average
 - 1.3 Generative statistical models
 - 1.4 Statistical estimation of the model parameters
 - 1.5 Experiments on USPS database and 2D medical images

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- 2 Bayesian Mixed Effect DTI Template
- 3 Noisy ICA

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Noisy ICA Model

- Observations : X_1^n such as $X_i = AY_i + \sigma \varepsilon$,
- A : source matrix
- σ^2 : variance of the Gaussian noise
- Y_1^n hidden variables.
- Model :

$$\begin{cases} Y_{1,1}^{n,p} \sim \otimes_{i=1}^{n} \otimes_{j=1}^{p} \nu_{\eta} \mid \eta, \\ X_{1}^{n} \sim \otimes_{i=1}^{N} \mathcal{N}(AY_{i}, \sigma^{2}Id) \mid A, \sigma^{2}, Y_{1}^{N}. \end{cases}$$

• Various choice of the distribution u_η

Same MCMC-SAEM algorithm to treat this estimation problem.

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Experime	Experiments : 101 subjects, 20 J.C.								



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Conclusion

• Generative statistical model = proper statistical framework for designing and inferring population average

• Stochastic algorithm of multiple uses even in difficult conditions

Thank you!

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MCMC-SAEM :

The Geometric Distribution : Between 2 classes :

Between 2 components in the same class :

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