

**Permutation test for classification
&
Risk bounds for mixture of densities**

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The learning problem

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Empirical risk minimization

$$\mathcal{A} : f_S \in \arg \min_{f \in \mathcal{H}} \frac{1}{n} \sum_{i=1}^n V(f(x_i), y_i).$$

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- Function output: $f_S \in \arg \min_{f \in \mathcal{H}} R_{\text{emp}}[f]$

Approximation and estimation errors

$$\mathbb{R}[f_S] - \mathbb{R}[f_0] = \mathbb{R}[f_S] - \mathbb{R}[f_{\mathcal{H}}] + \mathbb{R}[f_{\mathcal{H}}] - \mathbb{R}[f_0]$$

generalization error = estimation error + approximation error

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As $|\mathcal{H}|$ increases: **approximation error** decreases
estimation error increases.

Analysis of the tradeoff

1. Niyogi and Girosi: Approximation-estimation analysis for Radial Basis Functions

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4. Smale and Zhou: Estimating the approximation error for RKHS

Two problems

1. Permutation tests for classification: uses label permutations to compute a **bias variance** tradeoff for classification.
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2. Risk bounds for mixture of densities: **approximation** and **estimation** bounds for mixture of densities models.
A. Rakhlin, D. Panchenko, and S. Mukherjee

Permutation tests for classification

The permutation procedure described here is used extensively in **gene expression analysis** and **image based clinical studies**.

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- **Image-based clinical studies**: detect neuroanatomical changes induced by diseases and predict disease development.
- **Gene expression analysis**: classify tissue morphology, lineage, treatment outcome, or drug sensitivity using DNA microarray data.

The practical problem

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We compute a statistic $\mathcal{T}[S]$:

Training error $\frac{1}{n} \sum_{i=1}^n V(f_S(\mathbf{x}_i), \mathbf{y}_i)$

Leave-one-out error $\frac{1}{n} \sum_{i=1}^n V(f_{S^i}(\mathbf{x}_i), \mathbf{y}_i).$

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Can we trust $\mathcal{T}[S]$?

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- Repeat $m = 1, \dots, M$ times
 - ★ permute the labels: $\pi_m(S)$,
 - ★ $t_m = \mathcal{T}[\pi_m(S)]$
- construct an empirical cumulative distribution

$$\hat{\mathbb{P}}(T \leq t) = \frac{1}{M} \sum_{m=1}^M \Theta(t - t_m),$$

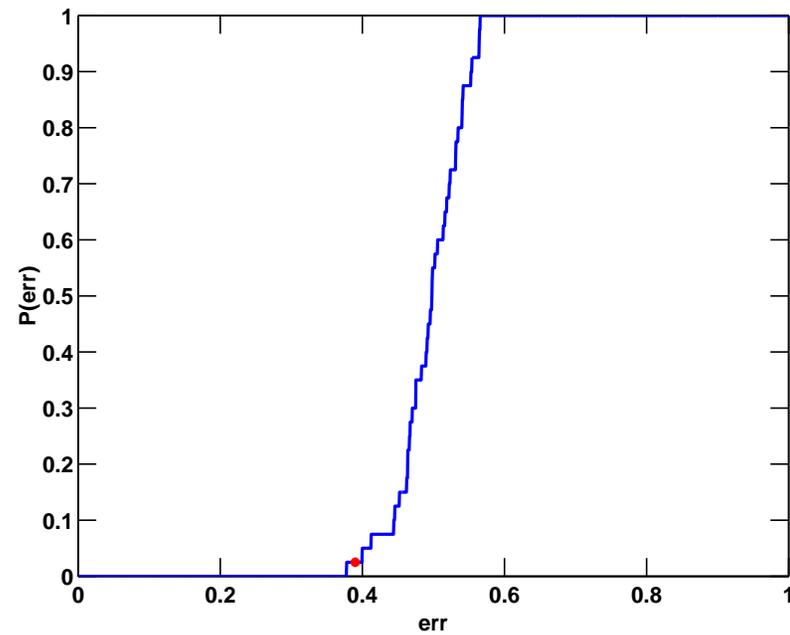
- the p-value of $\mathcal{T}[S]$ is $\hat{\mathbb{P}}(T \leq \mathcal{T}[S])$.

Toy example

$$\mathcal{T}[\mathcal{S}] = .39, .27, .25, .2 \text{ for } \mathcal{H}_4 \subset \mathcal{H}_3 \subset \mathcal{H}_2 \subset \mathcal{H}_1.$$

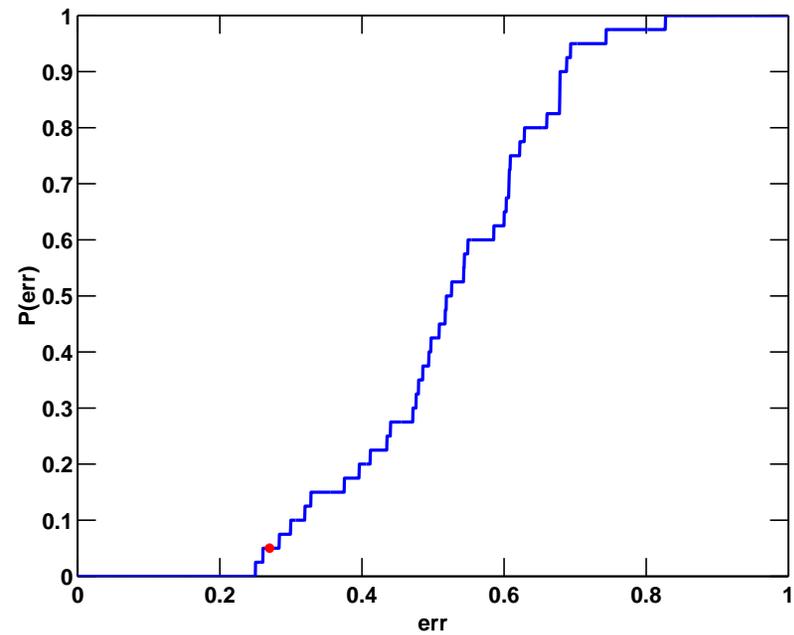
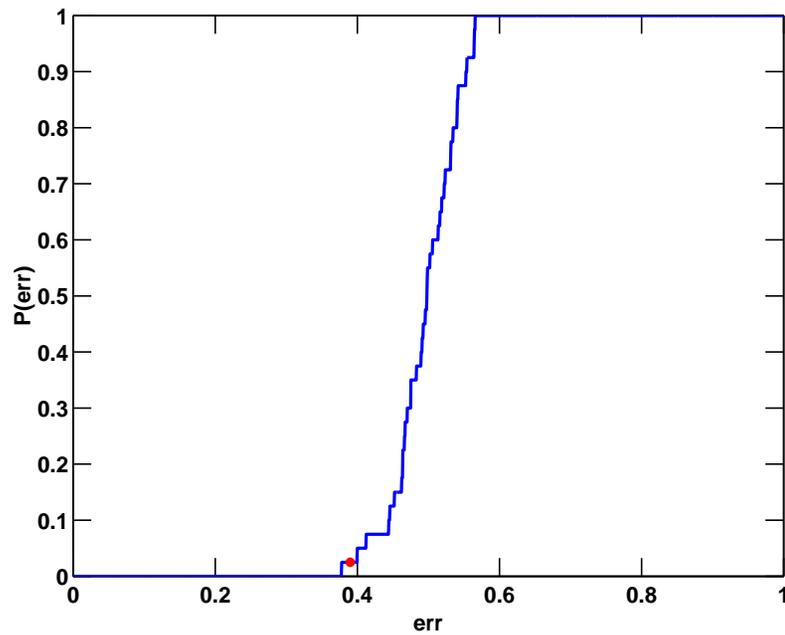
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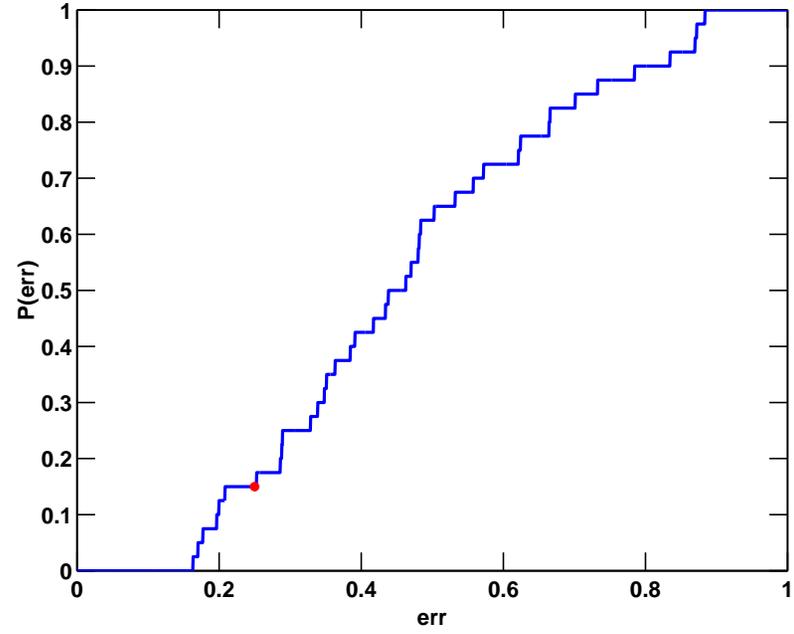
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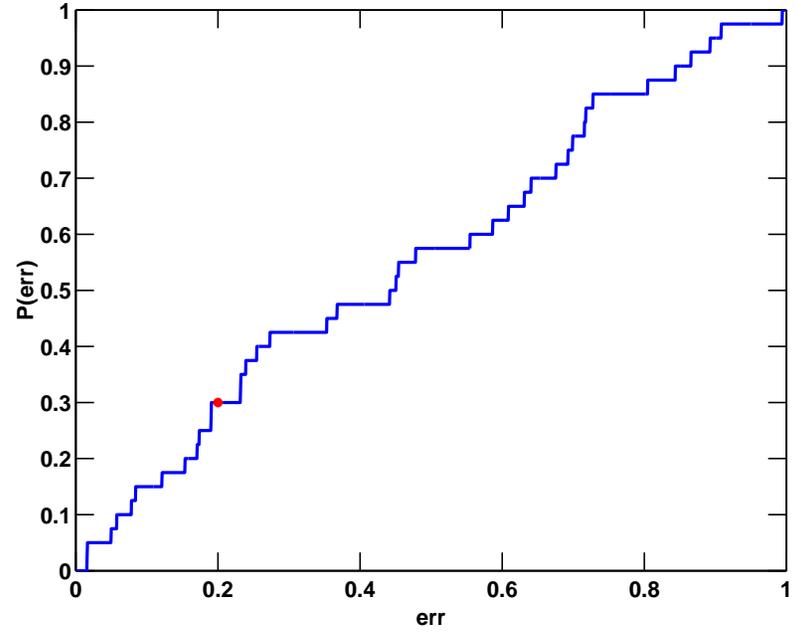
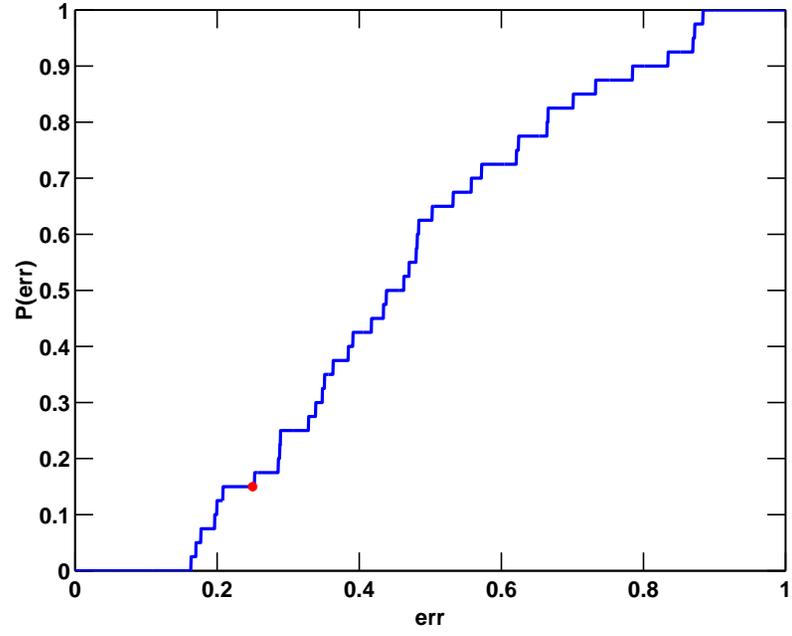


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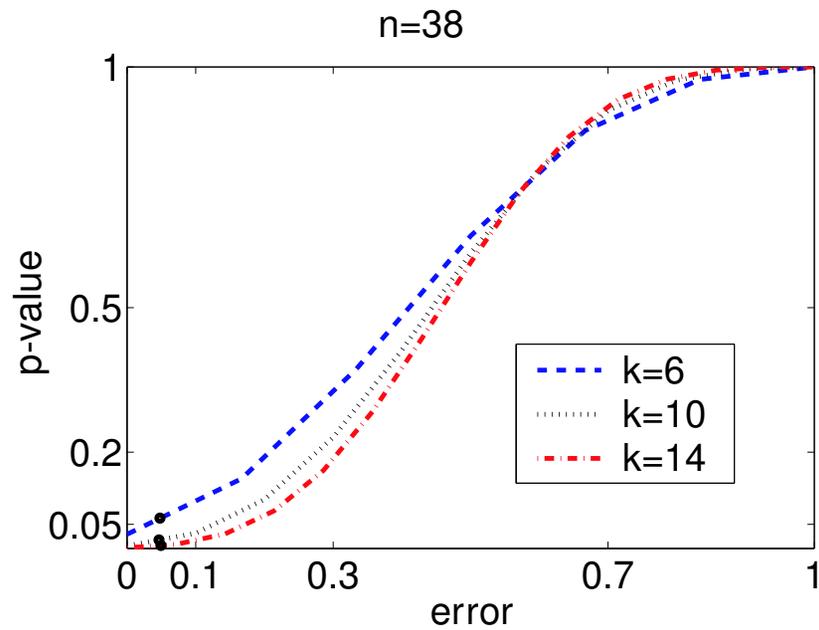


Leukemia

38 samples from 2 types of leukemia, picking k in leave- k -out.

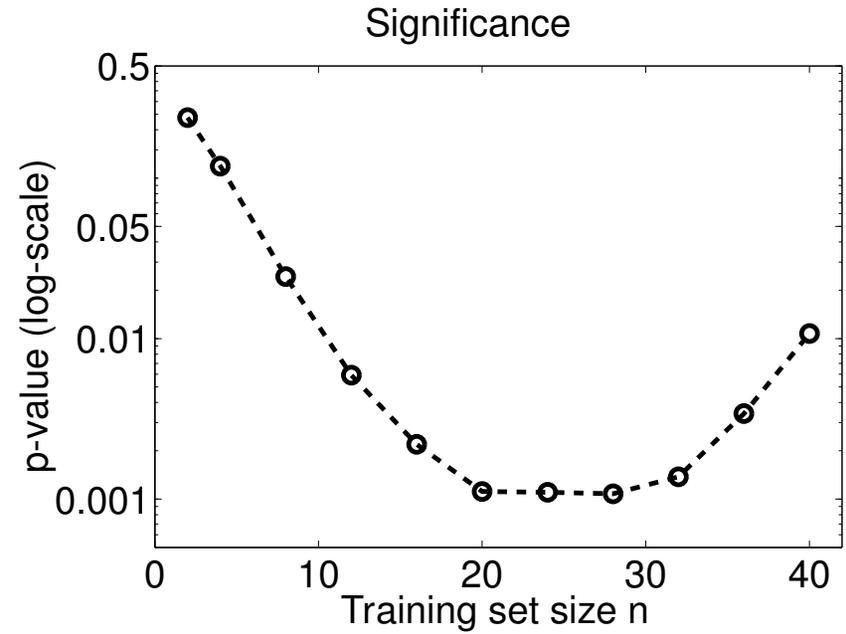
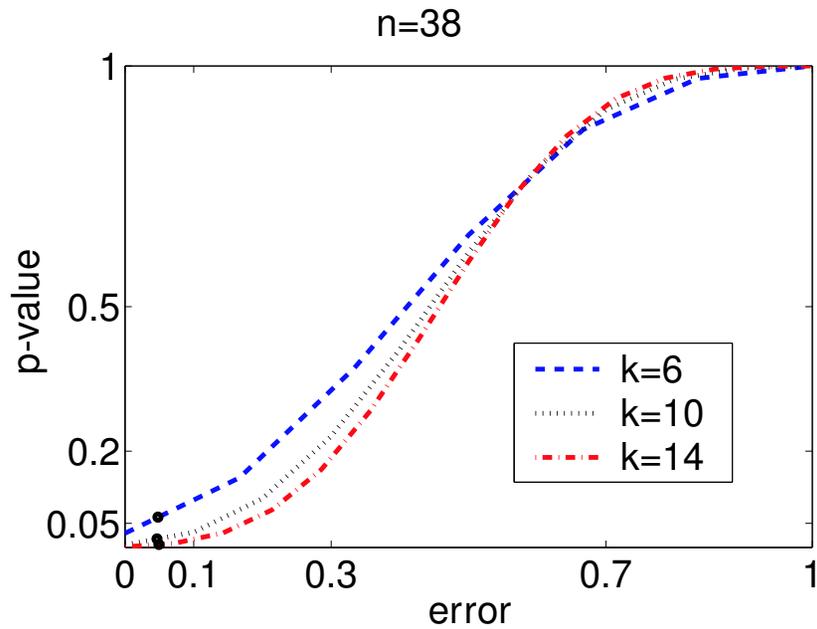
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Generalization of the permutation process

Given \mathcal{H} with target f_0 . For a permutation $\pi(S)$ the smallest training error is

$$\begin{aligned} e_n(\pi(S)) &= \min_{f \in \mathcal{H}} P_n(f \Delta f_0) \\ &= \min_{f \in \mathcal{H}} \left[\frac{1}{n} \sum_{i=1}^n I(z_i \in f, z_i^\pi \notin f_0) + I(z_i \notin f, z_i^\pi \in f_0) \right], \end{aligned}$$

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$$\mathbb{E}_\pi P_n(f \Delta f_0) = P(z \in f)(1 - P(z \in f_0)) + (1 - P(z \in f))P(z \in f_0) = P(z \in f_0) \equiv P(f_0).$$

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For appropriate complexity assumptions on \mathcal{H} prove that $e_n(\pi(S))$ is close to $P(f_0)$.

Concentration of the permutation process

The following maximization problem is equivalent to minimizing the empirical error on permuted data

$$e_n(\pi(S)) = P_n(z \in f_0) - \sup_{f \in \mathcal{H}} \left[\frac{1}{n} \sum_{i=1}^n I(z_i \in f) (2I(z_i^\pi \in f_0) - 1) \right].$$

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So we need only bound the following process

$$G_n(\pi(\mathcal{S})) = \sup_{f \in \mathcal{H}} \left[\frac{1}{n} \sum_{i=1}^n I(z_i \in f) (2I(z_i^\pi \in f_0) - 1) \right].$$

Bound on $G_n(\pi(S))$

Theorem 1. *If the \mathcal{H} has VC dimension V then with probability $1 - Ke^{-t/K}$*

$$G_n(\pi(S)) \leq K \min \left(\sqrt{\frac{V \log n}{n}}, \frac{V \log n}{n(1 - 2P(f_0))^2} \right) + \sqrt{\frac{Kt}{n}}.$$

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Therefore with probability $1 - Ke^{-t/K}$

$$P(z \in f_0) \leq P_n(z \in f_0) + K \min \left(\sqrt{\frac{V \log n}{n}}, \frac{V \log n}{n(1 - 2P(f_0))^2} \right) + \sqrt{\frac{Kt}{n}}.$$

Proof sketch (Part 1)

The process can be rewritten

$$G_n(\pi(S)) = \sup_{f \in \mathcal{H}} \left[\frac{1}{n} \sum_{i=1}^n I(z_i \in f) \varepsilon_i \right],$$

where $\varepsilon_i = 2I(z_i^\pi \in f_0) - 1 = \pm 1$ are Bernoulli random variables with $P(\varepsilon_i = 1) = P(f_0)$ and (ε_i) depend on (z_i) only through the cardinality of $\{z_i \in f_0\}$.

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The second term can be bounded by applying Chernoff's inequality twice.

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We need to bound the following process

$$\sup_{f \in \mathcal{H}} R[f, \varepsilon'] = \sup_{f \in \mathcal{H}} \left[\frac{1}{n} \sum_{i=1}^n I(z_i \in f) \varepsilon'_i \right],$$

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By Talagrand's inequality on the cube with probability $1 - e^{-Kt}$

$$\sup_{f \in \mathcal{H}} R[f, \varepsilon'] \leq \mathbb{E}_{\varepsilon'} \sup_{f \in \mathcal{H}} R[f, \varepsilon'] + \sqrt{\frac{t}{n}}.$$

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If $P(\varepsilon'_i = 1) = 1/2$ this is a Rademacher process and Dudley's entropy integral can be used to control $\mathbb{E}_{\varepsilon'} \sup_{f \in \mathcal{H}} R[f, \varepsilon']$.

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where $\eta_i = (\varepsilon_i' - \varepsilon_i'')/2$ takes values $\{-1, 0, 1\}$ and $P(\eta_i = 1) = P(\eta_i = -1)$.

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Since η_i satisfy

$$\mathbb{P} \left(\sum_{i=1}^n \eta_i a_i > t \right) \leq e^{-\frac{t^2}{2 \sum_{i=1}^n a_i^2}}$$

we can use the entropy integral.

Proof sketch (Part 4)

By the entropy integral bound

$$\mathbb{E}_{\eta_i} \sup_{f \in \mathcal{H}} \left[\frac{1}{n} \sum_{i=1}^n I[z_i \in f] \eta_i \right] \leq K \frac{1}{\sqrt{n}} \int_0^{\sqrt{\mu}} \sqrt{\log \mathcal{N}(u, \mathcal{H})} du$$

where $\mu = \frac{1}{n} \sum_{i=1}^n I[z_i \in f]$.

The result of the theorem is obtained by computing the entropy integral and optimizing. \square

Moral

- If $P(f_0) < 1/2$ then ignoring the “one dimensional terms” the rate of convergence is

$$O\left(\frac{V \log n}{n}\right).$$

- The weak dependency between (z_i) and a sequence (ε_i) can be broken with very little cost.

Risk bounds for mixture of densities

Given a dataset $S = \{x_1, \dots, x_n\}$ drawn i.i.d. from an unknown bounded (from above and below) density f_0 estimate this density using k -component mixtures f_k where

$$f_k \in \mathcal{C}_k = \text{conv}_k(\mathcal{H}) = \left\{ f : f(x) = \sum_{i=1}^k \lambda_i \phi_{\theta_i}(x), \sum_{i=1}^k \lambda_i = 1, \theta_i \in \Theta \right\},$$

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We are given an algorithm

$$\mathcal{A} : S \rightarrow \hat{f}_k$$

where $\hat{f}_k \in \mathcal{C}_k$.

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where $\mathcal{H} = \{\phi_{\theta}(x) : \theta \in \Theta \subset \mathbb{R}^d\}$.

We are given an algorithm

$$\mathcal{A} : S \rightarrow \hat{f}_k$$

where $\hat{f}_k \in \mathcal{C}_k$.

We want to bound

$$\mathbb{E}_S[D(f_0 || \hat{f}_k)] \leq \text{Approx}(\mathcal{C}_k) + \text{Est}(\mathcal{C}_k, n),$$

Risk bounds for mixture of densities

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The algorithms and some definitions

The following algorithms will be used

$$\mathcal{A}_{\text{MLE}} : \hat{f}_k = \arg \max_{\lambda, \theta} \sum_{i=1}^n \log \left[\sum_{j=1}^k \lambda_j \phi_{\theta_j}(z_i) \right]$$

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We define the class

$$\mathcal{C} = \text{conv}(\mathcal{H}) = \left\{ f : f(x) = \int_{\Theta} \phi_{\theta}(x) P(d\theta) \right\}$$

and

$$D(f_0 \| \mathcal{C}) = \inf_{g \in \mathcal{C}} D(f_0 \| g).$$

Approximation estimation tradeoff

Li and Barron proved the following:

Theorem 2. *Assume that Θ bounded and Lipschitz*

$$\sup_{x \in \mathcal{X}} |\log \phi_{\theta}(x) - \log \phi_{\theta'}(x)| \leq B \sum_{j=1}^d |\theta_j - \theta'_j|$$

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For either \mathcal{A}_{MLE} or $\mathcal{A}_{\text{Greedy}}$

$$\mathbb{E}_{\mathcal{S}} [D(f_0 \| \hat{f}_k)] - D(f_0 \| \mathcal{C}) \leq \frac{c_1}{k} + \frac{c_2 k}{n} \log(nc_3).$$

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The rate of convergence for optimal k is $O\left(\sqrt{\frac{\log n}{n}}\right)$.

There is no tradeoff in this problem

Alexander Rakhlin proved the following

Theorem 3. For any bounded f_0 ($a \leq f_0 \leq b$) then for either \mathcal{A}_{MLE} or $\mathcal{A}_{\text{Greedy}}$

$$\mathbb{E}_S [D(f_0 \|\hat{f}_k)] - D(f_0 \|\mathcal{C}) \leq \frac{c_1}{k} + \mathbb{E}_S \left[\frac{c_2}{\sqrt{n}} \int_0^b \sqrt{\log \mathcal{N}(\mathcal{H}, u, d_x)} du \right],$$

where $\mathcal{N}(\mathcal{H}, u, d_x)$ is the covering number of \mathcal{H} with respect to the empirical distance.

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There is no optimal k and only the complexity of \mathcal{H} is involved.

Why only \mathcal{H} (Part 1)

By McDiarmid's inequality with probability $1 - e^{-t}$

$$\sup_{h \in \mathcal{C}} \left| \frac{1}{n} \sum_{i=1}^n \log \frac{h(x_i)}{f_0(x_i)} - \mathbb{E} \log \frac{h}{f_0} \right| \leq \mathbb{E}_S \sup_{h \in \mathcal{C}} \left| \frac{1}{n} \sum_{i=1}^n \log \frac{h(x_i)}{f_0(x_i)} - \mathbb{E} \log \frac{h}{f_0} \right| + C\sqrt{t/n}.$$

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By symmetrization

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We will see that the Rademacher average above can be controlled only using \mathcal{H} .

Why only \mathcal{H} (Part 2)

Lemma 1. *Comparison inequality for Rademacher processes*

If $G : \mathbb{R} \rightarrow \mathbb{R}$ convex and non-decreasing and $\phi_i : \mathbb{R} \rightarrow \mathbb{R}$

($i = 1, \dots, n$) contractions ($\psi_i(0) = 0$ and $|\psi_i(s) - \psi_i(t)| \leq |s - t|$), then

$$\mathbb{E}_\varepsilon G \left[\sup_{f \in \mathcal{F}} \sum_{i=1}^n \varepsilon_i \psi_i(f(\mathbf{x}_i)) \right] \leq \mathbb{E}_\varepsilon G \left[\sup_{f \in \mathcal{F}} \sum_{i=1}^n \varepsilon_i f(\mathbf{x}_i) \right].$$

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Applying the above lemma multiple times gives us the following bound

$$\mathbb{E}_{S,\varepsilon} \sup_{h \in \mathcal{C}} \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i \log \frac{h(\mathbf{x}_i)}{f_0(\mathbf{x}_i)} \right| \leq K_1 \mathbb{E}_{S,\varepsilon} \sup_{h \in \mathcal{C}} \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i h(\mathbf{x}_i) \right|.$$

Why only \mathcal{H} (Part 3)

The Rademacher averages of a class are equal to those of the convex hull, since a linear functional of convex combinations achieves its maximum value at the vertices. Therefore,

$$\mathbb{E}_\varepsilon \sup_{h \in \mathcal{C}} \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i h(x_i) \right| = \mathbb{E}_\varepsilon \sup_{\theta \in \Theta} \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i \phi_\theta(x_i) \right|.$$

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Using the Dudley integral bound

$$\mathbb{E}_\varepsilon \sup_{\phi \in \mathcal{H}} \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i \phi(x_i) \right| \leq \frac{c_1}{\sqrt{n}} \int_0^b \log \sqrt{\mathcal{N}(\mathcal{H}, u, d_x)} du.$$

□

Moral

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- For estimates that are convex combinations of Lipschitz functionals the **estimation error** bound should be a function of the complexity of the base class \mathcal{H} and not the convex combination \mathcal{C} .
- When using mixture models control the complexity of the base class and use as many combinations as you want.