Graph inference via multiple testing Applications in neuroscience

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2 April 2019 1 / 40

The brain, an anatomical and functional network



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The brain, an anatomical and functional network



The brain, an anatomical and functional network



Interpretation of graph metrics



[Latora et al. 2001] [Bullmore et al. 2009] [Csárdi et al. 2006]

Interpretation of graph metrics



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Exploring the brain using networks analysis: summary



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2 April 2019 4 / 40

From dead salmon to dead rat

Sournal of Serendipitous and Unexpected Results

Neural Correlates of Interspecies Perspective Taking in the Post-Mortem Atlantic Salmon: An Argument For **Proper Multiple Comparisons Correction**

Craig M. Bennett 1*, Abigail A. Baird 2, Michael B. Miller 1 and George L. Wolford 3

(2011)





Dead rat results





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Reproducibility of global efficiency



Reproducibility of global efficiency



100 subjects rs-fMRI 2 sessions 14 min. TR=0.7s Intra class correlation $ICC = \frac{s_b - s_w}{s_b + s_w}$ s_b variance between subj.

 s_w variance within session



Data are available to download from my website!

[Termenon et al. 2016]

2 April 2019 7 / 40

Spatial pattern of reproducible local metrics



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Robustness to wrong identification of edges



Objectives



- Find the largest size of correct edges
- Control the number of non correct edges
- Optimize the procedures based on the size of data

Main questions

• What are the keys parameters to be robust?

• How to quantify the robustness of graph metrics?

• How to reveal alterations of networks?

Not large graphs, around hundreds of nodes. However complex multivariate time series.

Existing methods to estimate covariance matrices

- Hard thresholding [Bickel et al. 2008] Threshold of covariance matrices; cross validation.
- Shrinkage [Ledoit *et al.* 2004] Optimisation towards identity; forcing zeros in the covariance matrices.
- Regularisation [Bien et al. 2011] Optimisation using lasso penalisation; non convex problem.
- Graph Signal Processing [Pasdeloup et al. 2017] Optimisation through smoothness of signal on graphs; eigenvalues decomposition.

Graph inference with multiple testing

Objectives of our method:

- Simple application to any graph measures
- Control the false positive
- Control the false negative
- Unweighted graphs

Simple test

- $(X_i, Y_i)_{1 \le i \le n}$ i.i.d $\mathcal{N}(0, \Sigma)$.
- " $H_0: r = 0$ " against " $H_1: r \neq 0$ ".
- Reject of H_0 if $T(X) = |\hat{r}|/(\sqrt{(n-2)(1-\hat{r}^2)}) > t_{\alpha}$.
- Type I error: $\mathbb{P}_{H_0}(T(X) > t_{\alpha})$ (false positive).
- t_{α} type I error of level $\leq \alpha$. Here, t_{α} = quantile of order $(1 - \alpha/2)$ of St(n - 2) (law of T(X) under H_0).

p-value

A *p*-value, p(X), is a transformation of T(X) such that :

• Under
$$H_0$$
, $p(X) \sim \mathcal{U}(0,1)$,

• Under
$$H_1$$
, $p(X)$ "close to" 0.

Then,

reject
$$H_0$$
 if $p(X) \leq \alpha \Rightarrow$ type I error $\leq \alpha$.

Approach in order to reject part of null hypothesis

- Simultaneous tests of *m* mull hypotheses such that $m_0 = |\mathcal{H}_0|$ are true.
- Computation of the *m p*-values.
- Define the set $\mathcal{R} \subseteq \{1, \ldots, m\}$ of the rejected null hypotheses.

Definition: graph inference via multiple testing

•
$$\mathbb{X} = (X^{(1)}, \dots, X^{(n)})$$
 i.i.d sample such that for all $j = 1, \dots, n$,
 $X^{(j)} \sim \mathcal{N}_d(0, \Sigma)$ and $\Sigma = (\rho_{ii'})_{1 \leq i, i' \leq d}$, $\rho_{ii} = 1$ for all
 $i = 1, \dots, d$.

•
$$\mathcal{H} = \{(i, i'), i < i'\}$$
. For all $h \in \mathcal{H}$, do $m = d(d - 1)/2$ tests
 $H_{0,h}$: " $\rho_h = 0$ " against $H_{1,h}$: " $\rho_h \neq 0$ ",

via the test statistic $T_{n,h}$ and p values $p_{n,h}$. For instance, with empirical correlation,

$$T_{n,h}\left(\mathbb{X}\right) = \sqrt{n} \frac{\frac{1}{n} \sum_{j=1}^{n} X_i^{(j)} X_{i'}^{(j)} - \bar{X}_i \bar{X}_{i'}}{\hat{\sigma}(X_i) \hat{\sigma}(X_{i'})} = \sqrt{n} \hat{\rho}_{n,h}.$$

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- *d* = 5.
- $\mathcal{H} = \{(1,2); (1,3); (1,4); (1,5); (2,3); (2,4); (2,5); (3,4); (3,5); (4,5)\}.$
 - $H_{0,h}$: " $\rho_h = 0$ " contre $H_{1,h}$: " $\rho_h \neq 0$ ".
- Construction of the graph:



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• Construction of the graph:

 $\mathcal{R} = \{(1,2)$

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Graph inference via multiple testing

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- *d* = 5.
- $\mathcal{H} = \{(1,2); (1,3); (1,4); (1,5); (2,3); (2,4); (2,5); (3,4); (3,5); (4,5)\}.$

 $H_{0,h}$: " $\rho_h = 0$ " contre $H_{1,h}$: " $\rho_h \neq 0$ ".

• Construction of the graph:

 $\mathcal{R} = \{(1,2); (1,3)\}$

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- *d* = 5.
- $\mathcal{H} = \{(1,2); (1,3); (1,4); (1,5); (2,3); (2,4); (2,5); (3,4); (3,5); (4,5)\}.$

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• Construction of the graph:



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Why it is important to take into account the multiplicity in the context of graphs?

- 51 brain regions: 1275 tests.
- If we test **true** null hypotheses **indepently** simultaneously at level of significance $\alpha = 5\%$.
- What is the probability to get at least one significant results by chance (*i.e.* reject a null hypothesis)?

 $\mathbb{P}(\text{``at least one significant result''}) = 1 - \mathbb{P}(\text{``no significant results''})$ $= 1 - (1 - 0.05)^{1275}$ $\simeq 1$

Control of the error due to multiplicity

Two main criteria are usually used based on \mathcal{Q} , the proportion of false discovery

$$\mathcal{Q} = rac{|\mathcal{R} \cap \mathcal{H}_0|}{|\mathcal{R}| ee 1}.$$

•
$$\mathsf{FWER}(\mathcal{R}) = \mathbb{P}(\mathcal{Q} > 0)$$

•
$$\mathsf{FDR}(\mathcal{R}) = \mathbb{E}[\mathcal{Q}]$$

Illustration FWER vs FDR

Low signal strength

Strong signal strength



[Roquain 2007]

asymptotic control of FWER

Our objective:

lim sup FWER(\mathcal{R}) $\leq \alpha$. $n \rightarrow +\infty$

- Asymptotically in *n*
- Valid as soon as the statistics are converging
- Valid given any graph structures
- Can we evaluate the power?
- Can we extend to FDR?

asymptotic control of FWER for correlations

For all $h \in \{1, ..., m\}$, we consider the test statistic $T_{n,h}(\mathbb{X}^{(n)})$. In the case of correlations,

$$T_{n,h}\left(\mathbb{X}^{(n)}\right) = \sqrt{n}\,\widehat{\rho}_{n,h}\left(\mathbb{X}^{(n)}\right).$$

Assume that,

$$\sqrt{n}\left(\, \hat{
ho}_{n,\cdot}(\mathbb{X}^{(n)}) -
ho(\mathcal{P})
ight) \, \stackrel{d}{\longrightarrow} \, \mathcal{N}_m(0,\Sigma), \quad ext{when} \, \, n o +\infty,$$

where \xrightarrow{d} denotes the convergence in distribution. We assume that Σ is invertible.

Let $(p_{n,h}(\mathbb{X}^{(n)}))_{1 \le h \le m}$ be a family of *p*-values resulting from each *m* individual test:

$$\forall h \in \{1, \ldots, m\}, \ p_{n,h}\left(\mathbb{X}^{(n)}\right) = 2\left[1 - \Phi\left(\frac{\left|\sqrt{n} \,\widehat{\rho}_{n,h}\left(\mathbb{X}^{(n)}\right)\right|}{\sqrt{1 - \widehat{\rho}_{n,h}^2\left(\mathbb{X}^{(n)}\right)}}\right)\right],$$

where $\boldsymbol{\Phi}$ is the standard Gaussian cumulative distributive function.

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2 April 2019 23 / 40

Bonferroni

•
$$\mathcal{R}^{bonf} = \Big\{ h \in \mathcal{H} : p_{n,h} \leqslant \alpha/2m \Big\}.$$

Theorem: [Dunn, 1961]

The procedure \mathcal{R}^{bonf} controls asymtotically the FWER at significance level α ,

 $\mathsf{FWER}(\mathcal{R}^{\mathsf{bonf}}) \leq \alpha.$

Proposition : Let $X \sim \mathcal{N}_m(0, \Sigma)$ with Σ invertible, then for all $b \in \mathbb{R}^m_+$, $\mathbb{P}(|X| \le b) \ge \prod_{i=1}^m \mathbb{P}(|X_i| \le b_i).$

Šidák

•
$$\mathcal{R}^s = \left\{ h \in \mathcal{H} : |T_{n,h}| > \Phi^{-1} \left(\frac{1}{2} (1-\alpha)^{(1/m)} + \frac{1}{2} \right) \right\}.$$

Theorem: [Westfall, Young 1993 and Drton, Perlmann 2004] The procedure \mathcal{R}^s controls asymtotically the FWER at significance level α ,

 $\limsup_{n \to +\infty} \mathsf{FWER}(\mathcal{R}^s) \le \alpha.$

• Procedure valid for all types of dependence (Gaussian case).

• FWER controlled such as in the simple case!

27 / 40

Summary of statistical framework

 $\mathbf{X} = {\mathbf{X}(k), k \in \mathbb{Z}}$ long memory process, $1 \leq \ell, m \leq d$, Statistical Hypotheses Choice of Statistics $\mathcal{H}_0^{(h)}: \rho_{\ell,m} = 0 \qquad \mathcal{H}_1^{(h)}: \rho_{\ell,m} \neq 0$ $M = \frac{d(d-1)}{2}$ hypotheses number $\mathcal{R} =$ set of rejected null hypotheses

Tests are dependent!

Controlling procedure

$$\widehat{\rho}_{n,h} \text{ s.t. when } n \to +\infty,$$

 $\sqrt{n} \left(\widehat{\rho}_{n,\cdot}(\mathbb{X}^{(n)}) - \rho(P) \right) \longrightarrow \mathcal{N}_m(0, \Sigma),$
evaluation of *p*-values = $p_{n,h}$

Choice of criterion

 $p_{n,h} < \alpha/M$

$$\hat{h} = \max\{h \in \{0, \dots, m\} : p_{(n,h)} \le \alpha h/m\}$$



 $FWER(\mathcal{R}, P) = \mathbb{P}(|\mathcal{R} \cap \mathcal{H}_0(P)| > 1)$

$$\mathsf{FDR}(\mathcal{R}, P) = \mathbb{E}\left[rac{|\mathcal{R} \cap \mathcal{H}_0(P)|}{|\mathcal{R}| \lor 1}
ight]$$

[Roux et al. 2018] [Hero et al. 2013] [Drton et al. 2004] 2 April 2019 28/40

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Simulations

Simulate **M** such that $\mathbf{M} = I + \rho A$, where **I** is the identity matrix, and ρ is a real.

M is positive definite if and only if $|\rho| < \frac{1}{|\lambda_{min}|}$, where λ_{min} is the smallest eigenvalue.

Let us take d = 51, m = 150. This is true for $\rho = 0.2$ for 4 different structures: Erdos-Renyi, Stochastic Block Model, Small-World and Scale-free.

Network structures used in simulations



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using Bonferroni: FWER control is preserved as expected



using Bonferroni: power



Comparisons with shrinkage (Ledoit-Wolf)



Comparisons with FDR



2 April 2019 34 / 40

Real data



- 4 anesthetics: Isoflurane, Medetomidine, Edetomidate, Urethane
- 5 groups of rats
- 30-minutes resting-state fMRI, TR=0.5s

[Becq et al. 2018, in preparation]

Significant edges detection

Opitimisation of signal to noise ratio on edge detection



Graphs for anesthetics

Spatial reorganization of 5% highest correlation pairs



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Anesthetics graphs profiles

Spatial reorganisation of 5% highest correlation pairs



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Conclusion and perspectives

Conclusion

- Control of false positive is not easy
- Graph networks properties depend on false positive

Perspectives

- Control of false discovery rate may be more appropriate
- Extension to any other mesures of correlation or partial correlation

Thanks to my collaborators



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