Testing Dependency of Databases

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Correlated data structures

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- Modern challenges: data structures are high-*d*, noisy, **unlabeled/scrambled**.
- This precludes "direct" inference/data junction.
- General goal: determine if ∃ a correspondence under which the sources are "correlated".

Pictorially...

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- The correspondence between different sources is unknown/obfuscated.





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- If "correlation" is sufficiently large maybe it is possible to glean something about the correspondence.



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- Valuable tool to recover missing information by labeling unlabeled features and allowing the junction of data coming from different sources.



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- Each source provides information for entities (e.g., users).
- Crucial to understand limitations of data alignment so as to assess the feasibility and reliability of alignment procedures.



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- Dataset: lists of features for a set of entities, say, users.

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- No side information: could be effective for protecting user privacy (while providing access to data).

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- [Narayanan&Shmatikov'08,09]: many Netflix user IDs can be matched with IMDb profiles.
- Netflix prize dataset (anonymized): User IDs, movie IDs, movie ratings.
- IMDb dataset (public): Usernames, movie names, movie ratings.

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- Side information is abundant in the public domain!
- Crucial to understand the conditions that allow/prevent privacy breaches, and vulnerability of de-anony. schemes.

"Interactions among users"

• In many modern applications, observations appear as graphs.



[Wu&Xu&Yu'21]

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- Node labels may be absent or scrambled.



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- Social network analysis: two friendship networks on different social platforms share structural similarities?
- Computational biology: assess the correlation of two biological networks in two different species.
- Natural language processing: uncovering the correlation between two knowledge graphs that are in either different languages.

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- Natural language processing: uncovering the correlation between two knowledge graphs that are in either different languages.
- Significant attention and beautiful strong results, e.g., [Barak et. al.'19], [Cullina,Kiyavash'16,20], [Wu,Xu,Yu'21], [Ding, Ma, Wu, Xu'21], [Hall,Massoulié'21], [Ding,Li'22], [Ding,Du'23], and many references therein.

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- For now, databases include the same set of users.



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- We will assume features are i.i.d.



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- Features (X_i, Y_{πi}) associated with user i are dependent, while different pairs are independent.



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- Received significant attention, e.g., [Cullina,Mittal,Kiyavash'18],[Dai,Mittal,Kiyavash'19], [Wang,Wu,Xu,Yolou'22].



Generative Correlation Model

• In this talk, we focus on the detection variant of this problem.



Detection/Hypothesis Testing

• Null: X and Y are Gaussian and independent, i.e.,

$$(\mathsf{X}_1,\mathsf{Y}_1),\ldots,(\mathsf{X}_n,\mathsf{Y}_n) \overset{\text{i.i.d.}}{\sim} \mathcal{N}^{\otimes d}(0_{2\times 1},\mathbf{I}_{2\times 2})$$



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- Alternative: cond. on $\pi \sim \mathsf{Unif}(\mathbb{S}_n)$ (or $\exists \pi \in \mathbb{S}_n$)

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• For a test $\phi:\mathbb{R}^{n\times d}\times\mathbb{R}^{n\times d}\to\{0,1\},$ the "risk" is:

$$\mathsf{R}(\phi) \triangleq \mathbb{P}_{\mathcal{H}_0}[\phi(\mathsf{X},\mathsf{Y}) = 1] + \mathbb{E}_{\pi \sim \mathsf{Unif}(\mathbb{S}_n)} \mathbb{P}_{\mathcal{H}_1|\pi}[\phi(\mathsf{X},\mathsf{Y}) = 0].$$

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• Minimal (optimal) risk $R^* = \inf_{\phi} R(\phi)$.
Detecting Correlated Databases

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• Possibility: strong detection if $R(\phi) = o(1)$, and weak detection if $\lim R(\phi) < 1$.

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- Impossibility: strong detection if $R^* = \Omega(1)$, and weak detection if $R^* = 1 o(1)$.

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- Possibility: strong detection if $\lim d_{\mathsf{TV}}(\mathbb{P}_{\mathcal{H}_0}, \mathbb{P}_{\mathcal{H}_1}) = 1$, and weak detection if $\liminf d_{\mathsf{TV}}(\mathbb{P}_{\mathcal{H}_0}, \mathbb{P}_{\mathcal{H}_1}) > 0$.
- Impossibility: strong detection if d_{TV}(ℙ_{H0}, ℙ_{H1}) ≤ 1 − Ω(1), and weak detection if d_{TV}(ℙ_{H0}, ℙ_{H1}) = o(1).

Known Results and Gaps

• [Dai,Cullina,Kiyavash'19]: Perfect recovery is possible if $\rho^2 = 1 - o(n^{-4/d})$ and impossible if $\rho^2 = 1 - \omega(n^{-4/d})$, assuming $1 \ll d = O(\log n)$.

E.g., if $d = \omega(\log n)$ then rec. is possible if $\rho^2 = \omega\left(\frac{\log n}{d}\right)$.

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- [Zeynep,Nazer'21,22]: (Efficient) strong detection *possible* if $\rho^2 d \to \infty$, and *impossible* if $\rho^2 d \sqrt{n} \to 0$ and $d = \Omega(\log n)$

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- [Tamir'22,23]: Joint correlation detection and recovery.

Main Results (Correlated Databases)

We show in [Elimelech, Huleihel'23, 24]

	Weak Detection		Strong Detection	
Asymptotics	Possible	Impossible	Possible	Impossible
$n, d \to \infty$	$\Omega(d^{-1})$	$o(d^{-1})$	$\omega(d^{-1})$	$(1-\varepsilon)d^{-1}$
$d ightarrow \infty$, n constant	$\Omega(d^{-1})$	$o(d^{-1})$	$\omega(d^{-1})$	$O(d^{-1})$
$n ightarrow \infty$, d constant	$\rho^2 = \Omega(1)$	o(1)	$1 - o(n^{-\frac{4}{d}})$	$ ho^{\star}(d)$

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- Fixed d is the interesting and more challenging regime.
- $\bullet \ \ {\rm We \ use:} \ \ d\rho^2 \to 0 \Leftrightarrow \rho^2 = o(d^{-1}) \Leftrightarrow d\rho^2 = o(1).$

• **Total sum** [Zeynep,Nazer'21,22]: Threshold the sum of inner-products

$$\phi_{\mathsf{sum}}(\mathsf{X},\mathsf{Y}) \triangleq \mathbbm{1} \left\{ \sum_{i,j=1}^n \mathsf{X}_i^T \mathsf{Y}_j > \frac{dn\rho}{2} \right\}.$$

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Chernoff's bound gives:

$$\mathsf{R}(\phi_{\mathsf{sum}}) \le 2 \cdot \exp\left(-\frac{d\rho^2}{60}\right).$$

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 Completely independent of n.
 If d is fixed, then strong detection using φ_{sum} is not possible.

• Counting products [Elimelech, Huleihel'24]: Consider

$$\phi_{\mathsf{count}}(\mathsf{X},\mathsf{Y}) \triangleq \mathbb{1}\left\{\sum_{i,j=1}^{n} \mathbb{1}\left\{\mathsf{L}(\mathsf{X}_{i},\mathsf{Y}_{j}) \geq d \cdot \tau_{\mathsf{count}}\right\} \geq \frac{n\mathcal{P}_{d}}{2}\right\}$$

where

$$\mathsf{L}(\mathsf{X}_i,\mathsf{Y}_j) \triangleq -\frac{d}{2}\log(1-\rho^2) - \frac{d\rho^2}{2(1-\rho^2)} + \frac{\rho}{1-\rho^2}\mathsf{X}_i^T\mathsf{Y}_j$$

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$$\begin{split} \mathsf{L}(\mathsf{X}_{i},\mathsf{Y}_{j}) &\triangleq -\frac{d}{2}\log(1-\rho^{2}) - \frac{d\rho^{2}}{2(1-\rho^{2})} + \frac{\rho}{1-\rho^{2}}\mathsf{X}_{i}^{T}\mathsf{Y}_{j} \\ &= \log\frac{P_{XY}^{\otimes d}(\mathsf{X}_{i},\mathsf{Y}_{j})}{Q_{XY}^{\otimes d}(\mathsf{X}_{i},\mathsf{Y}_{j})}. \end{split}$$

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Theorem (Count test strong detection)

Fix $d \in \mathbb{N}$. Then, $\mathsf{R}(\phi_{\mathsf{count}}) \to 0$, as $n \to \infty$, if $\rho^2 = 1 - o(n^{-4/d})$.

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Coincides with the recovery threshold (via ML).

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$$\phi_{\mathsf{count}}(\mathsf{X},\mathsf{Y}) \triangleq \mathbbm{1}\left\{\sum_{i,j=1}^n \mathbbm{1}\left\{\mathsf{L}(\mathsf{X}_i,\mathsf{Y}_j) \geq d \cdot \tau_{\mathsf{count}}\right\} \geq \frac{n\mathcal{P}_d}{2}\right\}$$

where

$$\mathsf{L}(\mathsf{X}_i,\mathsf{Y}_j) \triangleq -\frac{d}{2}\log(1-\rho^2) - \frac{d\rho^2}{2(1-\rho^2)} + \frac{\rho}{1-\rho^2}\mathsf{X}_i^T\mathsf{Y}_j$$

Theorem (Count test strong detection)

Fix $d \in \mathbb{N}$. Then, $\mathsf{R}(\phi_{\mathsf{count}}) \to 0$, as $n \to \infty$, if $\rho^2 = 1 - o(n^{-4/d})$.



• Counting products [Elimelech, Huleihel'24]: Consider

$$\phi_{\mathsf{count}}(\mathsf{X},\mathsf{Y}) \triangleq \mathbb{1}\left\{\sum_{i,j=1}^{n} \mathbb{1}\left\{\mathsf{L}(\mathsf{X}_{i},\mathsf{Y}_{j}) \geq d \cdot \tau_{\mathsf{count}}\right\} \geq \frac{n\mathcal{P}_{d}}{2}\right\}$$

Proof sketch: first moment

$$\mathbb{P}_{\mathcal{H}_0}\left(\phi_{\mathsf{count}}=1\right) = \mathbb{P}_{\mathcal{H}_0}\left(\sum_{i,j=1}^n \mathsf{G}_{ij} \geq \frac{n\mathcal{P}_d}{2}\right) \leq \frac{2n\mathcal{Q}_d}{\mathcal{P}_d},$$

where

$$\begin{split} \mathcal{Q}_{d} &\triangleq \mathbb{P}_{\mathcal{N}^{\otimes d}(0,\mathbf{I})}[\mathsf{L}(\mathsf{A},\mathsf{B}) \geq d \cdot \tau_{\mathsf{count}}] \leq e^{-d \cdot E_{Q}(\tau_{\mathsf{count}})} \\ \mathcal{P}_{d} &\triangleq \mathbb{P}_{\mathcal{N}^{\otimes d}(0,\Sigma_{\rho})}[\mathsf{L}(\mathsf{A},\mathsf{B}) \geq d \cdot \tau_{\mathsf{count}}] \geq 1 - e^{-d \cdot E_{P}(\tau_{\mathsf{count}})} \end{split}$$

• Counting products [Elimelech, Huleihel'24]: Consider

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<u>Proof sketch</u>: second moment (w.l.o.g. $\pi = Id$),

$$\begin{split} \mathbb{P}_{\mathcal{H}_1} \left(\phi_{\mathsf{count}} = 0 \right) &= \mathbb{P}_{\mathcal{H}_1} \left(\sum_{i,j=1}^n \mathsf{G}_{ij} < \frac{n\mathcal{P}_d}{2} \right) \\ &\leq \mathbb{P}_{\mathcal{H}_1} \left(\sum_{i=1}^n \mathsf{G}_{ii} < \frac{n\mathcal{P}_d}{2} \right) \\ &\leq \frac{4 \cdot \mathsf{Var}_\rho \left(\sum_{i=1}^n \mathsf{G}_{ii} \right)}{n^2 \mathcal{P}_\rho^2} = \frac{4(1-\mathcal{P}_d)}{n\mathcal{P}_d} \leq \frac{4}{n\mathcal{P}_d}. \end{split}$$

• Comparison test [Elimelech,Huleihel'24]: Define,

$$\phi_{\mathsf{comp}}(\mathsf{X},\mathsf{Y}) \triangleq \mathbb{1}\left\{ \left| \sum_{i,j} (\mathsf{X}_{ij} - \mathsf{Y}_{ij}) \right| \leq \theta \right\}$$

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Take θ as the value for which

$$d_{\mathsf{TV}}\left(\mathcal{N}(0,1), \mathcal{N}\left(0,1-|\rho|\right)\right) \\ = \mathbb{P}\left(|\mathsf{G}| \ge \frac{\theta}{\sqrt{2nd}}\right) - \mathbb{P}\left(|\mathsf{G}'| \ge \frac{\theta}{\sqrt{2nd}}\right),$$

where $\mathsf{G} \sim \mathcal{N}(0,1)$ and $\mathsf{G}' \sim \mathcal{N}\left(0,1-|\rho|\right).$

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Theorem

Fix $d \in \mathbb{N}$. If $\rho^2 = \Omega(1)$ then $\lim_{n \to \infty} \mathsf{R}(\phi_{\mathsf{comp}}) < 1$.

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$$\phi_{\mathsf{comp}}(\mathsf{X},\mathsf{Y}) \triangleq \mathbb{1}\left\{ \left| \sum_{i,j} (\mathsf{X}_{ij} - \mathsf{Y}_{ij}) \right| \leq \theta \right\}$$

<u>Proof sketch</u>: Let $G_1 \triangleq \sum_{ij} X_{ij}$ and $G_2 \triangleq \sum_{ij} Y_{ij}$. Then, $G_1 - G_2 \stackrel{\mathcal{H}_0}{\sim} \mathcal{N}(0, 2nd)$ and $G_1 - G_2 \stackrel{\mathcal{H}_1}{\sim} \mathcal{N}(0, 2nd(1 - \rho))$. Therefore,

$$\begin{split} 1 - \mathsf{R}(\phi_{\mathsf{comp}}) &= \mathbb{P}_{\mathcal{H}_0}(|\mathsf{G}_1 - \mathsf{G}_2| \ge \theta) - \mathbb{P}_{\mathcal{H}_1}(|\mathsf{G}_1 - \mathsf{G}_2| \ge \theta) \\ &= \mathbb{P}(|\mathcal{N}(0, 2nd)| \ge \theta) \\ &- \mathbb{P}(|\mathcal{N}(0, 2n(1-\rho))| \ge \theta) \\ &= d_{\mathsf{TV}}\left(\mathcal{N}(0, 1), \mathcal{N}\left(0, 1-\rho\right)\right) = \Omega(1). \end{split}$$

We start with the regime where at least $d \to \infty$.

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Second moment calculation: let $L_n(X, Y) \triangleq \frac{\mathbb{P}_{\mathcal{H}_1}(X, Y)}{\mathbb{P}_{\mathcal{H}_0}(X, Y)}$, then

$$\mathsf{R}^{\star} = 1 - d_{\mathsf{TV}}(\mathbb{P}_{\mathcal{H}_1}, \mathbb{P}_{\mathcal{H}_0})$$

$$\mathbb{E}_{\mathcal{H}_0} \left[\mathsf{L}_n^2 \right] = O(1)$$

$$\implies d_{\mathsf{TV}}(\mathbb{P}_{\mathcal{H}_1}, \mathbb{P}_{\mathcal{H}_0}) \le 1 - \Omega(1)$$

$$\mathbb{E}_{\mathcal{H}_0} \left[\mathsf{L}_n^2 \right] = 1 + o(1)$$

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$$\implies d_{\mathsf{TV}}(\mathbb{P}_{\mathcal{H}_1}, \mathbb{P}_{\mathcal{H}_0}) \le o(1)$$

Thus, it is suffice to analyze the second moment of the likelihood.

Recall that

$$\begin{split} \mathsf{L}_{n}(\mathsf{X},\mathsf{Y}) &= \frac{\mathbb{P}_{\mathcal{H}_{1}}(\mathsf{X},\mathsf{Y})}{\mathbb{P}_{\mathcal{H}_{0}}(\mathsf{X},\mathsf{Y})} \\ &= \frac{\mathbb{E}_{\pi}[\mathbb{P}_{\mathcal{H}_{1}\mid\pi}(\mathsf{X},\mathsf{Y})]}{\mathbb{P}_{\mathcal{H}_{0}}(\mathsf{X},\mathsf{Y})} = \mathbb{E}_{\pi}\left[\frac{\mathbb{P}_{\mathcal{H}_{1}\mid\pi}(\mathsf{X},\mathsf{Y})}{\mathbb{P}_{\mathcal{H}_{0}}(\mathsf{X},\mathsf{Y})}\right] \end{split}$$

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Then,

Thus, Ingster-Suslina method (Fubini's theorem)

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi \perp \perp \pi'} \left[\mathbb{E}_{\mathcal{H}_0} \left[\frac{\mathbb{P}_{\mathcal{H}_1 \mid \pi}}{\mathbb{P}_{\mathcal{H}_0}} \cdot \frac{\mathbb{P}_{\mathcal{H}_1 \mid \pi'}}{\mathbb{P}_{\mathcal{H}_0}} \right] \right]$$

.

.
Invariance: fix $\pi' = Id$,

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi}\left[\mathbb{E}_{\mathcal{H}_0}\left[\frac{\mathbb{P}_{\mathcal{H}_1|\pi}}{\mathbb{P}_{\mathcal{H}_0}}\cdot\frac{\mathbb{P}_{\mathcal{H}_1}|\mathsf{Id}}{\mathbb{P}_{\mathcal{H}_0}}\right]\right].$$

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Recall that pairs $\{(\mathsf{X}_i,\mathsf{Y}_{\pi_i})\}_{i\in[n]}$ are i.i.d.,

$$\frac{\mathbb{P}_{\mathcal{H}_1|\pi}(\mathsf{X},\mathsf{Y})}{\mathbb{P}_{\mathcal{H}_0}(\mathsf{X},\mathsf{Y})} = \prod_{i=1}^n \mathsf{L}(\mathsf{X}_i,\mathsf{Y}_{\pi_i})$$
$$\frac{\mathbb{P}_{\mathcal{H}_1|\mathsf{Id}}(\mathsf{X},\mathsf{Y})}{\mathbb{P}_{\mathcal{H}_0}(\mathsf{X},\mathsf{Y})} = \prod_{i=1}^n \mathsf{L}(\mathsf{X}_i,\mathsf{Y}_i),$$

where
$$\mathsf{L}(\mathsf{X}_{i},\mathsf{Y}_{i}) \triangleq rac{P_{XY}^{\otimes d}(\mathsf{X}_{i},\mathsf{Y}_{i})}{Q_{XY}^{\otimes d}(\mathsf{X}_{i},\mathsf{Y}_{i})}.$$

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$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi} \left[\mathbb{E}_{\mathcal{H}_0} \left[\frac{\mathbb{P}_{\mathcal{H}_1 \mid \pi}}{\mathbb{P}_{\mathcal{H}_0}} \cdot \frac{\mathbb{P}_{\mathcal{H}_1 \mid \mathsf{ld}}}{\mathbb{P}_{\mathcal{H}_0}} \right] \right].$$

Thus,

$$\frac{\mathbb{P}_{\mathcal{H}_1|\pi}}{\mathbb{P}_{\mathcal{H}_0}} \cdot \frac{\mathbb{P}_{\mathcal{H}_1|\mathsf{Id}}}{\mathbb{P}_{\mathcal{H}_0}} = \prod_{i=1}^n \mathsf{L}(\mathsf{X}_i,\mathsf{Y}_{\pi_i})\mathsf{L}(\mathsf{X}_i,\mathsf{Y}_i) \triangleq \prod_{i=1}^n \mathsf{Z}_i$$

Invariance: fix $\pi' = Id$,

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi}\left[\mathbb{E}_{\mathcal{H}_0}\left[\frac{\mathbb{P}_{\mathcal{H}_1|\pi}}{\mathbb{P}_{\mathcal{H}_0}}\cdot\frac{\mathbb{P}_{\mathcal{H}_1}|\mathsf{Id}}{\mathbb{P}_{\mathcal{H}_0}}\right]\right].$$

Accordingly,

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi}\left[\mathbb{E}_{\mathcal{H}_0}\left(\prod_{i=1}^n \mathsf{Z}_i\right)\right].$$

Problem: $\{Z_i\}_{i=1}^n$ are dependent random variables Solution: cycle decomposition!

Facts on cycles (orbits)

• For each element $a \in [n]$, its orbit is a cycle (a_0, \ldots, a_{k-1}) , where $a_i = \pi^i(a)$, for $i = 0, \ldots, k-1$ and $\pi(a_{k-1}) = a$.

For example: Consider $\pi \in \mathbb{S}_7$ that

- Keeps 1 in the same place
- Swaps 2 with 3
- Occilically shifts 4567

Then, π consists of <u>three</u> orbits in canonical notation

 $\pi = (1)(23)(4567)$

Facts on cycles (orbits)

- For each element $a \in [n]$, its orbit is a cycle (a_0, \ldots, a_{k-1}) , where $a_i = \pi^i(a)$, for $i = 0, \ldots, k-1$ and $\pi(a_{k-1}) = a$.
- If |O| = k, we call O a k-orbit.
- Set of orbits of a permutation induce a partition of [n]

Let $\{O\}_{O \in \mathcal{O}}$ be the orbit/cycle decomposition of π . For $O \in \mathcal{O}$,

$$\mathsf{Z}_O \triangleq \prod_{i \in O} \mathsf{Z}_i \implies \prod_{i=1}^n \mathsf{Z}_i = \prod_{O \in \mathcal{O}} \mathsf{Z}_O$$

The random variables $\{Z_O\}_O$ are independent (under $\mathbb{P}_{\mathcal{H}_0}$),

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi}\mathbb{E}_{\mathcal{H}_0}\left[\prod_{i=1}^n \mathsf{Z}_i\right] = \mathbb{E}_{\pi}\mathbb{E}_{\mathcal{H}_0}\left[\prod_{O\in\mathcal{O}}\mathsf{Z}_O\right] = \mathbb{E}_{\pi}\prod_{O\in\mathcal{O}}\mathbb{E}_{\mathcal{H}_0}[\mathsf{Z}_O].$$

For a fixed orbit O of a permutation π ,

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{Z}_O] = \frac{1}{(1-\rho^{2|O|})^d}.$$

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If $N_k(\pi)$ is the number of k-orbits of π , then

$$\mathbb{E}_0[\mathsf{L}_n^2] = \mathbb{E}_{\pi}\left[\prod_C \mathbb{E}_{\mathcal{H}_0}[\mathsf{Z}_O]\right] = \mathbb{E}_{\pi}\left[\prod_{k=1}^n \frac{1}{(1-\rho^{2k})^{d \cdot N_k}}\right]$$

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Use statistical properties of k-orbits of $\pi \sim \text{Unif}(\mathbb{S}_n)$.

.

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In particular, [Arratia, Tavaré'92]

$$d_{\mathsf{TV}}\left(\mathcal{L}\left(N_1, N_2, \dots, N_k\right), \mathcal{L}\left(P_1, P_2, \dots, P_k\right)\right) \leq F\left(\frac{n}{k}\right),$$

for any $1 \le k \le n$, and $\{P_i\}_{i=1}^n$ independent sequence with $P_i \sim \text{Poisson}(i^{-1})$, and $\log F(x) = -x \log x(1 + o(1))$ as $x \to \infty$

In the Poisson world, for any m,

$$\mathbb{E}\left[\prod_{k=1}^m \frac{1}{(1-\rho^{2k})^{d \cdot P_k}}\right] \le \exp\left(\frac{d\rho^2}{1-\rho^2} + \frac{c(d,\rho^2)\rho^4}{1-\rho^4}\right)$$

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Decompose,

$$\prod_{k=1}^{n} \frac{1}{(1-\rho^{2k})^{d \cdot N_k}} = \prod_{k=1}^{\lceil \log n \rceil} \frac{1}{(1-\rho^{2k})^{d \cdot N_k}} \prod_{k=\lceil \log n \rceil+1}^{n} \frac{1}{(1-\rho^{2k})^{d \cdot N_k}}$$

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For the tail $(m = \lceil \log n \rceil)$,

$$\prod_{k=m+1}^{n} \frac{1}{(1-\rho^{2k})^{d \cdot N_k}} \le \left(\frac{1}{1-\rho^{2m}}\right)^{d \sum_{k=m}^{n} N_k} = \left(\frac{1}{1-\rho^{2m}}\right)^{dn} \le \exp\left(\frac{dn\rho^{2m}}{1-\rho^{2m}}\right) = 1 + o(1),$$

for $d\rho^2 = o(1)$.

In the Poisson world, for any m,

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Thus,

$$\prod_{k=1}^{n} \frac{1}{(1-\rho^{2k})^{d \cdot N_k}} = (1+o(1)) \cdot \prod_{k=1}^{\lceil \log n \rceil} \frac{1}{(1-\rho^{2k})^{d \cdot N_k}}$$

In the Poisson world, for any m,

$$\mathbb{E}\left[\prod_{k=1}^m \frac{1}{(1-\rho^{2k})^{d \cdot P_k}}\right] \le \exp\left(\frac{d\rho^2}{1-\rho^2} + \frac{c(d,\rho^2)\rho^4}{1-\rho^4}\right)$$

Now,

$$\mathbb{E}_{\pi}\left[\prod_{k=1}^{m} \left(\frac{1}{1-\rho^{2k}}\right)^{dN_{k}}\right] \leq \mathbb{E}_{\pi}\left[\prod_{k=1}^{m} \left(\frac{1}{1-\rho^{2k}}\right)^{dP_{k}}\right] + d_{\mathsf{TV}}\left(\mathcal{L}\left(N_{1}^{m}\right), \mathcal{L}\left(P_{1}^{m}\right)\right) \cdot \left(\frac{1}{1-\rho^{2}}\right)^{dn}$$

In the Poisson world, for any m,

$$\mathbb{E}\left[\prod_{k=1}^m \frac{1}{(1-\rho^{2k})^{d \cdot P_k}}\right] \le \exp\left(\frac{d\rho^2}{1-\rho^2} + \frac{c(d,\rho^2)\rho^4}{1-\rho^4}\right)$$

Now,

$$\mathbb{E}_{\pi} \left[\prod_{k=1}^{m} \left(\frac{1}{1 - \rho^{2k}} \right)^{dN_k} \right] \leq \mathbb{E}_{\pi} \left[\prod_{k=1}^{m} \left(\frac{1}{1 - \rho^{2k}} \right)^{dP_k} \right] + d_{\mathsf{TV}} \left(\mathcal{L} \left(N_1^m \right), \mathcal{L} \left(P_1^m \right) \right) \cdot \left(\frac{1}{1 - \rho^2} \right)^{dn} \leq \exp \left(\frac{d\rho^2}{1 - \rho^2} + \frac{c(d, \rho^2)\rho^4}{1 - \rho^4} \right) + F \left(\frac{n}{\lceil \log n \rceil} \right) \left(\frac{1}{1 - \rho^2} \right)^{dn} = 1 + o(1),$$

if $d\rho^2 = o(1)$.

When d is fixed, and $n \to \infty,$ the above technique gives

Theorem (Impossibility)

Strong detection is impossible if $d < \frac{\log(\rho^2)}{\log(1-\rho^2)}$.

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Consider the simple case of d = 1,



When d is fixed, and $n \to \infty$, the above technique gives

Theorem (Impossibility)

Strong detection is impossible if $d < \frac{\log(\rho^2)}{\log(1-\rho^2)}$.

Lower bound: for d = 1, we get the condition $\rho^2 < 1/2$.

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What is the source for this significant gap? Computational?

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What is the source for this significant gap? Computational?

Not clear yet! But, we can prove a better lower bound.

We have the following result.

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Theorem (Impossibility for d = 1)
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Strong detection is impossible for any $\rho^2 < 1$.

Proof sketch: We use polynomial decomposition:

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- We want to analyze $\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2]$.
- Idea: decompose L_n into its orthogonal components.
- Univariate Hermite polynomials: for $k \in \mathbb{N}$,

$$h_k(x) \triangleq (-1)^k e^{x^2/2} \frac{d^k}{dx^k} e^{-x^2/2},$$

are orthonormal w.r.t. the standard Gaussian measure,

$$\mathbb{E}_{X \sim \sim N(0,1)} \left[h_k(X) h_\ell(X) \right] = \delta[k - \ell].$$

Proof sketch: We use polynomial decomposition:

- We want to analyze $\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2]$.
- Idea: decompose L_n into its orthogonal components.
- Multivariate Hermite polynomials:

Let $H_{\theta}(x) = \prod_{i=1}^{n} h_{\theta_i}(x_i)$ for $\theta \in \mathbb{N}^n$, and it holds

 $\mathbb{E}_{\mathsf{X} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})} \left[H_{\alpha}(\mathsf{X}) H_{\gamma}(\mathsf{X}) \right] = \delta[\alpha - \gamma].$

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$$\mathbb{E}_{\mathsf{X} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})} \left[H_{\alpha}(\mathsf{X}) H_{\gamma}(\mathsf{X}) \right] = \delta[\alpha - \gamma].$$

• Form a complete orthonormal system in $L^2(\mathcal{H}_0)$,

$$\mathsf{L}_{n}(\mathsf{X},\mathsf{Y}) = \sum_{\alpha,\beta\in\mathbb{N}^{n}} \langle H_{\alpha,\beta}(\mathsf{X},\mathsf{Y}),\mathsf{L}_{n}(\mathsf{X},\mathsf{Y})\rangle_{\mathcal{H}_{0}} H_{\alpha,\beta}(\mathsf{X},\mathsf{Y}),$$

where $H_{\alpha,\beta}(\mathsf{X},\mathsf{Y}) \triangleq H_{\alpha}(\mathsf{X})H_{\beta}(\mathsf{Y})$, and

$$\langle \phi, \psi \rangle_{\mathcal{H}_0} \triangleq \mathbb{E}_{\mathcal{H}_0} \left[\psi(\mathsf{X}, \mathsf{Y}) \cdot \phi(\mathsf{X}, \mathsf{Y}) \right].$$

Proof sketch: We use polynomial decomposition:

- We want to analyze $\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2]$.
- Idea: decompose L_n into its orthogonal components.
- Parseval's identity,

$$\mathbb{E}_{\mathcal{H}_0}\left[\mathsf{L}_n^2\right] = \|\mathsf{L}_n\|_{\mathcal{H}_0}^2 = \sum_{\alpha,\beta\in\mathbb{N}^n} \langle H_{\alpha,\beta}(\mathsf{X},\mathsf{Y}),\mathsf{L}_n(\mathsf{X},\mathsf{Y})\rangle_{\mathcal{H}_0}^2$$

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It can be shown that

$$\langle H_{\alpha,\beta}(\mathsf{X},\mathsf{Y}),\mathsf{L}_n(\mathsf{X},\mathsf{Y})\rangle_{\mathcal{H}_0} = \rho^{|\alpha|} \cdot \mathbb{P}[\pi(\beta) = \alpha]$$

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Goal: find $\mathbb{P}[\pi(\beta) = \alpha]$.

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Then,

$$\mathbb{P}[\pi(\beta) = \alpha] = \frac{1}{|[\alpha]|} \mathbb{1}_{\alpha \equiv \beta}.$$

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• We want to analyze $\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2]$.

• Thus,

$$\begin{split} \mathbb{E}_{\mathcal{H}_0} \left[\mathsf{L}_n^2 \right] &= \sum_{\alpha, \beta \in \mathbb{N}^n} \left\langle H_{\alpha, \beta}(\mathsf{X}, \mathsf{Y}), \mathsf{L}_n(\mathsf{X}, \mathsf{Y}) \right\rangle_{\mathcal{H}_0}^2 \\ &= \sum_{\alpha, \beta \in \mathbb{N}^n} \rho^{2|\alpha|} \frac{1}{|[\alpha]|^2} \mathbb{1}_{\alpha \equiv \beta} \\ &= \sum_{m=0}^{\infty} |\{[\alpha] : |\alpha| = m\}| \cdot \rho^{2m} \\ &= \sum_{m=0}^{\infty} |\mathsf{Par}(m, \leq_n)| \cdot \rho^{2m} \end{split}$$

where $\mathsf{Par}(m,\leq_n)$ is the set of integer partitions of m to at most n elements.

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where $|\operatorname{Par}(m, \leq_{\infty})|$ is the number of integer partitions of the number m.

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• We want to analyze $\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2]$.

$$\mathbb{E}_{\mathcal{H}_0}\left[\mathsf{L}_n^2\right] \le \sum_{m=0}^{\infty} |\mathsf{Par}(m, \le_{\infty})| \cdot \rho^{2m} \quad (\star)$$

• Hardy-Ramanujan Formula: $\exists c > 0$, s.t.

$$|\mathsf{Par}(m, \leq_{\infty})| \le c \cdot \frac{1}{4\sqrt{3}m} \exp\left(\pi \sqrt{\frac{2m}{3}}\right)$$

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Thus, |Par(m, ≤_∞)| is sub-exponential in m, and hence (⋆) converges to a finite number, for any ρ² < 1.

Theorem (Impossibility for $d \in \mathbb{N}$)

Strong detection is impossible for any $d\rho^2 < 1$.

This is proved using the same techniques ending up with complicated high-dimensional distribution functions.

• What if the databases are not Gaussian?

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- Consider the following detection problem:

$$\mathcal{H}_0: (\mathsf{X}_1, \mathsf{Y}_1), \dots, (\mathsf{X}_n, \mathsf{Y}_n) \stackrel{\text{i.i.d}}{\sim} P_X^{\otimes d} \times P_Y^{\otimes d}$$
$$\mathcal{H}_1: (\mathsf{X}_1, \mathsf{Y}_{\pi_1}), \dots, (\mathsf{X}_n, \mathsf{Y}_{\pi_n}) \stackrel{\text{i.i.d}}{\sim} P_{XY}^{\otimes d},$$

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Theorem (Impossibility of weak detection) Weak detection is impossible if

$$d \cdot \chi^2(P_{XY} || Q_{XY}) = o(1).$$

where $\chi^2(\mathbb{P}||\mathbb{Q}) = \int \frac{\mathrm{d}\mathbb{P}^2}{\mathrm{d}\mathbb{Q}} - 1.$

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Theorem (Possibility of strong detection) *If*

$$d \cdot \frac{d_{\mathsf{SKL}}^2(P_{XY} || Q_{XY})}{\mathsf{Var}_{Q_{XY}}\left(\mathcal{K}(A,B)\right)} = \omega(1)$$

then, $\mathsf{R}(\phi_{\mathsf{sum}}) \to 0$, as $d \to \infty$.

Proof sketch:

• For any $x \in \mathcal{X}$ and $y \in \mathcal{Y}$, we let $\mathcal{L}(x, y) \triangleq \frac{P_{XY}(x, y)}{Q_{XY}(x, y)}$.

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- Recall that

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{L}_n^2] = \mathbb{E}_{\pi} \left[\prod_{O \in \mathcal{O}} \mathbb{E}_{\mathcal{H}_0}[\mathsf{Z}_O] \right].$$

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• Then, with the notation above, it can be shown that,

$$\mathbb{E}_{\mathcal{H}_0}[\mathsf{Z}_C] = \left(\sum_{i \in \mathbb{N}} \lambda_i^{2|C|}\right)^d$$

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• Substituting, massaging, it can be shown that weak detection is impossible if

$$d \cdot \sum_{i \ge 1} \frac{\lambda_i^2}{1 - \lambda_i^2} = o(1).$$

• What if the databases are only partially correlated?

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- Consider the following detection problem:

where $\pi \sim \text{Unif}(\mathbb{S}_n)$ and $\mathcal{K} \sim \text{Unif}\binom{[n]}{k}$.

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$$\mathcal{H}_{0}: (\mathsf{X}_{1}, \mathsf{Y}_{1}), \dots, (\mathsf{X}_{n}, \mathsf{Y}_{n}) \stackrel{\text{i.i.d}}{\sim} \mathcal{N}^{\otimes d}(0, I_{2 \times 2})$$
$$\mathcal{H}_{1}: \begin{cases} \{(\mathsf{X}_{i}, \mathsf{Y}_{\pi_{i}})\}_{i \in \mathcal{K}} \stackrel{\text{i.i.d}}{\sim} \mathcal{N}^{\otimes d}(\mathbf{0}, \Sigma_{\rho}) \\ \{(\mathsf{X}_{i}, \mathsf{Y}_{\pi_{i}})\}_{i \notin \mathcal{K}} \stackrel{\text{i.i.d}}{\sim} \mathcal{N}^{\otimes d}(\mathbf{0}, \mathbf{I}_{2 \times 2}) \\ \{(\mathsf{X}_{i}, \mathsf{Y}_{\pi_{i}})\}_{i \notin \mathcal{K}} \coprod \{(\mathsf{X}_{i}, \mathsf{Y}_{\pi_{i}})\}_{i \in \mathcal{K}} \end{cases}$$

where $\pi \sim \mathsf{Unif}(\mathbb{S}_n)$ and $\mathcal{K} \sim \mathsf{Unif}\binom{[n]}{k}$.

• So, only a planted set ${\mathcal K}$ of $k \leq n$ "users" is common to the two databases.

Theorem (Impossibility weak detection) *If*,

$$\left(\frac{k}{n}\right)^2 \left(\prod_{i=1}^k \frac{1}{1 - (d\rho^2)^i} - 1\right) = o(1),$$

then weak detection is impossible.

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then weak detection is impossible.

For example, if $k = O(\log n)$, then we get

$$\rho^2 < \frac{1}{d} \left[1 - \left(\mathsf{C}\frac{k}{n}\right)^{\frac{2}{k}} \right],$$

and we note that $(k/n)^{\frac{2}{k}} = o(1)$.

• In many modern applications, the observations may be in the form of graphs.

• Consider the following detection problem:

$$\begin{split} \mathcal{H}_{0} &: \left(\mathsf{A}_{ij}, \mathsf{B}_{ij}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{Q}_{\mathsf{A}\mathsf{B}} = \mathcal{P}_{\mathsf{A}} \times \mathcal{P}_{\mathsf{B}} \\ \mathcal{H}_{1} &: \left(\mathsf{A}_{ij}, \mathsf{B}_{\pi_{i}\pi_{j}}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{P}_{\mathsf{A}\mathsf{B}} | \pi \sim \mathsf{Unif}(\mathbb{S}_{n}), \end{split}$$

where $\mathcal{P}_A = \mathcal{P}_B$.

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where $\mathcal{P}_A = \mathcal{P}_B$.

 The Bernoulli case was analyzes thoroughly in the literature^a, both from the statistical and computational point of views! Here, P_A = P_B = Bernoulli(τp), for some p ∈ (0, 1) and τ ∈ [0, 1]. Under P_{AB}, we have A ~ Bernoulli(τp), and

$$\mathsf{B}|\mathsf{A} \sim \begin{cases} \mathsf{Bernoulli}(\tau), & \text{if } X = 1\\ \mathsf{Bernoulli}\left(\frac{\tau p(1-\tau)}{1-\tau p}\right), & \text{if } X = 0. \end{cases}$$

^aE.g., [Wu,Xu,Yu'21], [Ding,Du,'23], [Ding,Du,Li'23].

• Consider the following detection problem:

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 The Gaussian case was studied from the statistical point of view [Wu,Xu,Yu'21].

• Consider the following detection problem:

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where $\mathcal{P}_A = \mathcal{P}_B$.

Theorem (Impossibility of weak detection)

Weak detection is statistically impossible if

$$\chi^{2}\left(\mathcal{P}||\mathcal{Q}\right) \leq \frac{(2-\epsilon)\log n}{\alpha n}, \text{ and} \\ d_{\mathsf{KL}}\left(\mathcal{P}||\mathcal{Q}\right) + \delta_{n} \cdot \mathsf{Var}_{\mathcal{P}}\left(\log \mathcal{L}\right) \leq \frac{(2-\epsilon)\log n}{n},$$

for any $\omega(1) = \delta_n = o(\log n)$, and any constant $\epsilon > 0$.

• Consider the following detection problem:

$$\begin{split} \mathcal{H}_{0} &: \left(\mathsf{A}_{ij},\mathsf{B}_{ij}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{Q}_{\mathsf{A}\mathsf{B}} = \mathcal{P}_{\mathsf{A}} \times \mathcal{P}_{\mathsf{B}} \\ \mathcal{H}_{1} &: \left(\mathsf{A}_{ij},\mathsf{B}_{\pi_{i}\pi_{j}}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{P}_{\mathsf{A}\mathsf{B}} | \pi \sim \mathsf{Unif}(\mathbb{S}_{n}), \end{split}$$

where $\mathcal{P}_A = \mathcal{P}_B$.

• For the class of distributions for which there is a constant C > 1 such that $\chi^2(\mathcal{P}||\mathcal{Q}) \leq C \cdot d_{\mathsf{KL}}(\mathcal{P}||\mathcal{Q})$, weak detection is impossible if

$$d_{\mathsf{KL}}\left(\mathcal{P}||\mathcal{Q}\right) \leq \frac{(2-\epsilon)\log n}{n}$$

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• Coincides with [Wu,Xu,Yu'21] for Bernoulli and Gaussian.

• Consider the following detection problem:

$$\begin{split} \mathcal{H}_{0} &: \left(\mathsf{A}_{ij},\mathsf{B}_{ij}\right) \stackrel{\mathrm{i.i.d.}}{\sim} \mathcal{Q}_{\mathsf{A}\mathsf{B}} = \mathcal{P}_{\mathsf{A}} \times \mathcal{P}_{\mathsf{B}} \\ \mathcal{H}_{1} &: \left(\mathsf{A}_{ij},\mathsf{B}_{\pi_{i}\pi_{j}}\right) \stackrel{\mathrm{i.i.d.}}{\sim} \mathcal{P}_{\mathsf{A}\mathsf{B}} | \pi \sim \mathsf{Unif}(\mathbb{S}_{n}), \end{split}$$

where $\mathcal{P}_A = \mathcal{P}_B$.

Theorem (Strong detection upper bound)

Suppose there is a $\bar{\theta}\in(-d_{\mathsf{KL}}(\mathcal{Q}||\mathcal{P}),d_{\mathsf{KL}}(\mathcal{P}||\mathcal{Q}))$ with

$$E_{\mathcal{Q}}(\bar{\theta}) \ge \frac{2\log(n/e)}{n-1} + O(n^{-2}\log n),$$

$$E_{\mathcal{P}}(\bar{\theta}) = \omega(n^{-2}).$$

Then, $\mathsf{R}_n(\phi_{\mathsf{GLRT}}) \to 0$, as $n \to \infty$.

• Consider the following detection problem:

$$\begin{split} \mathcal{H}_{0} &: \left(\mathsf{A}_{ij},\mathsf{B}_{ij}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{Q}_{\mathsf{A}\mathsf{B}} = \mathcal{P}_{\mathsf{A}} \times \mathcal{P}_{\mathsf{B}} \\ \mathcal{H}_{1} &: \left(\mathsf{A}_{ij},\mathsf{B}_{\pi_{i}\pi_{j}}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{P}_{\mathsf{A}\mathsf{B}} | \pi \sim \mathsf{Unif}(\mathbb{S}_{n}), \end{split}$$

where $\mathcal{P}_A = \mathcal{P}_B$.

• For pairs of distributions $(\mathcal{P}, \mathcal{Q})$ with sub-exponential likelihood function, strong detection is possible if

$$d_{\mathsf{KL}}\left(\mathcal{P}||\mathcal{Q}\right) \ge \frac{2\log n}{n-1}$$

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• Complements lower bound.

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- Complements lower bound.
- GLRT is exhibits exponential computational complexity. What about poly-time algorithms?

• Consider the following detection problem:

$$\begin{split} \mathcal{H}_{0} &: \left(\mathsf{A}_{ij},\mathsf{B}_{ij}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{Q}_{\mathsf{A}\mathsf{B}} = \mathcal{P}_{\mathsf{A}} \times \mathcal{P}_{\mathsf{B}} \\ \mathcal{H}_{1} &: \left(\mathsf{A}_{ij},\mathsf{B}_{\pi_{i}\pi_{j}}\right) \stackrel{\text{i.i.d.}}{\sim} \mathcal{P}_{\mathsf{A}\mathsf{B}} | \pi \sim \mathsf{Unif}(\mathbb{S}_{n}), \end{split}$$

where $\mathcal{P}_A = \mathcal{P}_B$.

Theorem (Weak detection upper bound)

$$|If| |\operatorname{corr}(\mathcal{Q}, \mathcal{P})| \triangleq \frac{|\operatorname{cov}_{\mathcal{P}}(A, B)|}{\operatorname{Var}_{\mathcal{Q}}(A)} = \Omega(1), \text{ and }$$

$$\frac{\mathbb{E}_{\mathcal{Q}}|A-B|^3}{\mathsf{Var}_{\mathcal{Q}}^{3/2}(A)}, \frac{\mathbb{E}_{\mathcal{P}}|A-B|^3}{\mathsf{Var}_{\mathcal{Q}}^{3/2}(A)(1-|\mathsf{corr}(\mathcal{Q},\mathcal{P})|)^{3/2}} = o(n),$$

then $\lim_{n\to\infty} \mathsf{R}_n(\phi_{\mathsf{sum}}) < 1$.

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- **Conjecture**: this is fundamental in the sense that this is a barrier for what can be achieved using polynomial-time algorithms.
- In the Bernoulli case [Ding,Du,Li'23] prove computational lower bound based on the low-degree polynomial conjecture.
Testing Dependency of Random Graphs [Oren, Paslev, H'24]

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Theorem (LDP computational lower bound)

For a certain class of pairs of distributions $(\mathcal{P}, \mathcal{Q})$, if $|\operatorname{corr}(\mathcal{Q}, \mathcal{P})| = o(1)$, then $\|L_{n, \leq D}\|_{\mathcal{H}_0} \leq O(1)$, for any $D = O(|\operatorname{corr}(\mathcal{Q}, \mathcal{P})|^{-1})$.

Here, $L_{n,\leq D}$ is the projection of L_n to the linear subspace of polynomials of degree at most $D \in \mathbb{N}$.

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Open Problems:

- Close the gap, and obtain sharp bounds.
- Prove existence of/close the computational gaps.

Thank You!