



Computational Time Series Analysis (IPAM, Part II, nonstationary methods)

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and

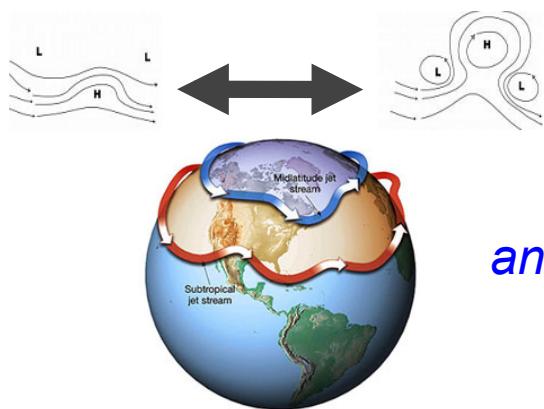
Institute of Computational Science
Università della Svizzera Italiana, Lugano
Switzerland



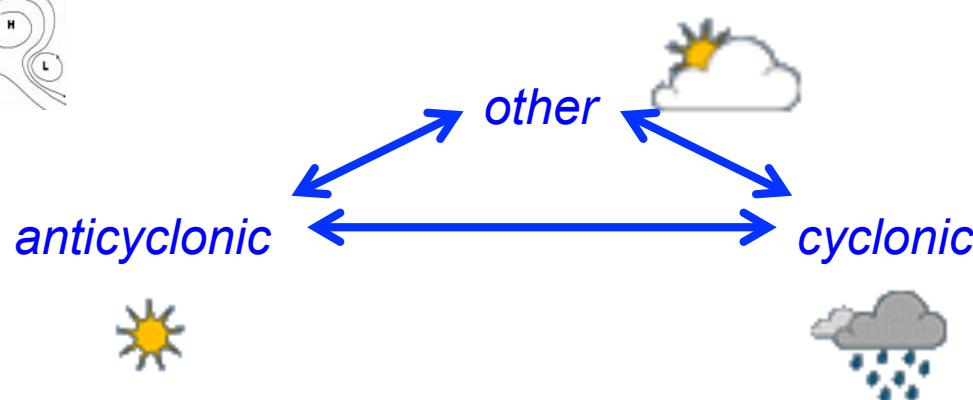
Università
della
Svizzera
italiana

Motivation (I): Data

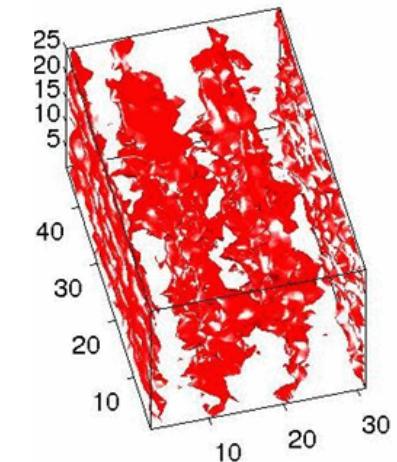
Inverse problem: *data*-based identification of *model* parameters



**time series of vectors
(continuous, finite dim.)**



**time series of regimes
(discrete, finite dim.)**



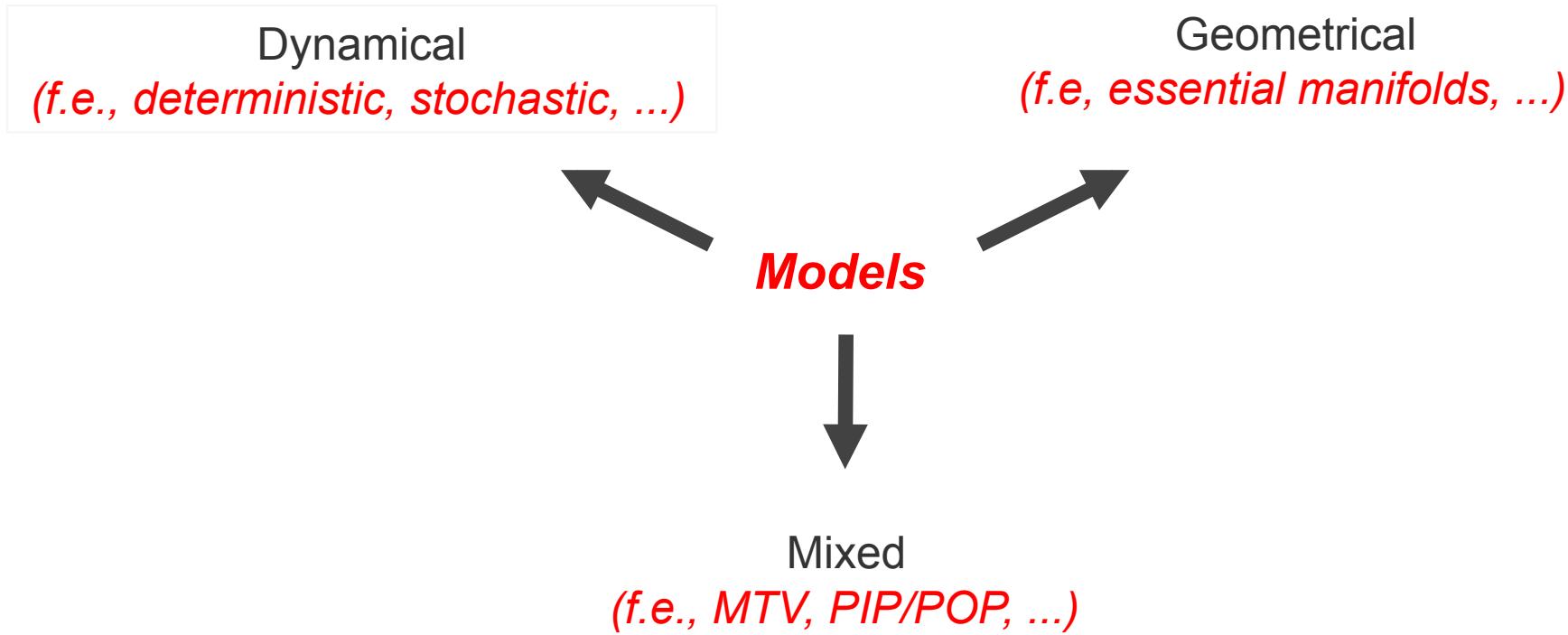
**time series of functions
(continuous, infin. dim.)**

Challenges:

**VERY LARGE (\sim Tb size) , non-Markovian (memory),
non-stationary (trends), multidimensional, ext. influence**

Motivation (II): Models

Inverse problem: *data*-based identification of *model* parameters



Challenges:

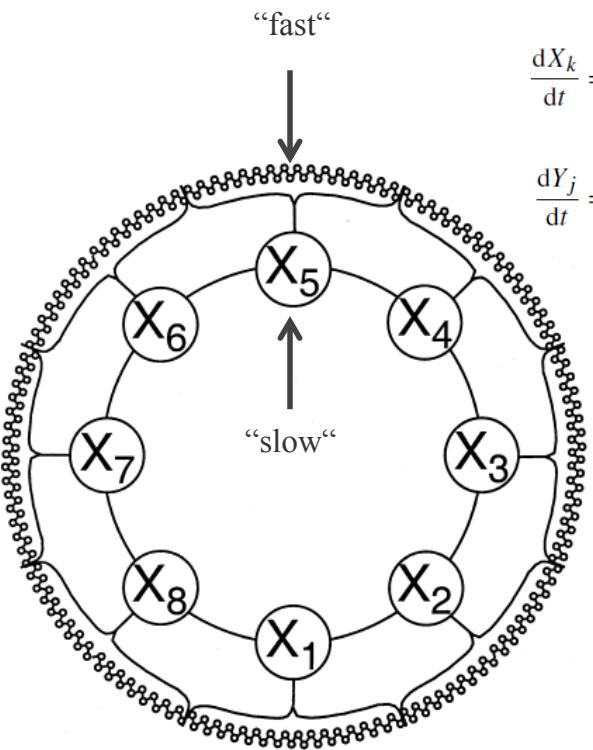
existence and uniqueness, robustness, conditioning

Yesterday (stationary case)

- “eagle eye“ perspective on **stochastic processes** from the viewpoint of **deterministic dynamical systems**
- geometric model inference: **EOF/SSA**
- multivariate dynamical model inference: **VARX**
- handling the ill-posed problem
- motivation for tomorrow : examples where stationarity assumption does not work

Lorenz96: “order zero“ atmosph. model

Wilks, Quart.J.Royal Met.Soc., 2005



$$\frac{dX_k}{dt} = -X_{k-1}(X_{k-2} - X_{k+1}) - X_k + F - \frac{hc}{b} \sum_{j=J(k-1)+1}^{kJ} Y_j; \quad k = 1, \dots, K$$

$$\frac{dY_j}{dt} = -cbY_{j+1}(Y_{j+2} - Y_{j-1}) - cY_j + \frac{hc}{b} X_{\text{int}[(j-1)/J]+1}; \quad j = 1, \dots, JK.$$

**stationary model
(time-independent parameters)**

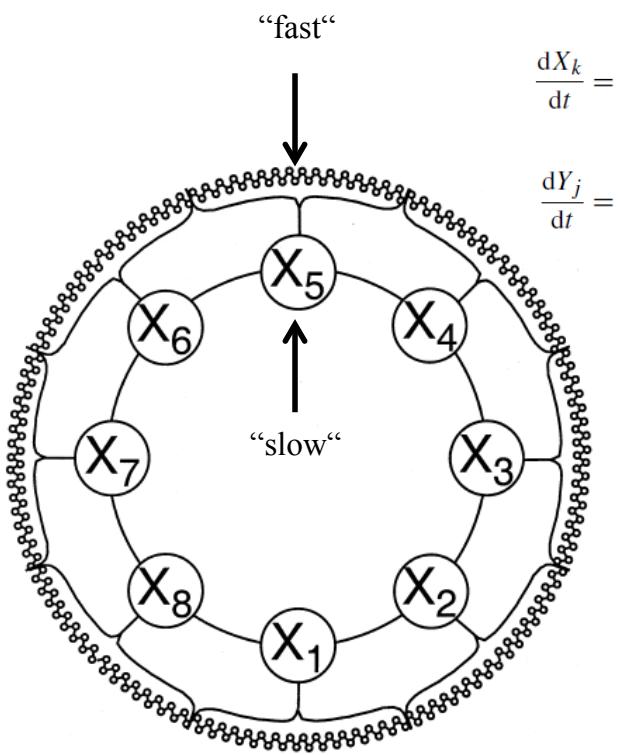
VARX (standard stationary stochastic Model)

$$y_t = \mu + \sum_{q=1}^m A_q y_{t-q\tau} + B\phi(x_t) + C\epsilon_t$$

- Lorenz, *Proc. Of. Sem. On Predict.*, 1996
- Majda/Timoffev/V.-Eijnden, *PNAS*, 1999
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**non-stationary model
(time-dependent parameters)**

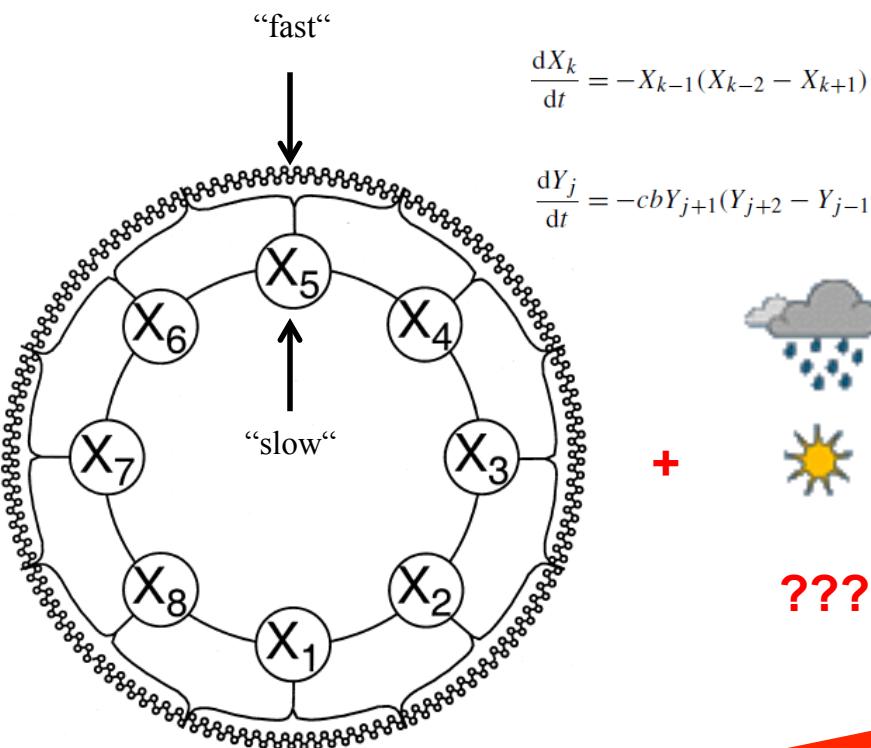
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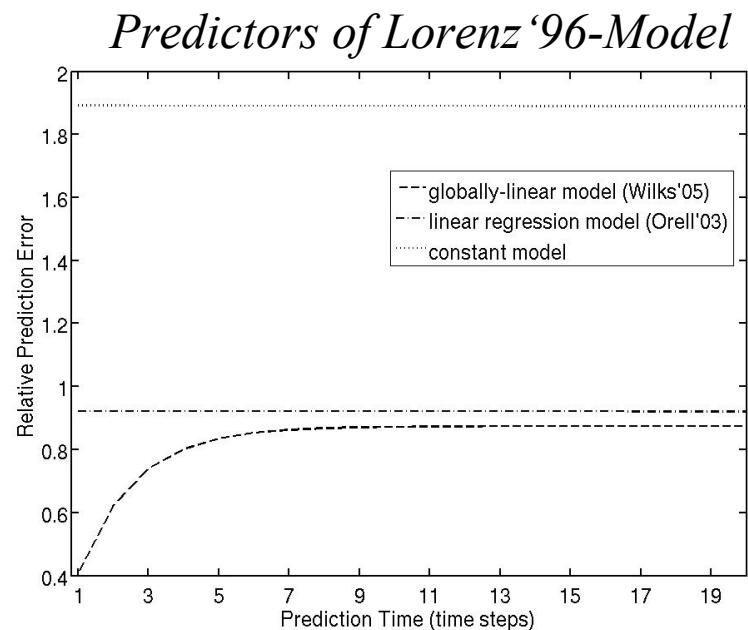
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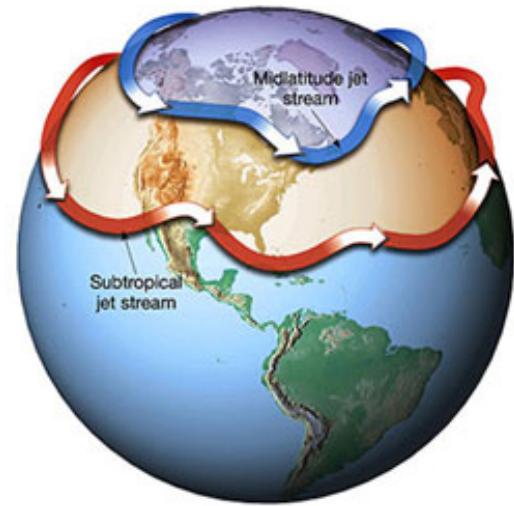
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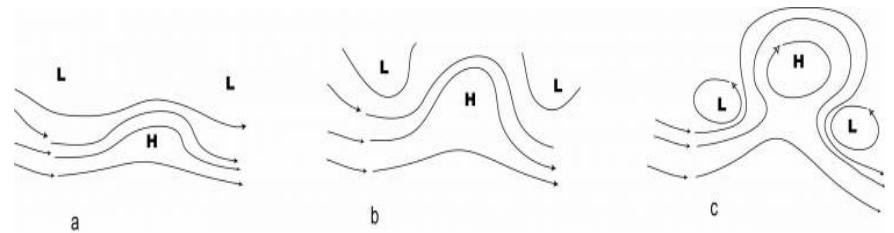


Analysis of geopotential data (Europe, 1958-2003, analysis of climate impact factors)

Wind Jets transport moisture from US to Europe

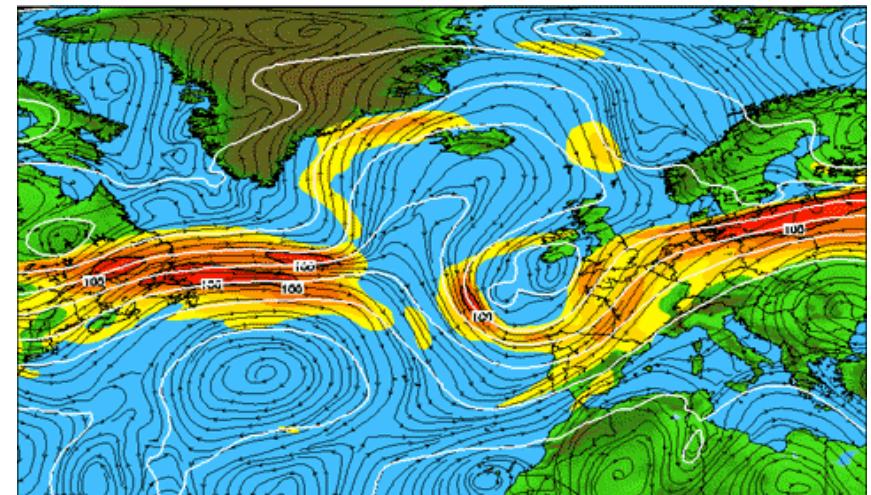


Jet-Blocking



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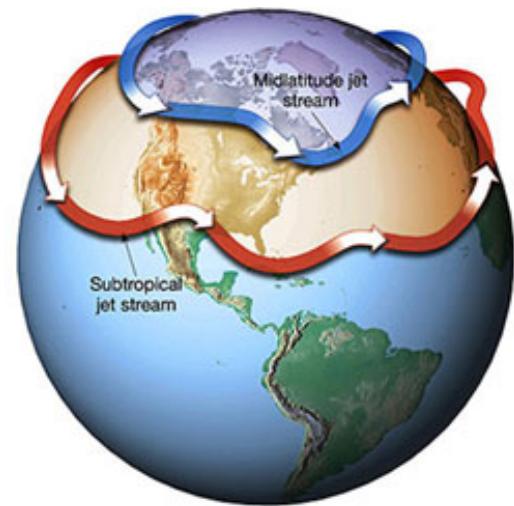


Jet-Blocking results in “indian summer” and “blackberry cold”

Analysis of ERA40 geopotential time series (PIK, Potsdam)

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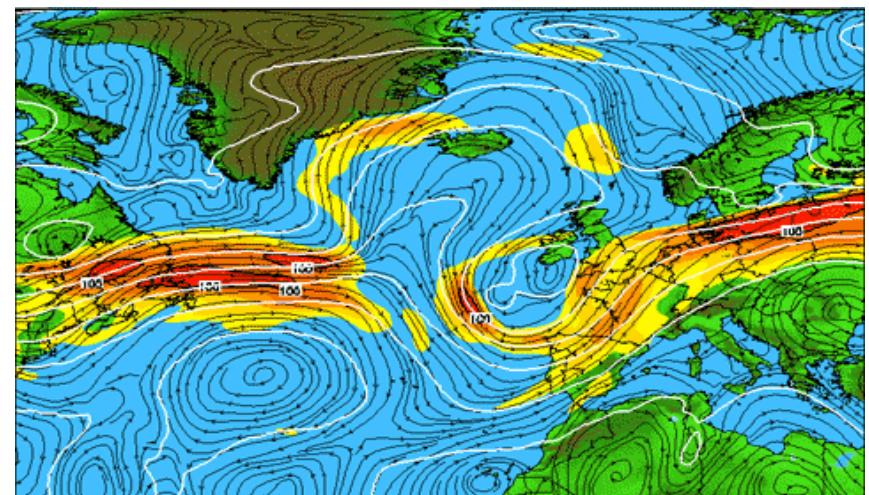
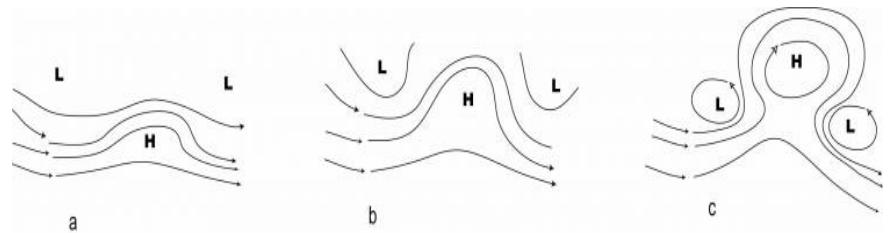
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Jet-Blocking



Standard statistics: seasonal factor, greenhouse gases and sun activity are all stat. insignificant!

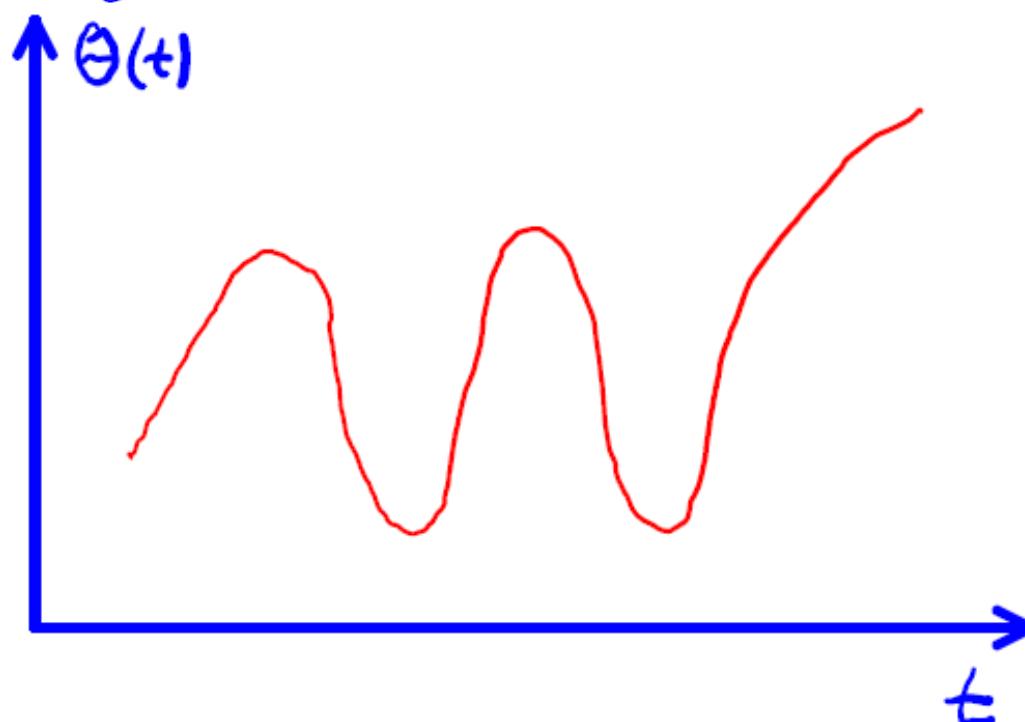
Non-stationary Processes

Direct problem:

$$F(x_t, \dots, x_{t-m\tau}, \theta(t), t) = 0$$

Inverse problem:

$$\int_0^T g(x_s, \theta(s)) ds \rightarrow \min_{\theta(t)}$$



Example:

$$x_t = \theta(t) + \epsilon_t \quad \epsilon_t \text{ (i.i.d.) } \mathbb{E} [\epsilon_t] = 0$$

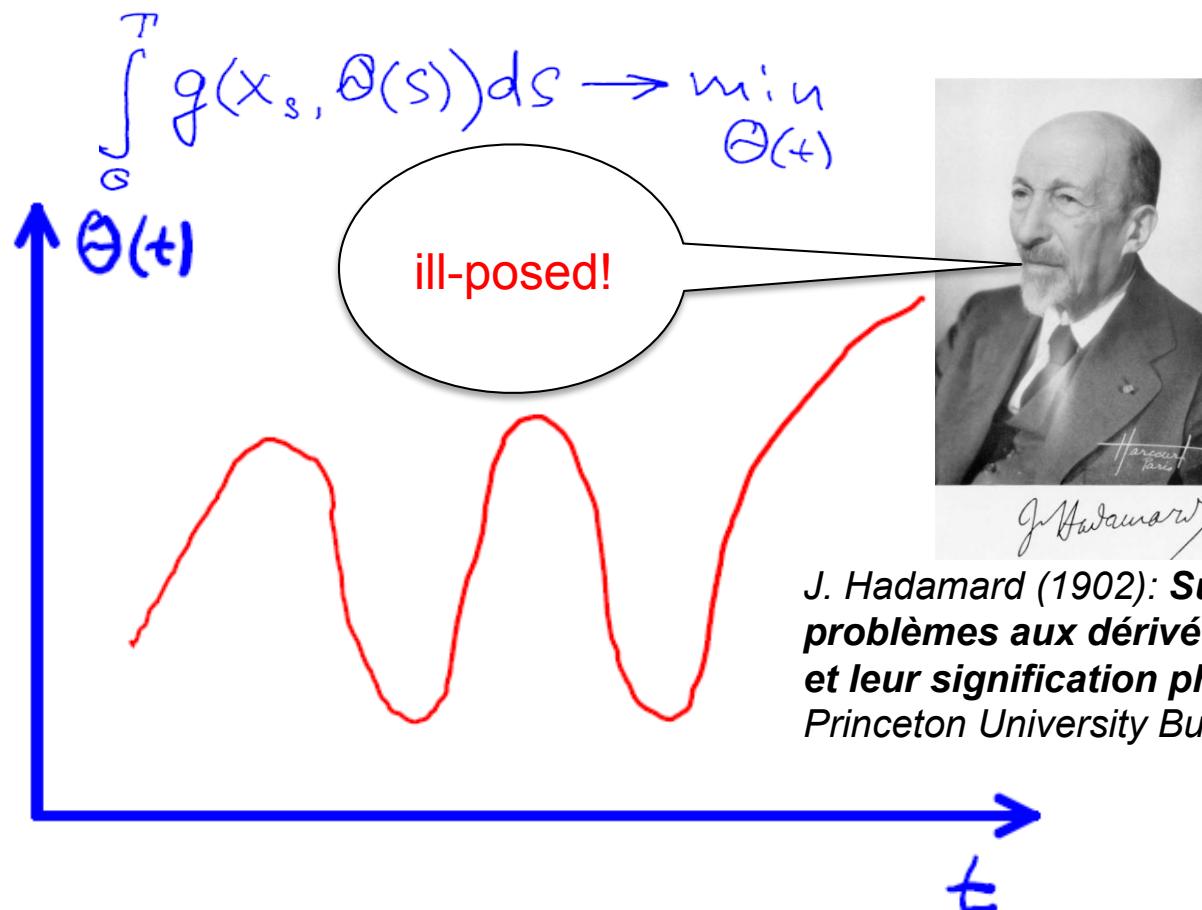
$$g(x_t, \theta(t)) = \|x_t - \theta_t\|_2$$

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III-posed problem

Example:

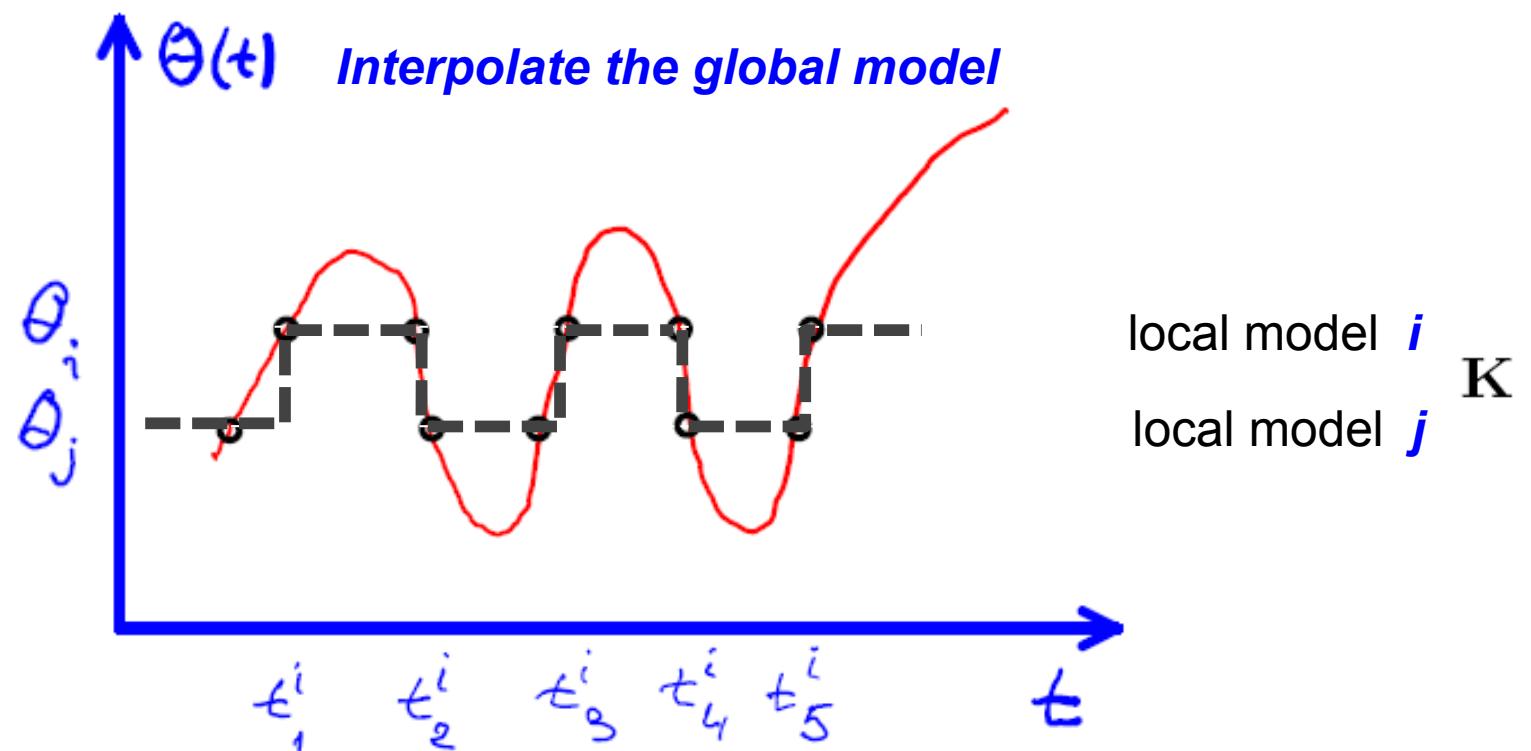
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Non-stationary Processes

$$g(x_t, \theta(t)) = \sum_{i=1}^K \gamma_i(t) g(x_t, \theta_i)$$

Local quality functional $g(x, \theta_i) : \Psi \times \Omega \rightarrow [0, \bar{g}]$, $0 < \bar{g} < +\infty$,



Averaged Clustering Functional

Find $\Gamma(t) = (\gamma_1(t), \dots, \gamma_{\mathbf{K}}(t))$ such that for each t :

$$\mathbf{L}(\Theta, \Gamma(t)) = \int_0^T \sum_{i=1}^{\mathbf{K}} \gamma_i(t) g(x_t, \theta_i) \rightarrow \min_{\Gamma(t), \Theta},$$

subjected to constraints:

$$\sum_{i=1}^{\mathbf{K}} \gamma_i(t) = 1, \quad \forall t \in [0, T]$$

$$\gamma_i(t) \geq 0, \quad \forall t \in [0, T], i = 1, \dots, \mathbf{K}$$

*Ill-posed problem:
Regularization !*

Incorporation of Temporal Information

Regularization : $\Theta(t)$ should be much “*slower*“ than x_t

$$\mathbf{L}(\Theta, \Gamma(t)) = \int_0^T \sum_{i=1}^K \gamma_i(t) g(x_t, \theta_i) \rightarrow \min_{\Gamma(t), \Theta},$$

$$\|\gamma_i(\cdot)\|_? \leq \|x_t\|_?$$
$$\begin{aligned}\sum_{i=1}^K \gamma_i(t) &= 1, \quad \forall t \in [0, T] \\ \gamma_i(t) &\geq 0, \quad \forall t \in [0, T], i = 1, \dots, K\end{aligned}$$

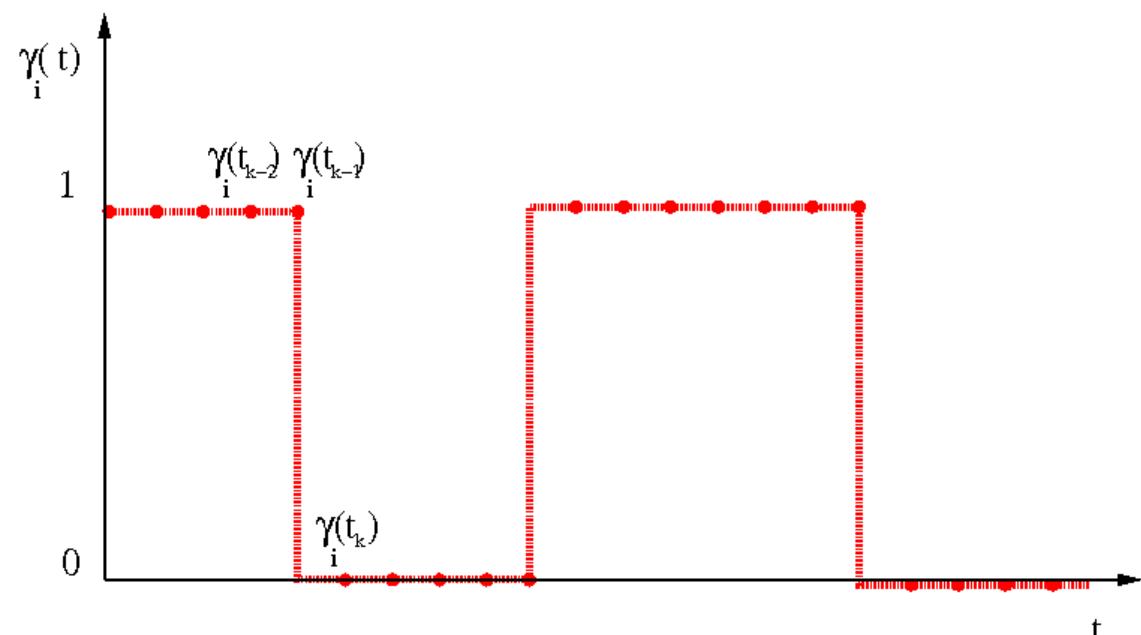
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H. 09, SIAM J. of Sci. Comp.
H. 09, accepted for publ. In JAS

Number Of Jumps
Time Interval

$$\|\cdot\|_{\mathcal{H}_1} : \sum_k \frac{(\gamma_i(t_{k+1}) - \gamma_i(t_k))^2}{\Delta t} \leq C_i$$

$$\|\cdot\|_{BV} : \sum_k \frac{|\gamma_i(t_{k+1}) - \gamma_i(t_k)|}{\Delta t} \leq C_i$$

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$$|\gamma_i|_{\mathcal{H}^1(0,T)} = \| \partial_t \gamma_i (\cdot) \|_{\mathcal{L}_2(0,T)} = \int_0^T (\partial_t \gamma_i(t))^2 dt \leq C_\epsilon^i < +\infty,$$

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A. N. Tikhonov
(<http://en.wikipedia.org>)

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Regularized clustering functional:

$$\mathbf{L}^\epsilon(\Theta, \Gamma(t), \epsilon^2) = \mathbf{L}(\Theta, \Gamma(t)) + \epsilon^2 \sum_{i=1}^K \int_0^T (\partial_t \gamma_i(t))^2 dt \rightarrow \min_{\Gamma(t), \Theta}$$

(H. 09, SIAM J. of Sci. Comp.)

FEM: Regularized Clustering Functional

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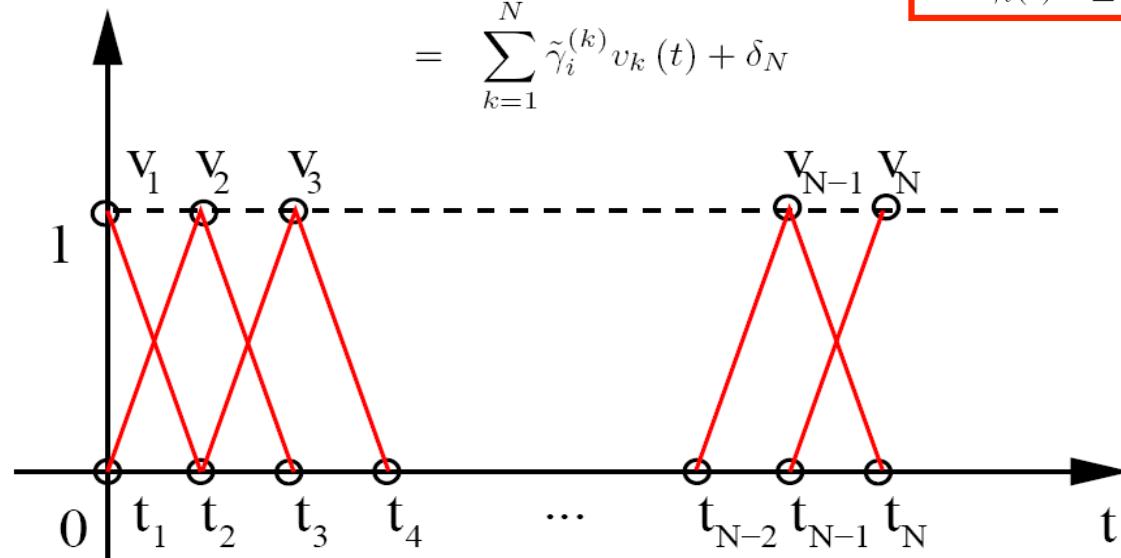
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(H. 09, SIAM J. of Sci. Comp.)

Galerkin-Ansatz :

$$\begin{aligned}\gamma_i(t) &= \tilde{\gamma}_i(t) + \delta_N \\ &= \sum_{k=1}^N \tilde{\gamma}_i^{(k)} v_k(t) + \delta_N\end{aligned}$$

$$\begin{aligned}\sum_{i=1}^K \gamma_i(t) &= 1, \quad \forall t \in [0, T] \\ \gamma_i(t) &\geq 0, \quad \forall t \in [0, T], i = 1, \dots, K\end{aligned}$$



Linear Finite Elements

FEM-Discretized Clustering Functional

$$\tilde{\mathbf{L}}^\epsilon = \sum_{i=1}^K [a^T(\theta_i) \bar{\gamma}_i + \epsilon^2 \bar{\gamma}_i^T \mathbf{H} \bar{\gamma}_i] \rightarrow \min_{\bar{\gamma}_i, \Theta}$$

subjected to

$$\sum_{i=1}^K \tilde{\gamma}_i^{(k+1)} = 1, \quad \forall k = 1, \dots, N,$$

$$\tilde{\gamma}_i^{(k+1)} \geq 0, \quad \forall k = 1, \dots, N; i = 1, \dots, K.$$

Iterative Subspace Minimization:
sparse QP can be used

where $a(\theta_i) = \left(\int_{t_1}^{t_2} v_1(t) g(x_t, \theta_i) dt, \dots, \int_{t_{N-1}}^{t_N} v_N(t) g(x_t, \theta_i) dt \right)$

- quadratic programming problem wrt $\Gamma(t)$ for a fixed Θ
- convex minim. problem wrt. Θ for a fixed $\Gamma(t)$

- Horenko, *SIAM Journal of Sci. Comp.*, 2008
- Horenko, *JAS*, 2009
- Horenko, *DYNAT*, 2009
- Franzke/Horenko/Majda/Klein, *JAS*, 2009

BV-case: making problem diff.-able

BV-constraint in the clustering functional minimization

$$\begin{aligned}\mathbf{L}(\Theta, \Gamma(t)) &= \int_0^T \sum_{i=1}^{\mathbf{K}} \gamma_i(t) g(x_t, \theta_i) \rightarrow \min_{\Gamma(t), \Theta}, \\ \sum_{i=1} \gamma_i(t) &= 1, \quad \forall t \in [0, T] \\ \gamma_i(t) &\geq 0, \quad \forall t \in [0, T], i = 1, \dots, \mathbf{K} \\ \|D\gamma_i^T\|_1 &\leq C \quad \forall i \quad v_i := D\gamma_i^T\end{aligned}$$

Theorem about the unique repr. of BV-func. (Moreau'88):

$$v_i = v_i^+ - v_i^-, \quad v_i^+ = \max(v_i, 0), v_i^- = \max(-v_i, 0)$$

Adaptive FEM-discretization:

$$\min_{x, \Theta} c^T(\Theta) x, \quad \text{subject to } A_{\text{eq}} x = b_{\text{eq}}, A_{\text{neq}} x \geq b_{\text{neq}}.$$

- linear programming problem wrt. x for a fixed Θ
- convex minim. problem wrt. Θ for a fixed x

Non-stationarity: Yes or No?

Information theory: golden ratio between numb. of parameters and “model quality”

(Ockham'XIVct.;Hartley'28;Shannon'48;Kolmogorov'68;Akaike'74;Tsai'98)



**How many angels can dance
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William of Ockham
(<http://en.wikipedia.org>)

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Let

$$G_{\mathbf{K}}(t) = \sum_{i=1}^{\mathbf{K}} \gamma_i(t) g(z_t, \theta_i) \quad M(\mathbf{K}) \text{ is the number of parameters}$$

Theorem 2.1 If for given \mathbf{K} and optimal Γ^* the $G_{\mathbf{K}}(t)$ are independently $\sigma_{\mathbf{K}} \chi_1^2$ distributed for some constant $\sigma_{\mathbf{K}}$, then the \mathbf{K} -dependent part of the Akaike Information Criterium is given by

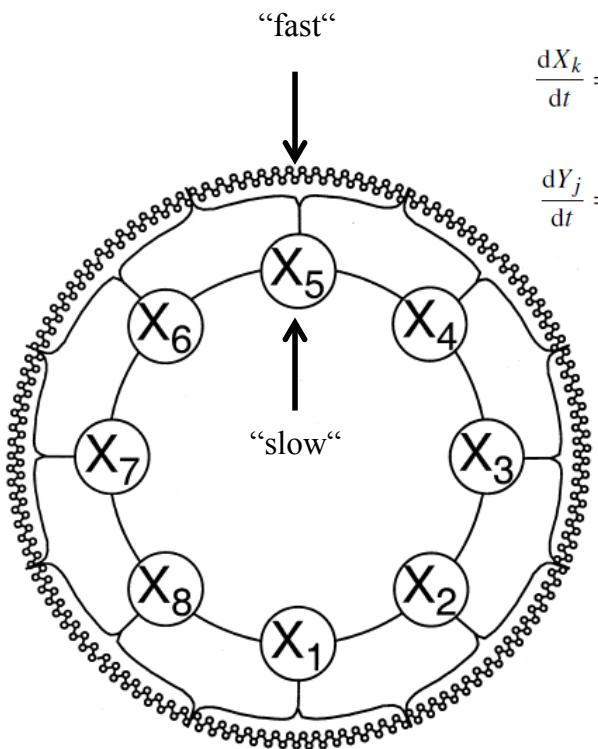
$$AIC = N \log L_{\mathbf{K}} + \frac{1}{2} \sum_{i=1}^N \log \frac{G_{\mathbf{K}}(t_i)}{L_{\mathbf{K}}} + M(\mathbf{K})$$

What does it mean:

Model Discrimination and identification of optimal \mathbf{K}

Example I: Lorenz96+forcing

Wilks, Quart.J.Royal Met.Soc., 2005



$$\frac{dX_k}{dt} = -X_{k-1}(X_{k-2} - X_{k+1}) - X_k + F - \frac{hc}{b} \sum_{j=J(k-1)+1}^{kJ} Y_j; \quad k = 1, \dots, K$$

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**non-stationary model
(time-dependent parameters)**

$$x_t = \mu(t) + \mathbf{A}(t) \phi_1(x_{t-\tau}, \dots, x_{t-m\tau}) + \mathbf{B}(t) \phi_2(u_t) + \mathbf{C}(t) \epsilon_t$$

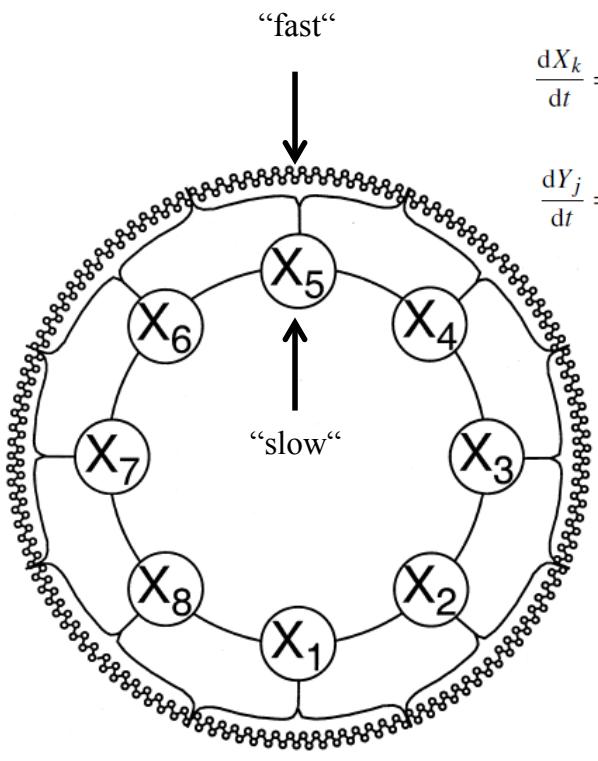
$$\mathbf{C}(t) = \mathbf{P}(t) \boldsymbol{\Lambda}(t)$$

$$g(x_t, \theta(t)) = \|x_t - \mu(t) - \mathbf{A}(t) \phi_1(x_{t-\tau}, \dots, x_{t-m\tau}) - \mathbf{B}(t) \phi_2(u(t))\|_{\mathbf{P}(t)}$$

(H. 10, JAS, in press)

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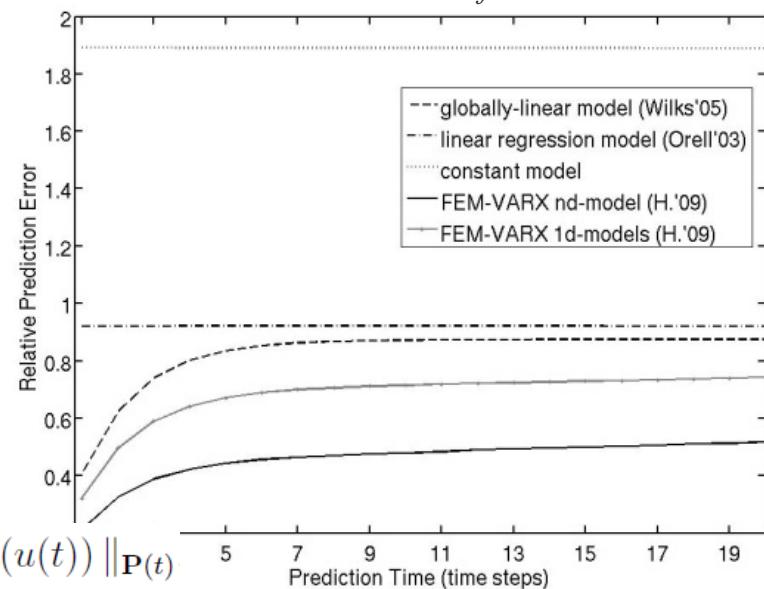


+



???

**non-stationary model
(time-dependent parameters)**
Predictors of Lorenz'96



$$x_t = \mu(t) + \mathbf{A}(t) \phi_1(x_{t-\tau}, \dots, x_{t-m\tau}) + \mathbf{B}(t) \phi_2(u_t) + \mathbf{C}(t) \epsilon_t$$

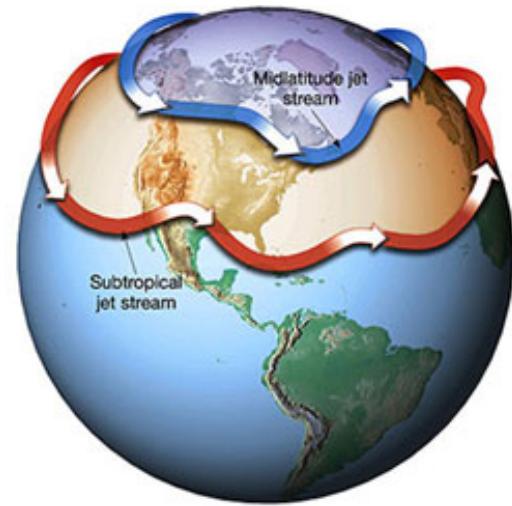
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Example II: analysis of geopotential data (Europe, 1958-2003, analysis of climate impact factors)

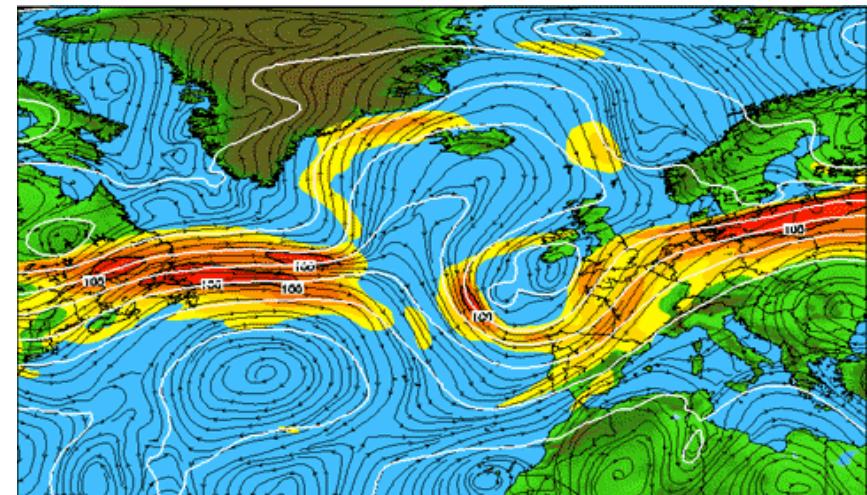
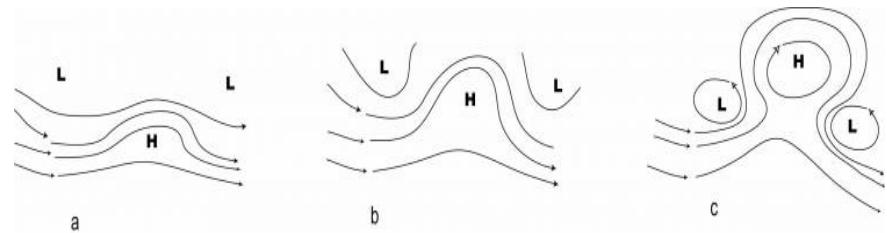
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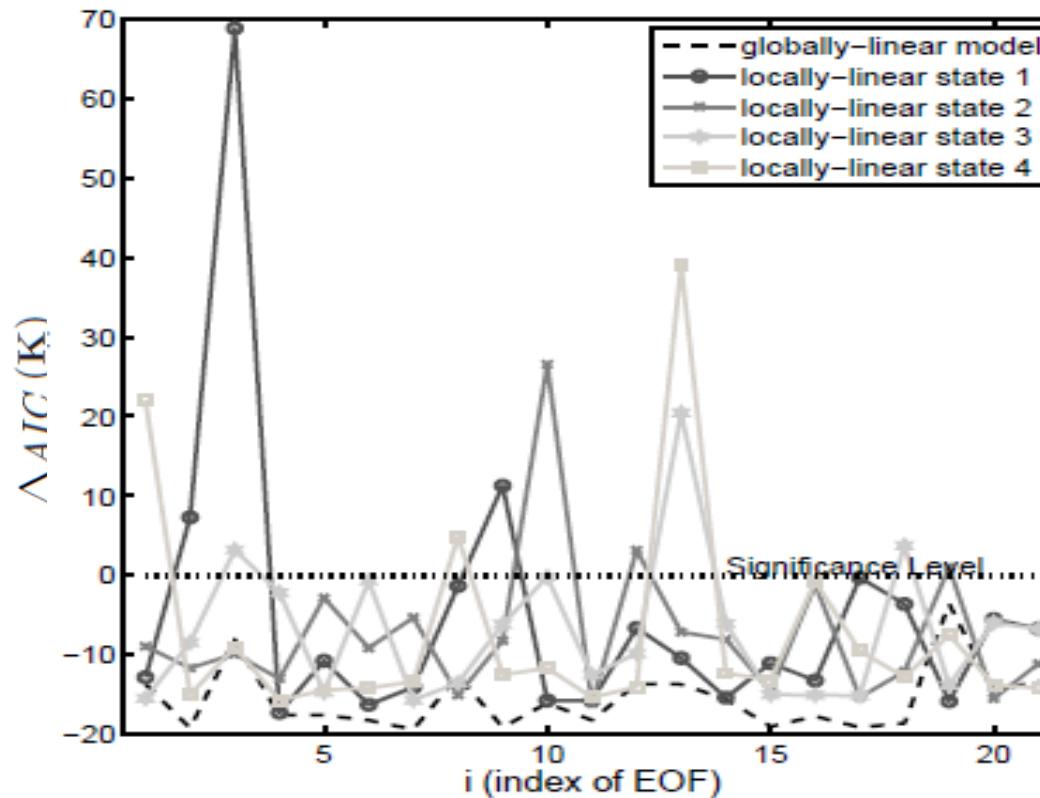


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non-stationary ($K=4$), stat. significant influence of **seasonal factor** and CO_2 ,
influence of **solar activity** insignificant

Stationarity: Yes or Not?
Answer: Not!

Example II: analysis of geopotential data (Europe, 1958-2003, analysis of climate impact factors)

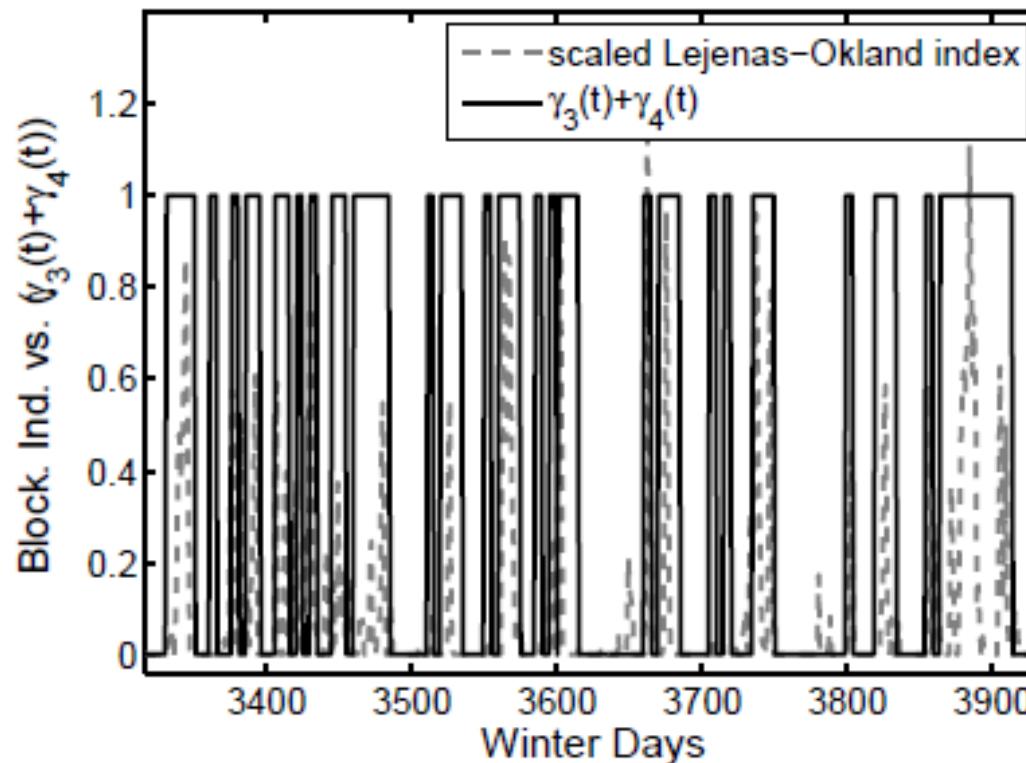


Figure 6: Comparison of the negative Lejenas-Oakland blocking index (dashed line) and the sum of cluster affiliations of locally-linear states 3 and 4 (solid line, calculated with FEM-VARX for $K = 4$, $N = 4000$, $C = 3000$, $q = 2$).

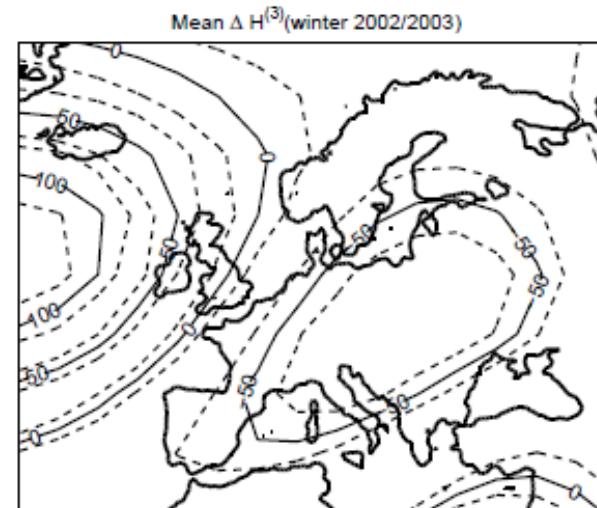
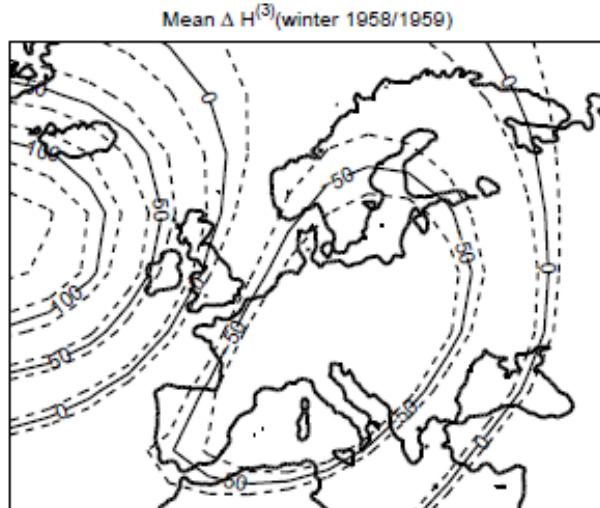
Example II: analysis of geopotential data (Europe, 1958-2003, analysis of climate impact factors)

$$\Delta H^{(i)}(u(t)) = \mathbb{E}^i[x_t]$$

$$x_t = \mu^i + \sum_{q=1}^m A_q^i x_{t-q\tau} + B^i \phi(u(t)) + C^i \epsilon_t$$

$\Rightarrow \Delta H^{(i)}(u(t)) = \mu^i + \sum_{q=1}^m A_q^i \Delta H^{(i)}(u(t)) + B^i \phi(u(t))$

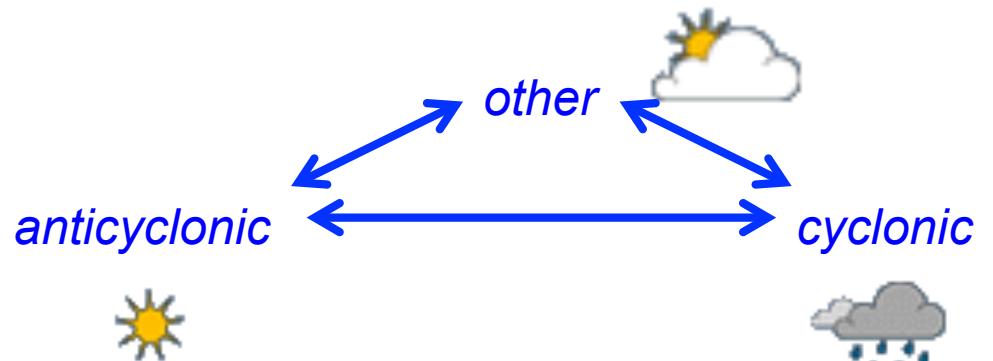
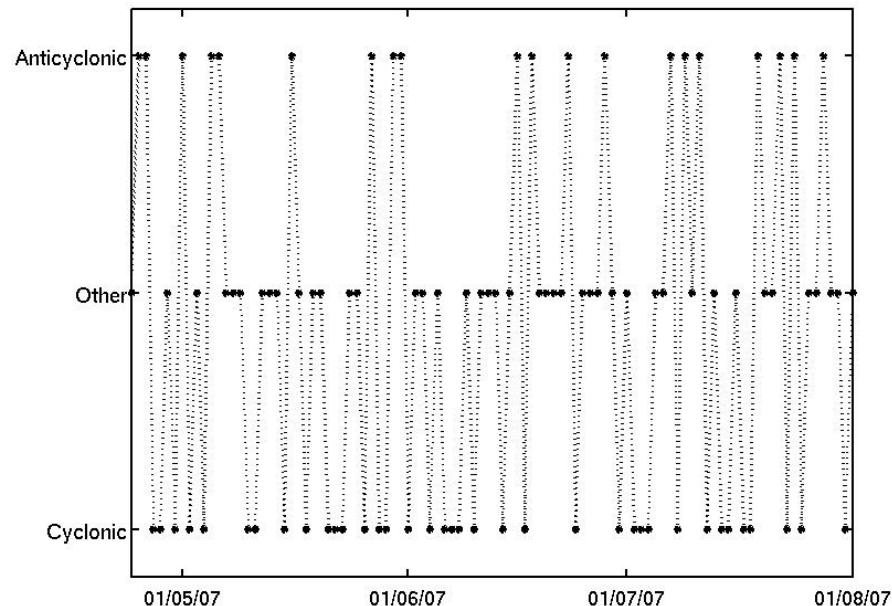
CO_2 impact on “indian summer“ local model ($i=3$)



Result: essential pressure impact factors
 CO_2 and *seasonal factor* (but **not solar activity**)

Lamb Circulation Patterns for UK (1945-2007)

*Historical Circulation Data: weather regimes
(Data from the Univ. of East Anglia)
3 atmospheric states considered*

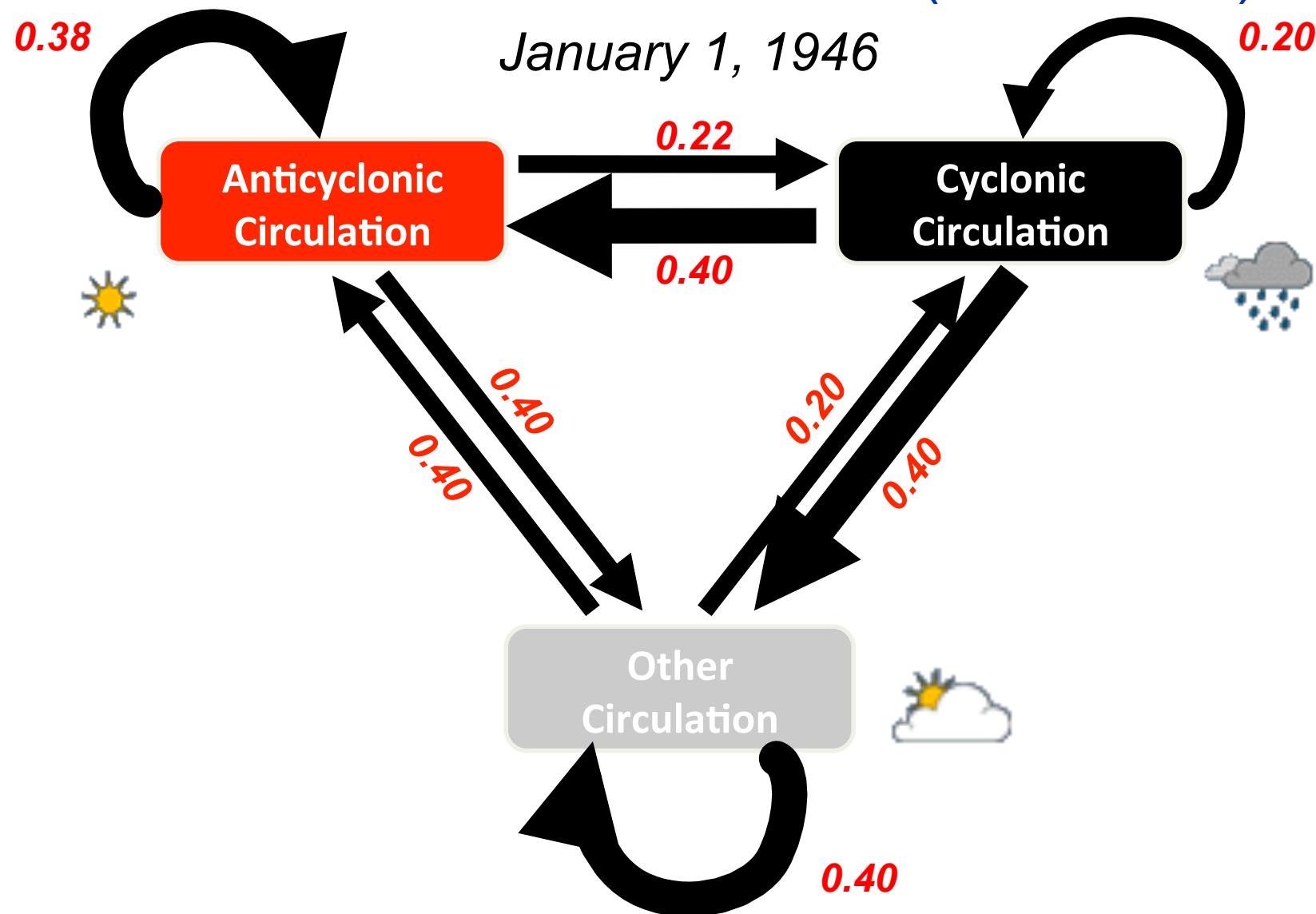


$$\mathbf{L}(\gamma_i(t), \mathbf{T}_i, \mu_i) \rightarrow \min$$

$$x_t : X_t \rightarrow X_{t+1}$$

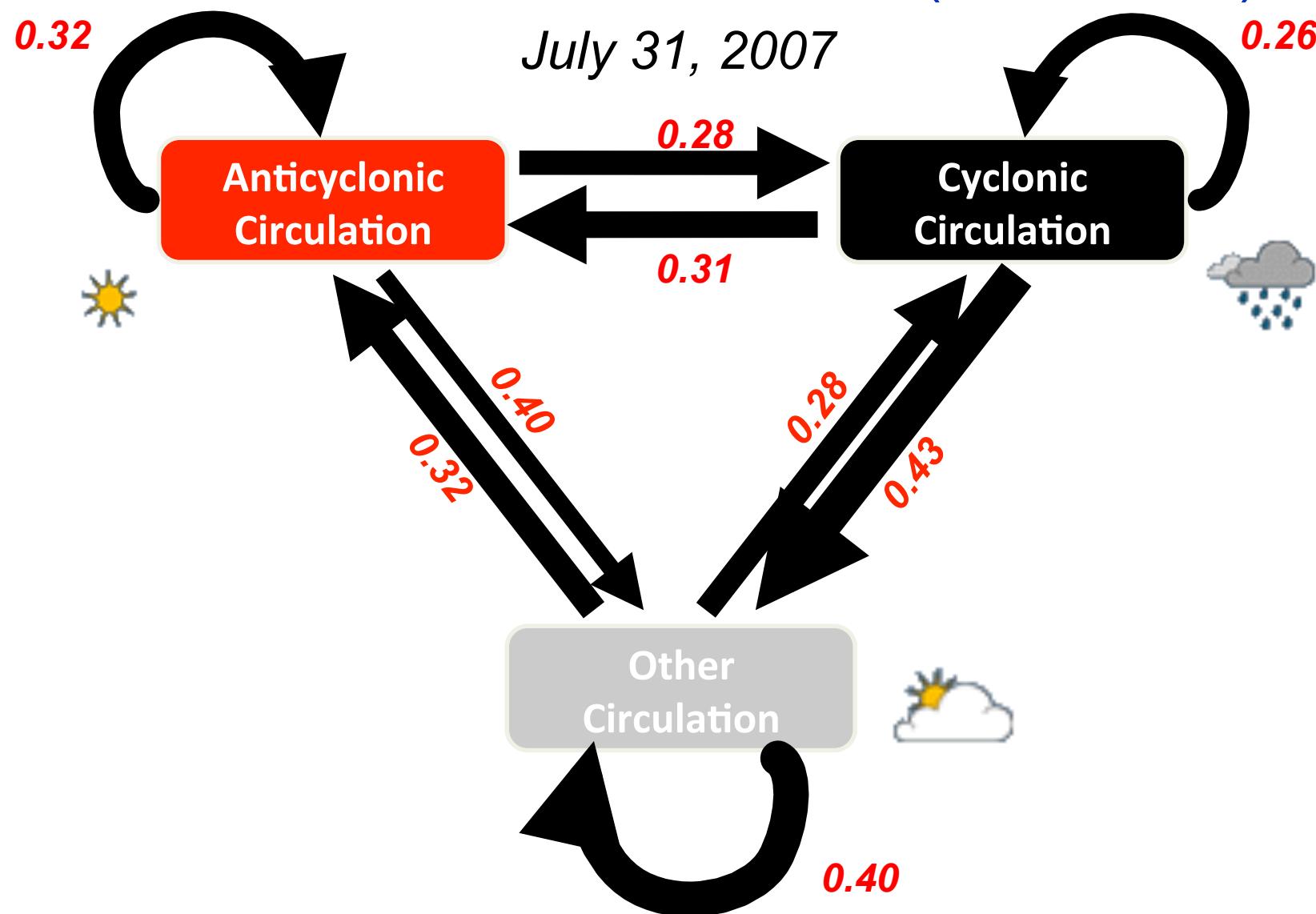
$$g(x_t, P^i) = -\log P^i_{X_t X_{t+1}}$$

Circulation Patterns for UK(1945-2007)



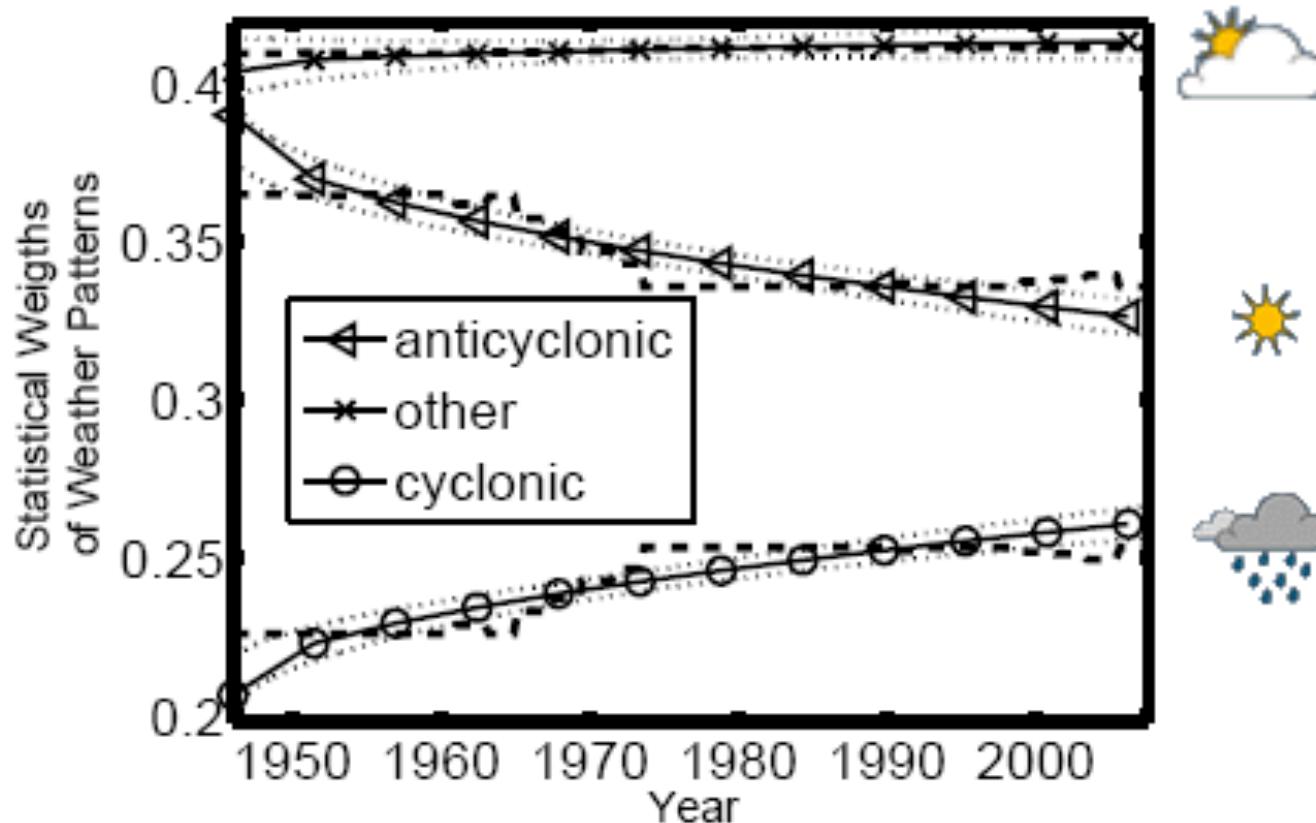
H., Journal of Atmos. Sci. (2009)

Circulation Patterns for UK(1945-2007)



Circulation Patterns for UK(1945-2007)

*Comparison of resulting a posteriori stat. weights
(obtained with two different sets of math. assumptions)*



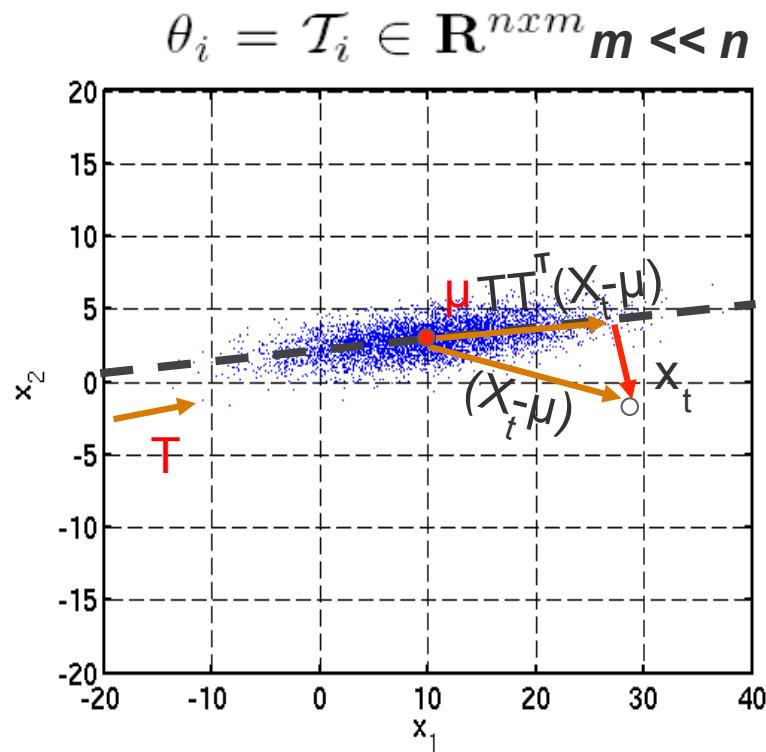
And now...



Dimension Reduction+Clustering

Problem 1: $\mathbf{L}(\gamma_i(t), \mathbf{T}_i, \mu_i) \rightarrow \min$

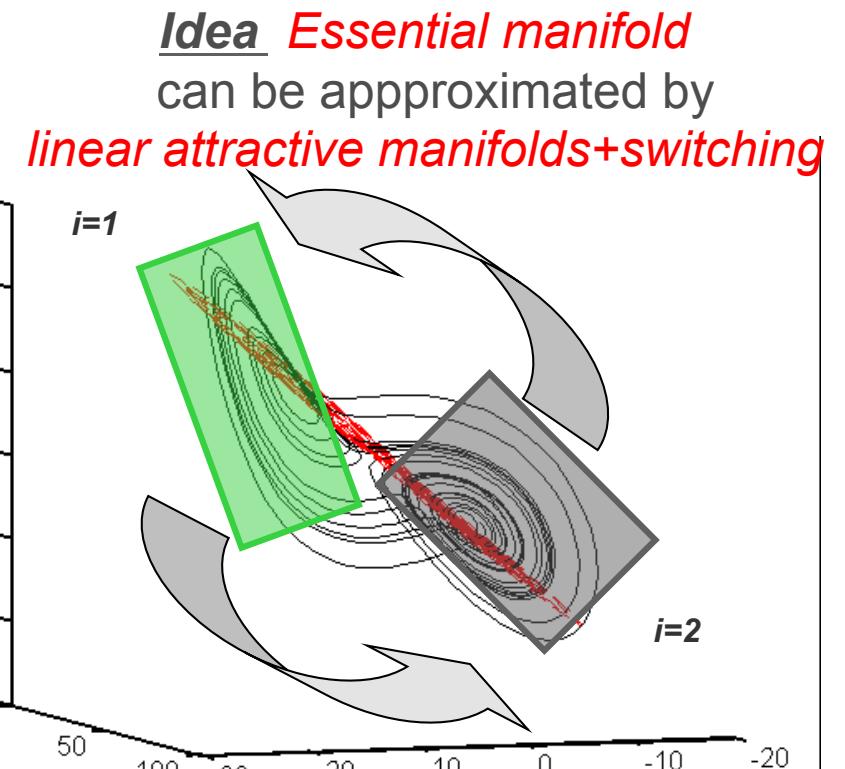
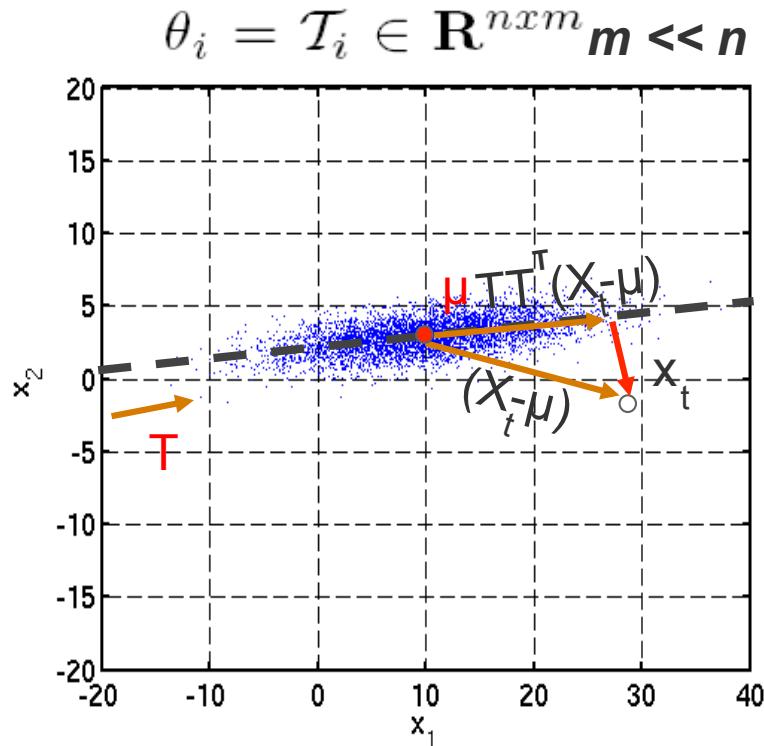
$$g(x, \theta_i) = \|x - \mathcal{T}_i \mathcal{T}_i^T x\|^2$$



Dimension Reduction+Clustering

Problem 1: $\mathbf{L}(\gamma_i(t), \mathbf{T}_i, \mu_i) \rightarrow \min$

$$g(x, \theta_i) = \|x - \mathcal{T}_i \mathcal{T}_i^T x\|^2$$



Wrap-Up: Methods

Given data \mathbf{x}_t , look for parameters $\theta(t)$

$$\mathbf{F}(x_t, \dots, x_{t-m\tau}, \theta(t), t) = 0 \\ (\text{direct problem})$$

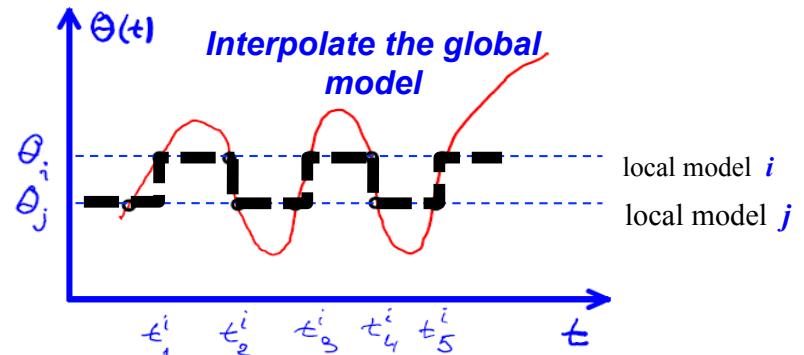
$$\int_0^T g(x_t, \theta(t)) dt \rightarrow \min_{\theta(t)} \\ (\text{inverse problem})$$

$$g(x_t, \theta(t)) = \sum_{i=1}^K \gamma_i(t) g(x_t, \theta_i) \\ (\text{ansatz})$$

- Extension from $W_{l,2}$ to BV -function spaces (allows parameter discontinuities)
- Data-based parameterization of non-stationary autoregressive factor models (stochastic modeling of sub-grid scale, **FEM-VARX**)

$$x_t = \mu(t) + \mathbf{A}(t) \phi_1(x_{t-\tau}, \dots, x_{t-m\tau}) + \mathbf{B}(t) \phi_2(u_t) + \mathbf{C}(t) \epsilon_t$$
- Deployment of *information theory* (optimal model discrimination)

$$\begin{aligned} \mathbf{L}(\Theta, \Gamma(t)) &= \int_0^T \sum_{i=1}^K \gamma_i(t) g(x_t, \theta_i) \rightarrow \min_{\Gamma(t), \Theta}, \\ \sum_{i=1}^K \gamma_i(t) &= 1, \quad \forall t \in [0, T] \\ \gamma_i(t) &\geq 0, \quad \forall t \in [0, T], \quad i = 1, \dots, K \\ \|\gamma_i(t)\| &\leq C_i \end{aligned}$$



Adaptive FEM-Clustering Methods: efficient numerical solvers

- Horenko, *SIAM Journal of Sci. Comp.*, 2009
- Horenko, *JAS*, 2009
- Horenko, *DYNAT*, 2010